

Bayesian Inference for a Stochastic Volatility Nelson-Siegel Model

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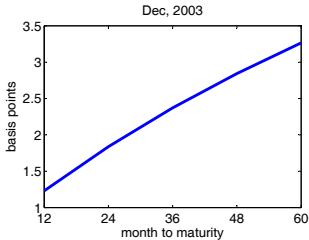
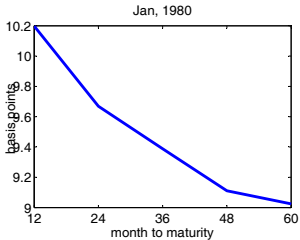
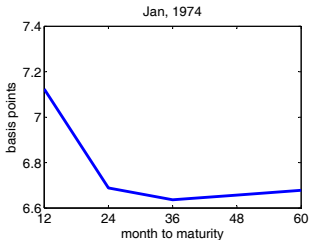
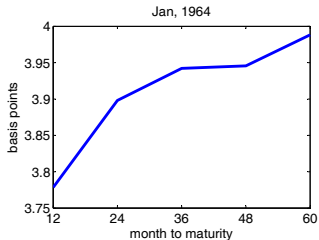
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Feb 12, 2010

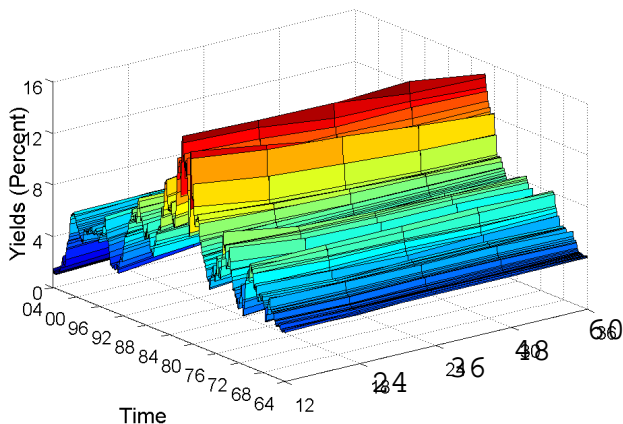
<mailto:fuyu.yang@cms.hu-berlin.de>



Different Shapes of Yield Curves



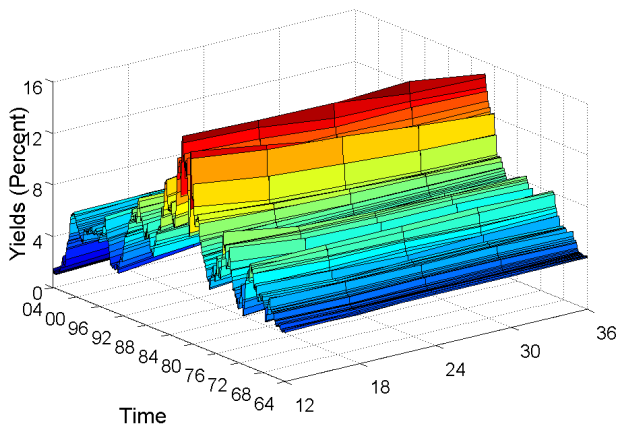
Yield Data Plot



Can we forecast yields?

- ▶ Use a parsimonious Nelson Siegel model to forecast yields

Yield Data Plot



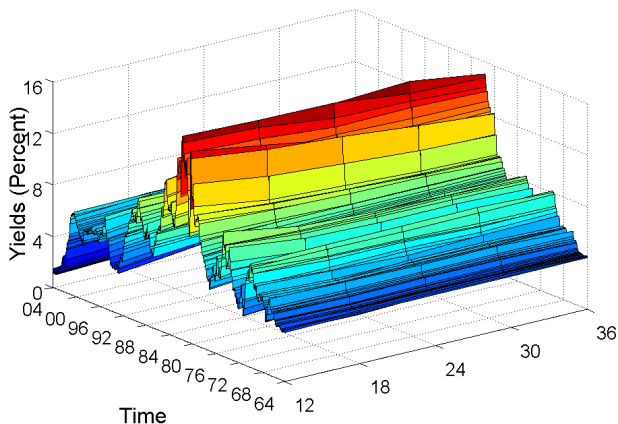
How can we capture the time varying variances in all yields?

- ▶ We propose a Stochastic Volatility Nelson Siegel model.

SVNS



Yield Data Plot



How we do density forecast for yields?

- ▶ Can be achieved in a Bayesian framework without extra cost.

Why we care about yield curve forecasts?

- ▶ For monetary policy makers:
 - ▷ The yield curve carries information about prospective evolution of economic activity, inflation etc. **e.g.** an inverted yield curve indicates that investors expect the economy to slow down in the future.
 - ▷ Interest rate spreads (difference between 10 year and 3 month rates, or 10 year and 2 year rates) are predictive indicators of recession and inflation **e.g.** steep yield curve can be observed at the beginning of economic expansion (after recession).

- ▶ For investors...

Outline

1. Model Specifications
2. Estimation in Bayesian Framework
3. Forecasting in Bayesian Framework
4. Empirical Results
5. Conclusion

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Nelson Siegel Model (1987)

$$y_t(n) = f_{1,t} + f_{2,t} \left[\frac{1 - e^{-\lambda \cdot n}}{\lambda \cdot n} \right] + f_{3,t} \left[\frac{1 - e^{-\lambda \cdot n}}{\lambda \cdot n} - e^{-\lambda \cdot n} \right]$$

- ▶ $y_t(n)$: yields at time t , with n period to maturity.

Diebold Li Nelson Siegel Model (2006)

- ▶ $[f_{1,t}, f_{2,t}, f_{3,t}]$: level, slope, curvature factors
- ▶ $\left[1, \frac{1 - e^{-\lambda \cdot n}}{\lambda \cdot n}, \frac{1 - e^{-\lambda \cdot n}}{\lambda \cdot n} - e^{-\lambda \cdot n} \right]$: factor loadings.

Denote: $\mathbf{y}_t = \{y_t(1), y_t(2), \dots, y_t(N)\}'$, $\mathbf{f}_t = \{f_{1,t}, f_{2,t}, f_{3,t}\}'$.

$$\mathbf{y}_t = \mathbf{A}\mathbf{f}_t + \varepsilon_t \quad (1)$$

$$\mathbf{f}_t - \mu_f = \Phi_f (\mathbf{f}_{t-1} - \mu_f) + \mathbf{u}_t \quad (2)$$

where $\varepsilon_t \stackrel{i.i.d}{\sim} f_{MN}(0_N, H_\varepsilon)$, and $\mathbf{u}_t \stackrel{i.i.d}{\sim} f_{MN}(0_N, H_u)$ and

$$\mathbf{A} = \begin{pmatrix} 1 & \frac{1-e^{-\lambda \cdot 1}}{\lambda \cdot 1} & \frac{1-e^{-\lambda \cdot 1}}{\lambda \cdot 1} & -e^{-\lambda \cdot 1} \\ 1 & \frac{1-e^{-\lambda \cdot 2}}{\lambda \cdot 2} & \frac{1-e^{-\lambda \cdot 2}}{\lambda \cdot 2} & -e^{-\lambda \cdot 2} \\ \vdots & \vdots & \vdots & \vdots \\ 1 & \frac{1-e^{-\lambda \cdot n}}{\lambda \cdot n} & \frac{1-e^{-\lambda \cdot n}}{\lambda \cdot n} & -e^{-\lambda \cdot n} \end{pmatrix} : N \times 3$$

Stochastic Volatility Nelson Siegel Model (2010)

$$\mathbf{f}_t - \mu_f = \Phi_f (\mathbf{f}_{t-1} - \mu_f) + \mathbf{u}_t \quad (3)$$

$$\mathbf{u}_t = \begin{pmatrix} \sqrt{\exp h_{1,t}} & 0 & 0 \\ 0 & \sqrt{\exp h_{2,t}} & 0 \\ 0 & 0 & \sqrt{\exp h_{3,t}} \end{pmatrix} \times \zeta_t \quad (4)$$

$$f_{j,t} - \mu_{j,f} = \phi_{j,f} (f_{j,t-1} - \mu_{j,f}) + \exp\left(\frac{h_{j,t}}{2}\right) \zeta_{j,t} \quad (5)$$

$$h_{j,t} - \mu_{j,h} = \phi_{j,h} (h_{j,t-1} - \mu_{j,h}) + \sigma_j \epsilon_{j,t} \quad (6)$$

where $j = 1, 2, 3$, $\zeta_{j,t} \stackrel{i.i.d}{\sim} f_N(0, 1)$ and $\epsilon_{j,t} \stackrel{i.i.d}{\sim} f_N(0, 1)$

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Estimation

Estimating the model is challenging

- ▶ Parameters have to be estimated:
 $\theta = \{\lambda, H_\varepsilon, [\mu_{j,f}, \phi_{j,f}], [\mu_{j,h}, \phi_{j,h}, \sigma_j]\}$ and the latent variables f_j and h_j for $j = 1, 2, 3$.
- ▶ The latent yield factors f and log Volatilities h are highly correlated, blocking f and h is needed (Chib and Greenberg 1996)
- ▶ An efficient MCMC algorithm is a solution.

Blocking the highly correlated latent variables

$$\begin{aligned} \mathbf{y}_t &= \mathbf{A}\mathbf{f}_t + \varepsilon_t \\ \mathbf{f}_t - \mu_{\mathbf{f}} &= \Phi_{\mathbf{f}}(\mathbf{f}_{t-1} - \mu_{\mathbf{f}}) + \mathbf{u}_t \\ f_{j,t}^* &= h_{j,t} + z_{j,t} \\ h_{j,t+1} - \mu_{j,h} &= \phi_{j,h}(h_{j,t} - \mu_{j,h}) + \sigma_j \epsilon_{j,t} : j = 1, 2, 3 \end{aligned}$$

where

$$\begin{aligned} f_{j,t}^* &= \ln \left\{ [f_{j,t} - \mu_{j,f} - \phi_{j,f}(f_{j,t-1} - \mu_{j,f})]^2 + c \right\} \\ z_{j,t} &= \ln \left(\zeta_{j,t}^2 \right) \end{aligned}$$

Remarks

1. c is the Fuller offset parameter, we choose $c = 0.0001$.
2. The SVNS model can be fitted in the state space form.
3. Approximate the distribution of $z_{j,t}$ by 7 mixture of normal densities (Kim, Chib, Shephard, 1998).

$$z_t | s_t \sim f_N \left(m_{s_t}, \nu_{s_t}^2 \right)$$

$$\Pr(s_t = i) = q_i, i \leq 7, t \leq T$$

4. This approximation allows us to draw the whole block $\{h_j\}$
5. The approximation error can be removed at the end of posterior sampling by a re-weighting procedure (Chib, 2002)

MCMC Algorithm

1. Initialize $\theta = \{\lambda, H_\varepsilon, [\mu_{j,f}, \phi_{j,f}], [\mu_{j,h}, \phi_{j,h}, \sigma_j]\}$ and the latent variables f_j and h_j with $f_{j,0} = \mu_{j,f}$ and $h_{j,0} = \mu_{j,h}$ for $j = 1, 2, 3$.
2. Run a Gibbs sampler for steps (a) - (d) using S replications, where the initial S_0 draws are discarded:
 - 2.1 Sample H_ε from $H_\varepsilon | y, \lambda, [\mu_{j,f}, \phi_{j,f}], [\mu_{j,h}, \phi_{j,h}, \sigma_j], f_j, h_j, \quad j = 1, 2, 3$.
 - 2.2 Sample λ from $\lambda | y, H_\varepsilon, [\mu_{j,f}, \phi_{j,f}], [\mu_{j,h}, \phi_{j,h}, \sigma_j], f_j, h_j$ using a Griddy-Gibbs sampling method.
 - 2.3 Sample $[\mu_{j,f}, \phi_{j,f}], f_j$ from $[\mu_{j,f}, \phi_{j,f}], f_j | y, \lambda, H_\varepsilon, h_j$ using a simulation smoother.

2.4 Run (i)-(iv) 3 times to estimate $h_j | f_j$, with $j = 1, 2, 3$, respectively:

- (i) Compute f_j^* and run the loop (ii)-(iv) 2000 times. Discard the results from the initial 500 loops.
- (ii) Sample $s | f_j^*, h_j$.
- (iii) Sample $h_j | f_j^*, s, [\mu_{j,h}, \phi_{j,h}, \sigma_j]$ using a simulation smoother.
- (iv) Sample $[\mu_{j,h}, \phi_{j,h}, \sigma_j] | f_j^*, h_j$ using the Metropolis-Hastings algorithm proposed by Chib et al. (2002) choosing a multivariate t -density as the proposal density.

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Forecasting of Yields

- ▶ *h*-step ahead Point forecast

$$\hat{f}_{t+h} = \hat{\mu}_f + \hat{\phi}_f \left(\hat{f}_t - \hat{\mu}_f \right), \quad (7)$$

$$\hat{y}_{t+h} = \hat{A} \hat{f}_{t+h}, \quad (8)$$

- ▶ *h*-step ahead Density forecast

$$\hat{f}_{t+h}^{(s)} = \hat{\mu}_f^{(s)} + \hat{\phi}_f^{(s)} \left(\hat{f}_t^{(s)} - \hat{\mu}_f^{(s)} \right) + \hat{\eta}_{t+h}^{(s)},$$

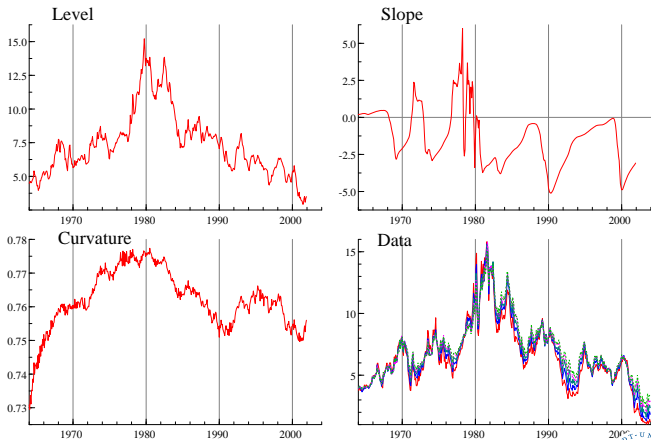
$$\hat{y}_{t+h}^{(s)} = \hat{A}^{(s)} \left(\hat{f}_{t+h}^{(s)} + \hat{\eta}_{t+h}^{(s)} \right) + \hat{\varepsilon}_{t+h}^{(s)},$$

where $\hat{\eta}_{t+h}^{(s)} = \Sigma_{\eta}^{(s)} u_{\eta,t+h}$ and $\hat{\varepsilon}_{t+h}^{(s)} = \Sigma_{\varepsilon}^{(s)} u_{\varepsilon,t+h}$, $u_{\varepsilon,t+h}$ and $u_{\eta,t+h}$ are i.i. standard normally distributed.

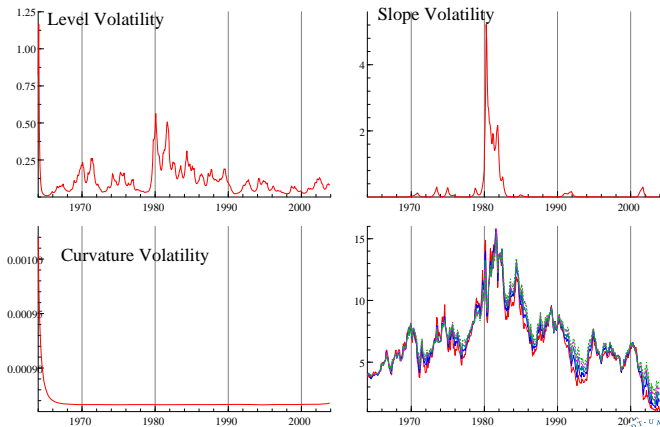
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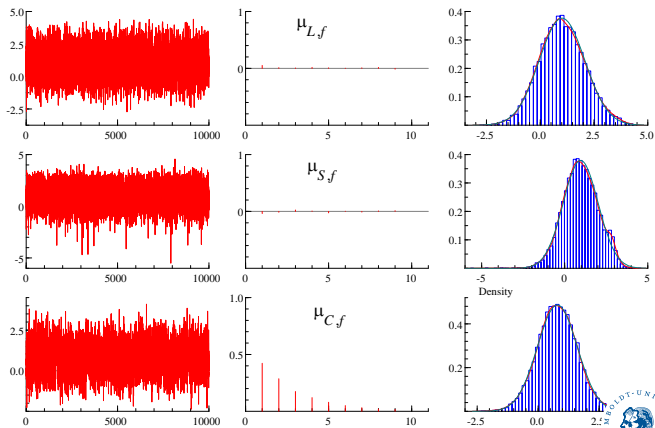
Extracted yield factors f



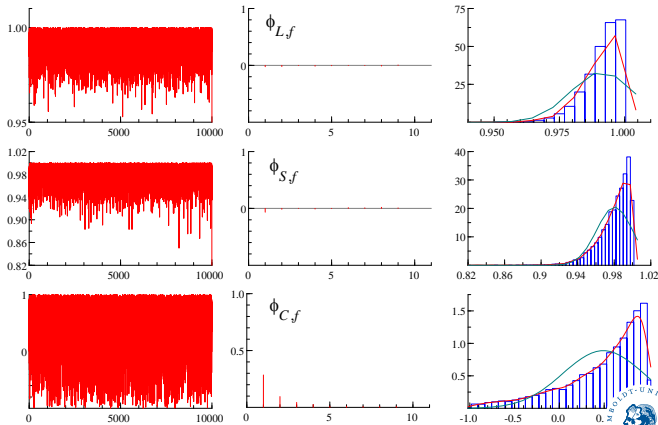
Extracted log Volatilities h



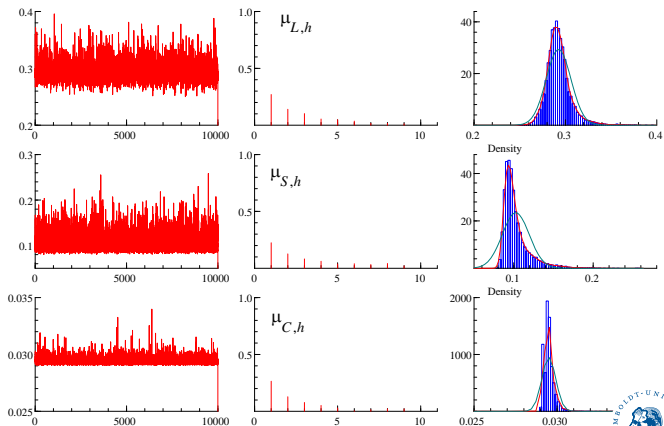
Draws, autocorrelations and posterior densities μ_f



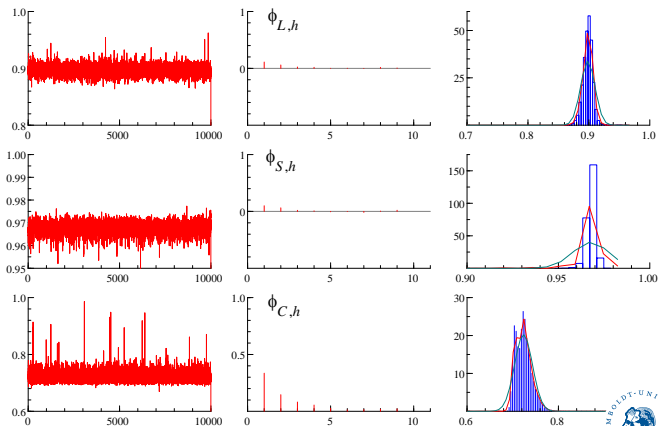
Draws, autocorrelations and posterior densities ϕ_f



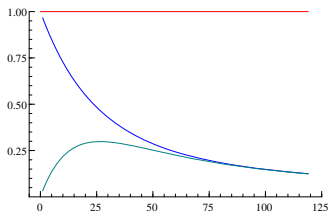
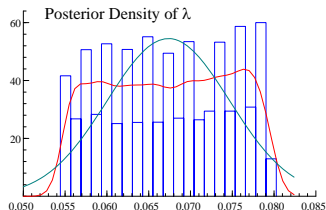
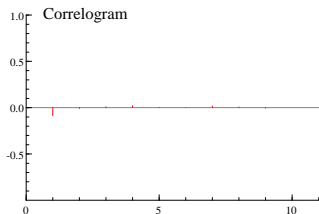
Draws, autocorrelations and posterior densities μ_h



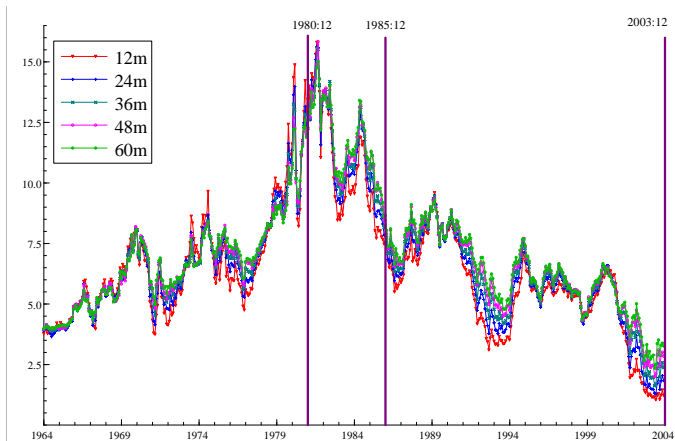
Draws, autocorrelations and posterior densities ϕ_h



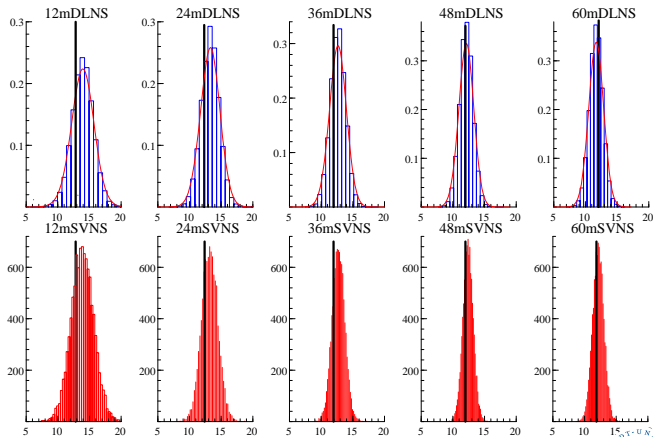
Draws, autocorrelations and posterior densities λ



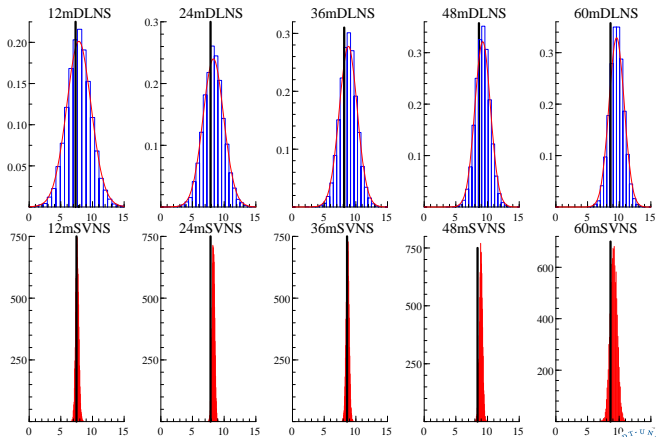
Forecast using 3 sub samples



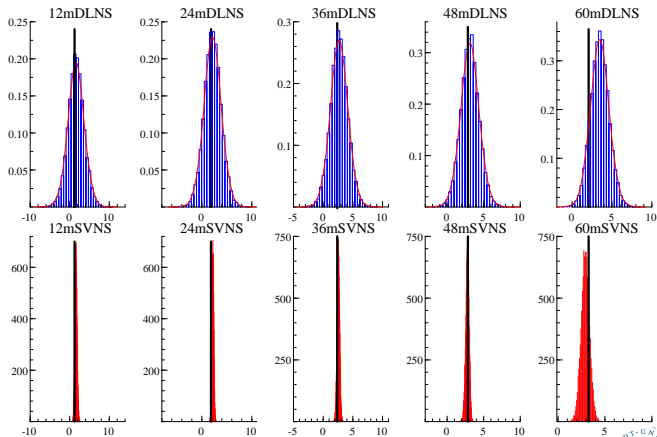
Forecast the yields for Dec, 1980



Forecast the yields for Dec, 1985



Forecast the yields for Dec, 2003



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Conclusion

- ▶ The proposed algorithm is computationally efficient.
- ▶ The SVNS fits the data well according to posterior predictive p-value, see Hautsch and Yang (2010).
- ▶ Assuming a constant variance(i.e. DLNS model), the forecast density reveals larger uncertainty than it should be, since the data with high volatilities "contaminate" the estimate of the variances.
- ▶ Using Bayesian method, the uncertainty associates with the point forecast can be quantified.