

Local Adaptive Multiplicative Error Models for High-Frequency Forecasts

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Statistical Challenges

□ Understanding high-frequency dynamics

- ▶ Time-varying parameters [▶ Parameter Dynamics](#)
- ▶ Regime shifts

□ Modelling using procrustean assumptions

- ▶ Time-invariant parameters
- ▶ Transition form, number of regimes, transition variable type



Objectives

► Economic Benefits

- (i) Localising Multiplicative Error Models (MEM)
 - ▶ Local parametric approach (LPA)
 - ▶ Balance between modelling bias and parameter variability
 - ▶ Estimation windows with potentially varying lengths
- (ii) Short-term forecasting
 - ▶ Case study: trading volume
 - ▶ Evaluation against standard approach - fixed estimation length on an ad hoc basis



Example

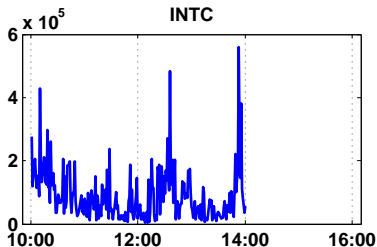


Figure 1: One-minute cumulated trading volume for Intel Corporation (INTC) on 20080902



Example

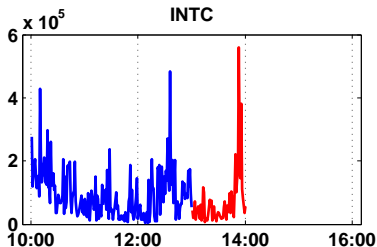


Figure 2: One-minute cumulated trading volume for Intel Corporation (INTC) on 20080902 with an **estimation window length of 60**



Example

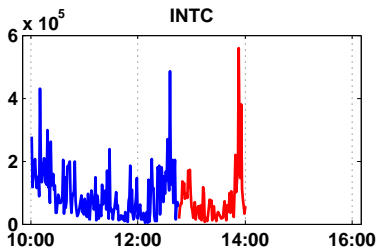


Figure 3: One-minute cumulated trading volume for Intel Corporation (INTC) on 20080902 with an **estimation window length of 75**



Example

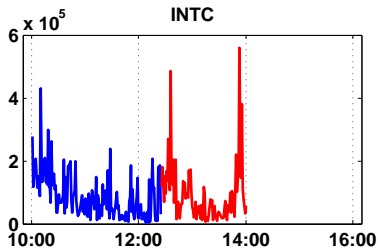


Figure 4: One-minute cumulated trading volume for Intel Corporation (INTC) on 20080902 with an **estimation window length of 95**



Example

Example: Short-term forecasting

Forecasting strategies up to the next one hour:

- (i) 'Standard' method - fixed estimation window (one day/week)
- (ii) LPA technique with adaptively selected interval of homogeneity



Outline

1. Motivation ✓
2. Multiplicative Error Models (MEM)
3. Local Parametric Approach
4. Empirical Results
5. Conclusions




Multiplicative Error Models (MEM)

- Engle (2002), $\text{MEM}(p, q)$, \mathcal{F}_i - information set up to i

$$y_i = \mu_i \varepsilon_i, \quad E[\varepsilon_i | \mathcal{F}_{i-1}] = 1$$

$$\mu_i = \omega + \sum_{j=1}^p \alpha_j y_{i-j} + \sum_{j=1}^q \beta_j \mu_{i-j}, \quad \omega > 0, \alpha_j, \beta_j \geq 0$$

- y_i - squared (de-meaned) log return: GARCH(p, q)
- y_i - volume, bid-ask spread, duration: ACD(p, q)

Engle, Robert F. on BBI: 



Parametric Modelling

1. Exponential-ACD, Engle and Russel (1998) ▶ EACD
 $\varepsilon_i \sim \text{Exp}(1)$, $\theta_E = (\omega, \alpha, \beta)^\top$, $\alpha = (\alpha_1, \dots, \alpha_p)^\top$, $\beta = (\beta_1, \dots, \beta_q)^\top$

2. Weibull-ACD, Engle and Russel (1998) ▶ WACD
 $\varepsilon_i \sim \mathcal{G}(s, 1)$, $\theta_W = (\omega, \alpha, \beta, s)^\top$

□ Quasi maximum likelihood estimates (QMLEs) of θ_E and θ_W

$$\tilde{\theta}_I = \arg \max_{\theta \in \Theta} L_I(y; \theta) \quad (1)$$

- ▶ Data interval (fix i_0 , length n) $I = [i_0 - n, i_0]$
- ▶ Quasi log likelihood $L_I(\cdot)$, see (2) and (3)

Weibull, E. H. Waloddi on BBI:



Parameter Dynamics

► Statistical Challenges

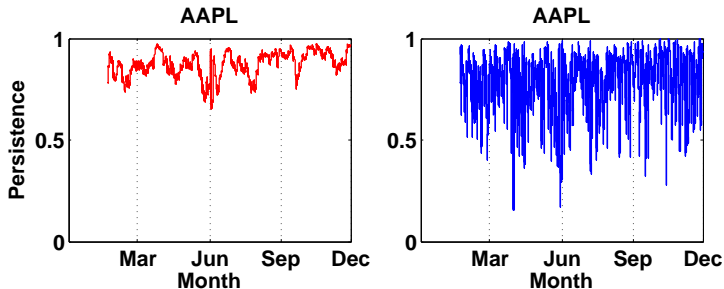


Figure 5: Estimated **weekly** ($n = 1800$) and **daily** ($n = 360$) persistence $\tilde{\alpha}_i + \tilde{\beta}_i$ for seasonally adjusted trading volume using an EACD(1, 1) at each minute in 2008



Estimation Quality

- Quality of estimating θ^* by QMLE $\tilde{\theta}_I$ in terms of Kullback-Leibler divergence; $\mathcal{R}_r(\theta^*)$ - risk bound

$$\mathbb{E}_{\theta^*} \left| L_I(\tilde{\theta}_I) - L_I(\theta^*) \right|^r \leq \mathcal{R}_r(\theta^*)$$

'Modest' risk, $r = 0.5$ (shorter intervals of homogeneity)

'Conservative' risk, $r = 1$ (longer intervals of homogeneity)

Kullback, Solomon and Leibler, Richard A. on BBI:



Local Parametric Approach (LPA)

- ▣ Volatility modelling - Mercurio and Spokoiny (2004)
- ▣ GARCH(1,1) models - Čížek et al. (2009)
- ▣ Realized volatility - Chen et al. (2010)

- ▣ LPA: time series parameters can be locally approximated
- ▣ Balance between modelling bias and parameter variability



Statistical Framework

- LPA: Spokoiny (1998, 2009)
- Small modelling bias (SMB) - 'Oracle' interval
- Local change point detection test - *Interval of homogeneity*
- Adaptive estimate $\hat{\theta}$ - QMLE at the *interval of homogeneity*



Interval Selection

□ $(K + 1)$ nested intervals with length $n_k = |I_k|$

$$\begin{matrix} I_0 \\ \tilde{\theta}_0 \end{matrix} \subset \begin{matrix} I_1 \\ \tilde{\theta}_1 \end{matrix} \subset \cdots \subset \begin{matrix} I_k \\ \tilde{\theta}_k \end{matrix} \subset \cdots \subset \begin{matrix} I_K \\ \tilde{\theta}_K \end{matrix}$$

Example: Trading volumes aggregated over 1-min periods

Fix i_0 , $I_k = [i_0 - n_k, i_0]$, $n_k = \lceil n_0 c^k \rceil$, $c > 1$

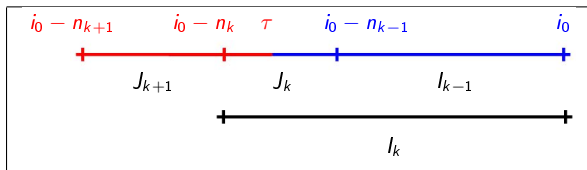
$\{n_k\}_{k=0}^{13} = \{60 \text{ min.}, 75 \text{ min.}, \dots, 1 \text{ week}\}$, $c = 1.25$



Local Change Point Detection ▶ Example

□ Fix i_0 , sequential test ($k = 1, \dots, K$)

H_0 : parameter homogeneity within l_k vs. H_1 : change point within l_k



$$T_k = \sup_{\tau \in J_k} \left\{ L_{A_{k,\tau}} \left(\tilde{\theta}_{A_{k,\tau}} \right) + L_{B_{k,\tau}} \left(\tilde{\theta}_{B_{k,\tau}} \right) - L_{l_{k+1}} \left(\tilde{\theta}_{l_{k+1}} \right) \right\},$$

with $J_k = l_k \setminus l_{k-1}$, $A_{k,\tau} = [i_0 - n_{k+1}, \tau]$ and $B_{k,\tau} = (\tau, i_0]$



Critical Values, z_k

- Simulate z_k under the null of the homogeneity of the interval sequence I_0, \dots, I_k
- Check z_k for different θ^*
 - ▶ Nine different parameters θ^* for each model (EACD and WACD) and risk level ($r = 0.5$ and $r = 1$)
 - ▶ Findings: z_k are virtually invariable w.r.t. θ^* given a scenario
Largest differences at first two or three steps



Critical Values, \mathfrak{z}_k

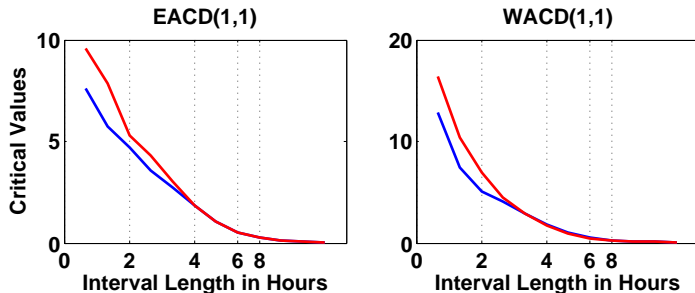


Figure 6: Critical values for low ($\tilde{\alpha} + \tilde{\beta} = 0.84$) and high ($\tilde{\alpha} + \tilde{\beta} = 0.93$) weekly persistence and 'modest' risk ($r = 0.5$)



Adaptive Estimation

- Compare T_k at every step k with z_k
- Data window index of the *interval of homogeneity* - \hat{k}
- Adaptive estimate

$$\hat{\theta} = \tilde{\theta}_{\hat{k}}, \quad \hat{k} = \max_{k \leq K} \{k : T_\ell \leq z_\ell, \ell \leq k\}$$

- Note: rejecting the null at $k = 1$, $\hat{\theta}$ equals QMLE at l_0
If the algorithm goes until K , $\hat{\theta}$ equals QMLE at l_K



Data

- NASDAQ Stock Market in 2008, 250 trading days
 - ▶ 5 stocks: AAPL, CSCO, INTC, MSFT and ORCL
 - ▶ \check{y}_i - one-minute cumulated trading volume from 10:00-16:00
 - ▶ y_i - seasonally adjusted trading volume

- Periodicity effect - FFS approximation, Gallant (1981)
 - ▶ 30-days rolling windows, Engle and Rangel (2008)
 - ▶ Order $M = 1, \dots, 6$ selected by BIC, $\bar{t} \in [0, 1]$

$$y_i = \check{y}_i / [\delta \cdot \bar{t} + \sum_{m=1}^M \{ \delta_{c,m} \cos(\bar{t} \cdot 2\pi m) + \delta_{s,m} \sin(\bar{t} \cdot 2\pi m) \}]$$



Intraday Periodicity

Figure 7: Estimated intraday periodicity components for AAPL, $M = 6$



Adaptive Estimation

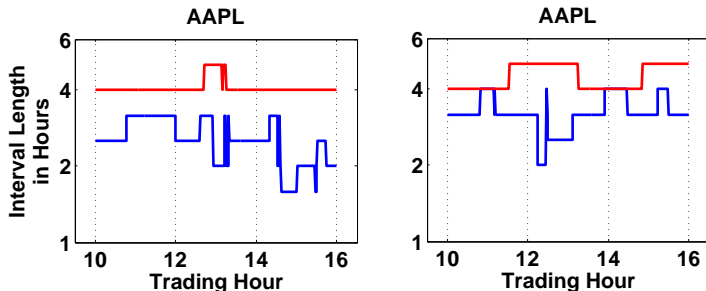


Figure 8: Estimated interval length $n_{\hat{k}}$ for the **conservative** ($r = 1$) and **modest** ($r = 0.5$) risk case on 20080902 (left, lowest volume) and 20081030 (right, highest volume) using the EACD model



Adaptive Estimation

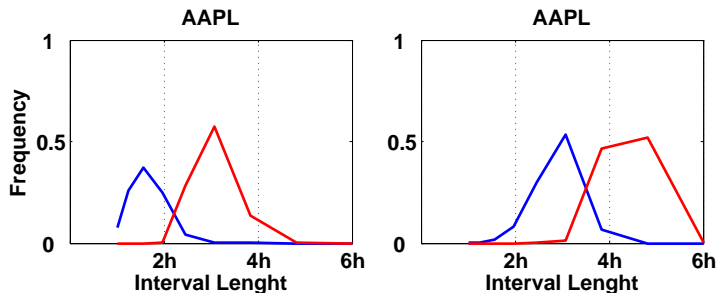


Figure 9: Estimated interval length $n_{\hat{k}}$ for the **conservative** ($r = 1$) and **modest** ($r = 0.5$) risk case in 2008 using the EACD (left) and WACD (right) models



Forecasting

Strategies

- ▣ LPA technique with adaptively selected interval of homogeneity
- ▣ 'Standard' method: 360 (1 day) or 1800 observations (1 week)

Setup

- ▣ Forecasting period: 20080222 - 20081222 (210 trading days)
- ▣ Forecasts at each minute (horizon $h = 1, \dots, 60$ min.)

(a) Sign test, (b) Diebold-Mariano test, (c) Forecasting regressions



Forecasting Accuracy

1. *Overall performance* - LPA qualitatively outperforms 'standard' methods, quantitatively equal to 1-day 'standard' method
2. *Forecasting horizon* - performance best at approx. 3-4 minutes
3. *Course of a typical trading day* - best results after 14:00
4. *Model specifications and tuning parameters* - WACD slightly better, insensitivity



Forecasting Accuracy

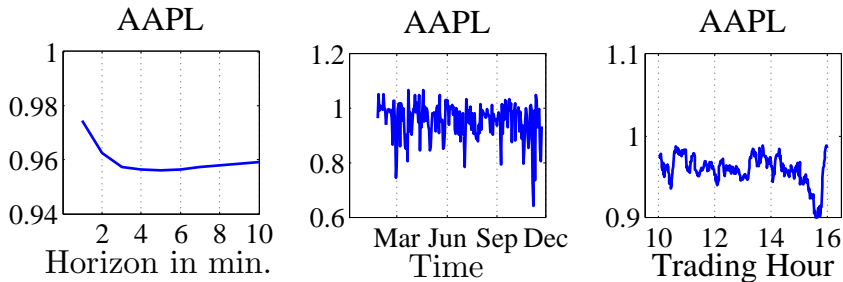


Figure 10: Ratio between the RMSPE of the LPA method and the RMSPE of the standard approach using an EACD(1,1) model for Apple Inc. (AAPL) from 20100222 to 20101222 (210 trading days)



Forecasting Accuracy

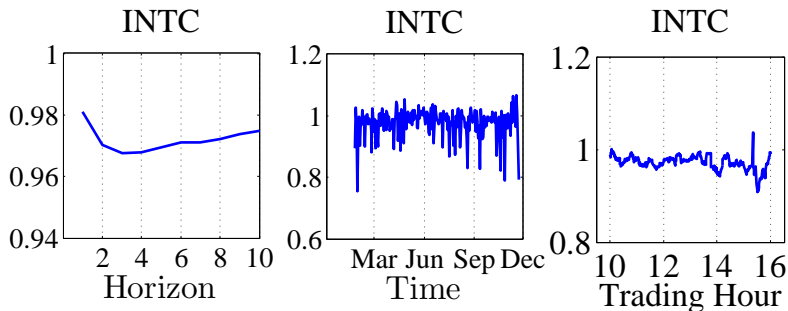


Figure 11: Ratio between the RMSPE of the LPA method and the RMSPE of the standard approach using an EACD(1, 1) model for Intel Corporation (INTC) from 20100222 to 20101222 (210 trading days)



Conclusions

Localising MEM

- Time-varying parameters and estimation quality
- LPA - 5 stocks in 2008 (79200 minutes): AAPL, CSCO, INTC, MSFT and ORCL
- Precise adaptive estimation ($r = 1$) requires 4-5 hours of data, modest risk approach ($r = 0.5$) requires 2-3 hours

Forecasting Trading Volume

- LPA outperforms the 'standard' method
- Overall performance, horizon, trading day, tuning parameters



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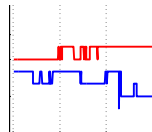
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Economic Benefits

► Objectives

Flexible statistical framework

- MEM parameter dynamics: strength and frequency
- Typical interval lengths of parameter homogeneity
- Tuning parameters (sensitivity)
- Short-term forecasting



Exponential-ACD (EACD)

► Parametric Modelling

□ Engle and Russel (1998), $\varepsilon_t \sim \text{Exp}(1)$

$$L_I(\theta) = \sum_{t \in I} \left(-\log \mu_t - \frac{y_t}{\mu_t} \right) \quad (2)$$

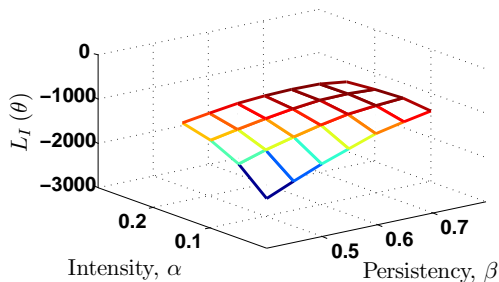


Figure 12: Quasi log likelihood - EACD(1,1), $\theta^* = (0.10, 0.20, 0.65)^\top$



Weibull-ACD (WACD)

► Parametric Modelling

□ Engle and Russel (1998), $\varepsilon_t \sim$ standardised $\mathcal{G}(s, 1)$

$$L_I(\theta) = \sum_{t \in I} \left[\log \frac{s}{y_t} + s \log \frac{\Gamma(1 + 1/s) y_t}{\mu_t} - \left\{ \frac{\Gamma(1 + 1/s) y_t}{\mu_t} \right\}^s \right] \quad (3)$$

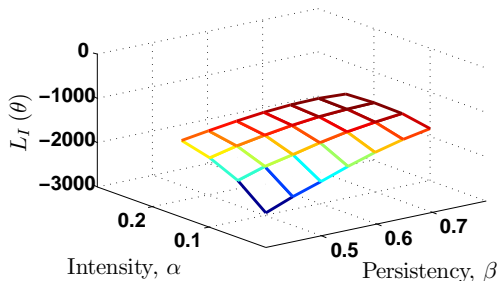


Figure 13: Quasi log likelihood - WACD(1,1), $\theta^* = (0.10, 0.20, 0.65, 0.85)^T$

Local Adaptive MEM



Local Change Point Detection ▶ LCP

Example: Trading volumes aggregated over 1-min periods

- Scheme with $(K + 1) = 14$ intervals and fix i_0
- Assume $I_0 = 60\text{min.}$ is homogeneous
- H_0 : parameter homogeneity within $I_1 = 75\text{min.}$
 - ▶ Define $J_1 = I_1 \setminus I_0$ - observations from y_{i_0-75} up to y_{i_0-60}
 - ▶ For each $\tau \in J_1$ fit log likelihoods over $A_{1,\tau}$, $B_{1,\tau}$ and I_2
 - ▶ Find the largest likelihood ratio - T_{I_1, J_1}

