

Flexible copula-based vector MEM

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Liquidity risk

Example 1

Problem: Let Q be an huge order influencing the market price; thus Q cannot be executed at the actual market price.

Solution: Split Q into smaller amounts $\delta_i \widehat{V}_i$, such that $Q = \sum_{i=1}^n \delta_i \widehat{V}_i$, with $\delta_i \in (0, 1)$, where \widehat{V}_i is the forecasted trading volumes, $i = 1, \dots, n$.



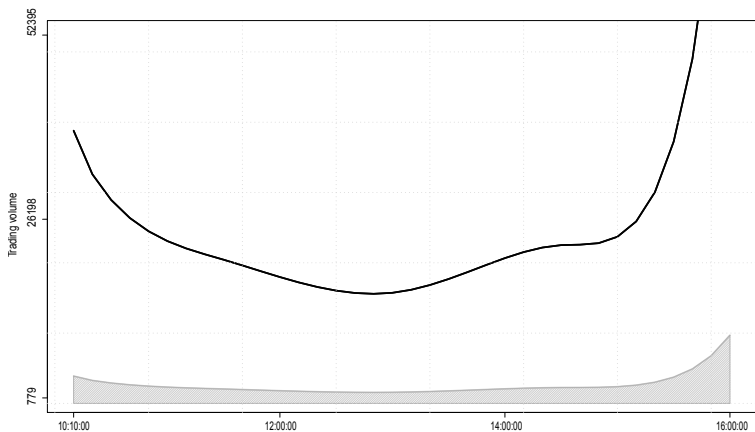


Figure 1: Trading volume of Skyworks Solutions Inc. (SWKS) at the 2009-08-08. Forecasted market volume; limited trade sizes with $\delta_i = 0.05$.



Multiplicative error model (MEM)

Engle (2002) proposed the MEM, i.e.

$$x_i = \mu_i \varepsilon_i,$$
$$\mu_i \stackrel{\text{def}}{=} E(x_i | \mathcal{F}_{i-1}; \xi),$$

with

- $\{x_i\}_{i=1}^n$ is a positive valued stochastic process,
- ε_i is iid with $E(\varepsilon_i) = 1$ and density $f : \mathbb{R}^+ \rightarrow \mathbb{R}^+$,
- Vector of parameters ξ ,
- μ_i measurable w.r.t. \mathcal{F}_{i-1} .



Research questions

- Joint modeling of volumes, price variations and trading intensities
- Basic ingredients are positive: *no* natural distribution on \mathbb{R}_+^d
- More accurate forecasts due to large information set
- Challenges:
 - ▶ Non-Gaussian dependencies
 - ▶ Two step estimation instead of computationally demanding full-ML estimation



Vector MEM (VMEM)

Cipollini and Gallo (2010) formalized the VMEM as

$$x_j = \mu_j \odot \varepsilon_j,$$
$$\mu_j \stackrel{\text{def}}{=} E(x_j | \mathcal{F}_{j-1}; \xi) = \underbrace{\omega}_{(d \times 1)} + \underbrace{A}_{(d \times d)} x_{j-1} + \underbrace{B}_{(d \times d)} \mu_{j-1},$$

where “ \odot ” is the componentwise Hadamard product, $\varepsilon_{ij} \geq 0 \forall i, j = 1, \dots, d$, $\xi = (\omega, A, B)$ and $\varepsilon_j | \mathcal{F}_{j-1} \sim D(1_d, \Sigma)$.

Example 2

$$\varepsilon_{ij} \sim \text{Exp}(1) = F \text{ and } D = C\{F(x_{i1}), \dots, F(x_{id}); \theta\}$$



Outline

1. Motivation ✓
2. Estimation
3. HAC
4. Data
5. Preliminary estimation results
6. Summary

Full ML estimation

From VMEM has x_i a conditional density

$$f_x(x_{i1}, \dots, x_{id} | \mathcal{F}_{i-1}; \theta, \xi, \alpha) = c\{F_1(\varepsilon_{i1}; \alpha_1), \dots, F_d(\varepsilon_{id}; \alpha_d); \theta\} \\ \times \prod_{j=1}^d \frac{f_j(\varepsilon_{ij}; \alpha_j)}{\mu_{ij}(\xi_j)},$$

for $\varepsilon_{ij} \stackrel{\text{def}}{=} \varepsilon_{ij}(\xi_j)$, $i = 1, \dots, n$ and copula density c . Then, the log-likelihood is given through

$$\ell(\theta, \xi, \alpha | \mathcal{F}_{i-1}) = \sum_{i=1}^n \sum_{j=1}^d [\log\{\varepsilon_{ij} f_j(\varepsilon_{ij}; \alpha_j)\} - \log x_{ij}] \\ + \sum_{i=1}^n \log c\{F_1(\varepsilon_{i1}; \alpha_1), \dots, F_d(\varepsilon_{id}; \alpha_d); \theta\}.$$



Advantage:

- Efficient and asymptotically unbiased estimates

Disadvantages:

- If C is flexible structured copula, ℓ has to be optimized for each possible structure
- Curse of dimensionality
 - ▶ $2d^2 + 4d - 1$ parameters to estimate (assuming 2 parameters for each margin and $d - 1$ for the copula)
- Parameter restrictions to ensure stationarity and positivity
- Computationally intensive
- ...



Multistage procedure

Similar to Chen and Fan (2006)

1. Estimate $\mu_i(\xi) = \{\mu_{i1}(\xi_1), \dots, \mu_{id}(\xi_d)\}^T$, $i = 1, \dots, n$
2. Extract the residuals $\hat{\varepsilon}_{i1}, \dots, \hat{\varepsilon}_{id}$ and determine the edfs \hat{F}_j
3. $\hat{\theta} = \arg \max_{\theta} \sum_{i=1}^n \log c \left\{ \hat{F}_1(\hat{\varepsilon}_{i1}), \dots, \hat{F}_d(\hat{\varepsilon}_{id}); \theta \right\}$



Advantages:

- ▣ Separate estimation of $\mu_i(\xi)$ and $C(\cdot; \theta)$
- ▣ Applicable for high dimensional time series, conditional on the copula

Disadvantages:

- ▣ Matrices A and B should be diagonal
- ▣ Estimation equation by equation is inefficient otherwise

Consequence: no real VMEM and no superior forecasting properties



Maximization by parts

Set $\delta = (\xi, \alpha)^\top$ and rewrite the log-likelihood

$$\begin{aligned}
 \ell(\theta, \delta | \mathcal{F}_{i-1}) &= \overbrace{\sum_{i=1}^n \sum_{j=1}^d [\log \{\varepsilon_{ij} f_j(\varepsilon_{ij}; \alpha_j)\} - \log x_{ij}]}^{\text{working}} \\
 &\quad + \underbrace{\sum_{i=1}^n \log c \{F_1(\varepsilon_{i1}; \alpha_1), \dots, F_d(\varepsilon_{id}; \alpha_d); \theta\}}_{\text{error}}
 \end{aligned}$$

in two parts

$$\ell(\delta, \theta) = \ell_w(\delta) + \ell_e(\delta, \theta).$$



Shortcuts: $\ell_w^\delta = \partial \ell_w / \partial \delta$, $\ell_e^\delta = \partial \ell_e / \partial \delta$ and $\ell_e^\theta = \partial \ell_e / \partial \theta$

Algorithm (Song et al. (2005))

For $k = 1$:

1. Solve $\ell_w^\delta(\delta) = \mathbf{0}$ for $\hat{\delta}^1$
2. Solve $\ell_e^\theta(\hat{\delta}^1, \theta) = \mathbf{0}$ for $\hat{\theta}^1$

For $k = 2, 3, \dots$:

1. Solve $\ell_w^\delta(\delta) = -\ell_e^\delta(\hat{\delta}^{k-1}, \hat{\theta}^{k-1})$ for $\hat{\delta}^k$
2. Solve $\ell_e^\theta(\hat{\delta}^{k-1}, \theta) = \mathbf{0}$ for $\hat{\theta}^k$



HAC

Example 3

3-dimensional fully nested HAC with generators ϕ and ψ :

$$C(u_1, u_2, u_3) = \phi \left\{ \phi^{-1} \circ \psi \left\{ \psi^{-1}(u_1) + \psi^{-1}(u_2) \right\} + \phi^{-1}(u_3) \right\}$$

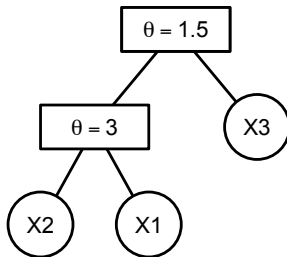


Figure 2: Structure of 3-dim HAC



Estimation

Okhrin et al. (2011) proposed a multi-stage QML-estimation, i.e.

$$\left(\frac{\partial \ell_1}{\partial \theta_1}, \dots, \frac{\partial \ell_{d-1}}{\partial \theta_{d-1}} \right)^\top = \mathbf{0}_{d-1}$$

with quasi log-likelihood

$$\ell(\theta|\mathbf{X}) = n^{-1} \sum_{i=1}^n \log \left[c \left\{ \widehat{F}_1(x_{i1}), \dots, \widehat{F}_d(x_{id}), \theta \right\} \right].$$



Data

- Datasource: NASDAQ (LOBSTER)
- Stock: SWKS
- 2nd of November 2009 to 31th of December 2009
- Series is cleaned and adjusted for deterministic intraday patterns
- Aggregation level:
 - ▶ 10 min
 - ▶ 1331 observations



Variables

High-frequency data sets provide information about

$$\{(p_j, s_j, t_j)\}_{j=1}^{n^*},$$

with

- Price series p_j ,
- Amount of traded shares s_j ,
- Time stamps t_j ,
- Number of observation within a day n^* .



Construct the lower frequent equi-distant series

$$\{(HL_i, Vol_i, NT_i, t_i)\}_{i=1}^n,$$

with

- $n < n^*$
- $HL_i = \max \{p_j | t_j \in (t_{i-1}, t_i]\} - \min \{p_j | t_j \in (t_{i-1}, t_i]\},$
- $NT_i = \# \{t_j | t_j \in (t_{i-1}, t_i]\}$ and
- $Vol_i = NT_i^{-1} \sum_{j \in (t_{i-1}, t_i]} s_j.$



Multistage estimation

$$\mu_i = c + ax_{i-1} + b\mu_{i-1}$$

	HL	NT	Vol
c	0.02	0.02	0.18
a	0.11	0.20	0.28
b	0.87	0.77	0.56
α_1	-	0.94	0.50
α_2	-	5.00	11.74
p_{LB10}	0.03	0.01	0.10
p_{LB20}	0.09	0.17	0.40
p_{LB30}	0.15	0.06	0.55

Table 1: Estimated parameters and p -values of the Ljung-Box (LB) test; sample size: 1331. Assumptions: $\varepsilon_{HL} \sim Exp(1)$, $\varepsilon_{NT} \sim Ga(\alpha_1, \alpha_2)$ and $\varepsilon_{Vol} \sim Ga(\alpha_1, \alpha_2)$.



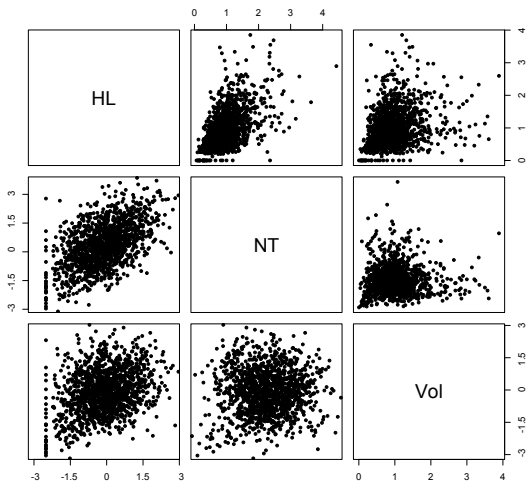


Figure 3: Scatterplot of the residuals.



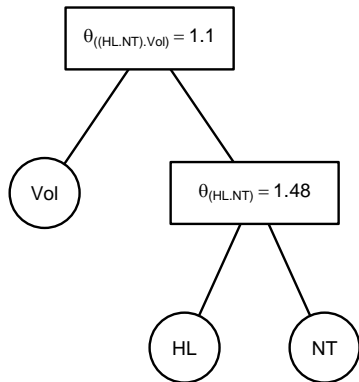


Figure 4: Estimated HAC for the residuals of the univariate processes.



Full ML

$$\mu_i = \begin{pmatrix} 0.22 \\ 0.22 \\ 0.24 \end{pmatrix} + \begin{pmatrix} 0.29 & 0.08 & 0.07 \\ 0.19 & 0.08 & 0.07 \\ 0.09 & 0.24 & 0.22 \end{pmatrix} x_{i-1} + \begin{pmatrix} 0.33 & 0.16 & 0.07 \\ 0.01 & 0.04 & 0.01 \\ 0.00 & 0.00 & 0.15 \end{pmatrix} \mu_{i-1}$$

Structure	log-lik
$((\text{HL.NT})_{1.85} \text{Vol})_{1.27}$	-2740
$((\text{HL.Vol})_{1.49} \text{NT})_{1.43}$	-2828
$((\text{NT.Vol})_{1.44} \text{HL})_{1.44}$	-2838

With $\alpha_1 = (1.50, 2.73, 3.68)^\top$, $\alpha_2 = (2.98, 2.07, 1.04)^\top$ for $(\text{HL}, \text{NT}, \text{Vol})^\top$.



Conclusion

- Aim at forecasting economic variables
- Two step estimation for VMEM
- Benefit from well developed univariate models and modern copulas

To do:

- Maximization by parts
- Time varying relationships
- Penalization of the parameter space's dimensionality



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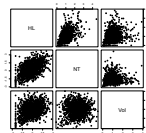
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