

Robust Inference for Quantile Predictive Regressions with Persistent Predictors

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Predictability of stock returns

- Test efficient market hypothesis
- Check policy effectiveness
- Look for potential investment opportunities
- Foresee financial crisis



Mean Predictive Regression

$$\begin{aligned}y_t &= \beta_0 + \beta_1 x_{t-1} + u_{0t}, \quad x_t = R_T x_{t-1} + u_{xt}, \\ R_T &= 1 + c/T^\alpha, \quad t = 1, \dots, T\end{aligned}\tag{1}$$

- y_t - stock returns
- x_{t-1} - predictors, e.g. monthly dividend-price ratio (DP), earnings-price ratio (EP), or macroeconomic variables etc.
- highly persistent and even non-stationary predictors, R_T is close to 1
- test the predictability of stock return, $H_0 : \beta_1 = 0$



Two Issues due to Persistent Predictors

- Embedded endogeneity
 - ▶ x_{t-1} and u_{0t} may be correlated
 - ▶ OLS estimator and its standard error are biased in finite sample
 - ▶ t -statistics tends to over-reject H_0

- Unknown degree of persistency
 - ▶ non-pivotal statistics, depending on nuisance parameter
 - ▶ robust inference theory across different persistence categories is required

▶ persistence categories



Corrections on the Issues

- Embedded endogeneity
 - ▶ first-order bias-correction estimator (Stambaugh, 1999)
 - ▶ linear projection of u_{0t} onto u_{xt} (Amihud et al., 2009; Cai and Wang, 2014)

- Unknown degree of persistency
 - ▶ approximations to finite-sample distributions, construct confidence intervals by Bonferroni's method (Cavanagh et al., 1995; Campbell and Yogo, 2006)
 - ▶ IVX filtering (Magdalinos and Phillips, 2009):
$$\tilde{z}_t = R_{TZ}\tilde{z}_{t-1} + \Delta x_t, R_{TZ} = 1 + \frac{c_z}{T^\delta}$$
 - ▶ include one more redundant lag term x_{t-2} (Ren et al., 2015)



Quantile Predictive Regression

□ Advantages

- ▶ examine the predicability under the entire conditional distribution
- ▶ skewed and heavy tailed errors in financial asset returns
- ▶ VaR in risk management

□ Existing works

- ▶ unit root predictor (Xiao, 2009) and nearly integrated case (Maynard et al., 2011)
- ▶ IVX-QR (Lee, 2016)



Our Contributions

- New two-step estimation procedure to deal with the two issues in turn
 - ▶ remove the endogeneity in the first step
 - ▶ construct a pivotal statistic in the second step
- Consistency and asymptotic normality of the estimator
- Robust inference theory across different persistence types to perform the tests



Outline

1. Motivation ✓
2. Model and Estimation
3. Inference Theory
4. Monte Carlo Simulation
5. Empirical Applications

Model Setting

- Univariate predictive regression model

$$y_t = \beta_0 + \beta_1 u_{xt} + \beta_2 x_{t-1} + v_t, t = 1, \dots, T$$
$$E(v_t | \mathcal{F}_{t-1}) = 0, E(v_t^2 | \mathcal{F}_{t-1}) = \sigma_v^2 \quad (2)$$

where $\mathcal{F}_t = \{(v_j, u_{xj})^\top, j \leq t\}$ is a natural filtration.

- Persistent predictor

$$x_t = R_T x_{t-1} + u_{xt}$$
$$R_T = 1 + c/T^\alpha \quad (3)$$

with $u_{xt} = \sum_{j=0}^{\infty} F_{xj} \varepsilon_{t-j}$, $\varepsilon_t \sim \text{m.d.s.}(0, \sigma_\varepsilon^2)$, $E|\varepsilon_1|^{2+\nu} < \infty$, $\forall \nu > 0$,
 $F_{x0} = 1$, $\sum_{j=0}^{\infty} j |F_{xj}| < \infty$, $F_x(z) = \sum_{j=0}^{\infty} F_{xj} z^j$, $F_x(1) = \sum_{j=0}^{\infty} F_{xj} > 0$



Two Step Quantile Regression - 2SQR(1)

Quantile Predictive Regression

$$Q_{y_t}(\tau | \mathcal{F}_{t-1}) = \beta_{0\tau} + \beta_{1\tau} u_{xt} + \beta_{2\tau} x_{t-1}, \tau \in (0, 1)$$

- ▶ step1:

$$(\hat{\beta}_{0\tau}, \hat{\beta}_{1\tau}, \hat{\beta}_{2\tau})^\top = \arg \min_{\beta_0, \beta_1, \beta_2} \sum_{t=2}^T \rho_\tau (y_t - \beta_0 - \beta_1 x_t - \beta_2 x_{t-1}),$$

where $\rho_\tau(e) = e[\tau - \mathbf{I}\{e < 0\}]$.

- ▶ step2:

$$(\hat{\beta}_{2\tau}, \hat{\beta}_{3\tau})^\top = \arg \min_{\beta_2, \beta_3} \sum_{t=3}^T \rho_\tau (y_{t\tau} - \beta_2 x_{t-1} - \beta_3 x_{t-2}),$$

with $y_{t\tau} = y_t - \hat{\beta}_{0\tau} - \hat{\beta}_{1\tau} \hat{u}_{xt}$, and \hat{u}_{xt} are the errors from OLS on (3).



Two Step Quantile Regression - 2SQR(2)

Quantile Predictive Regression

$$Q_{y_t}(\tau | \mathcal{F}_{t-1}) = \beta_{0\tau} + \beta_{1\tau} u_{xt} + \beta_{2\tau} x_{t-1}, \tau \in (0, 1)$$

- step1:

$$(\hat{\beta}_{0\tau}, \hat{\beta}_{1\tau}, \hat{\beta}_{2\tau})^\top = \arg \min_{\beta_0, \beta_1, \beta_2} \sum_{t=2}^T \rho_\tau (y_t - \beta_0 - \beta_1 x_t - \beta_2 x_{t-1}),$$

where $\rho_\tau(e) = e[\tau - \mathbf{I}\{e < 0\}]$.

- step2:

$$\hat{\beta}_{2\tau} = \arg \min_{\beta_2} \sum_{t=2}^T \rho_\tau (y_{t\tau} - \beta_2 x_{t-1}),$$

with $y_{t\tau} = y_t - \hat{\beta}_{0\tau} - \hat{\beta}_{1\tau} \hat{u}_{xt}$, and \hat{u}_{xt} are the errors from OLS on (3).

Asymptotic Distributions

Theorem 1: Asymptotic distribution of $\hat{\beta}_{2\tau}$ by 2SQR(1)

$$\sqrt{T} \left(\hat{\beta}_{2\tau} - \beta_{2\tau} \right) \xrightarrow{\mathcal{L}} \frac{\sqrt{\tau(1-\tau)}}{f_{v\tau}(0)} \mathbf{N} \left(0, \sigma_{xx}^{-1} \right),$$

where $\sigma_{xx} = \sum_{h=-\infty}^{\infty} \mathbf{E} (u_{xt} u_{xt-h})$, $f_{v\tau}$ is the pdf of the residuals in the 2nd step regression.



Test Statistics

Theorem 2: t -statistics of $\hat{\beta}_{2\tau}$ by 2SQR(1)

Under $H_0 : \beta_{2\tau} = 0$,

$$t = \frac{\sqrt{T} \hat{f}_{v\tau}(0) \hat{\beta}_{2\tau} \hat{\sigma}_{xx}}{\sqrt{\tau(1-\tau)}} \xrightarrow{\mathcal{L}} N(0, 1),$$

where $\hat{f}_{v\tau}(0)$ and $\hat{\sigma}_{xx}$ are consistent estimators of $f_{v\tau}(0)$ and σ_{xx} .

Test Statistics

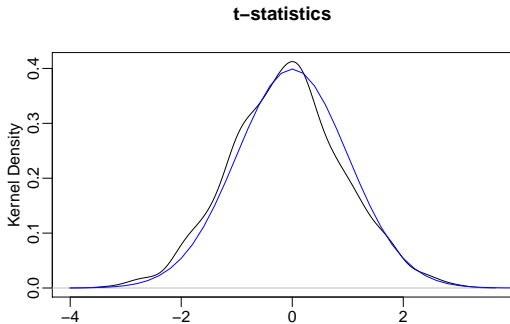



Figure 1: Kernel density of the t -statistics and [standard normal pdf](#), with $R_T = 0.99$, $\tau = 0.1$, $\text{cov}(u_{xt}, u_{0t}) = 0.95$, normal errors.

 BQPR_statdensity

Asymptotic Distributions

Theorem 3: Asymptotic distribution of $\hat{\beta}_{2\tau}$ by 2SQR(2)

$$D_T \left(\hat{\beta}_{2\tau} - \beta_{2\tau} \right) \xrightarrow{\mathcal{L}} \begin{cases} \frac{\sqrt{\tau(1-\tau)}}{f_{v\tau}(0)} N(0, \sigma_{xx}^{-1}), & \text{for } I0; \\ \frac{\sqrt{\tau(1-\tau)}}{f_{v\tau}(0)} N(0, \omega_{xx}^{-1}), & \text{for } MI; \\ \frac{\sqrt{\tau(1-\tau)}}{f_{v\tau}(0)} MN \left(0, \left\{ \int_0^\infty \bar{J}_x^c(r)^2 dr \right\}^{-1} \right), & \text{for } NI \text{ or } I1; \\ \frac{\sqrt{\tau(1-\tau)}}{f_{v\tau}(0)} MN(0, \tilde{\omega}_{xx}^{-1}), & \text{for } ME; \end{cases}$$

where $\omega_{xx} = \int_0^\infty \exp(2rc) \sigma_{xx} dr$, $\tilde{\omega}_{xx} = \int_0^\infty \exp(-2rc) Y_c^2 dr$, $Y_c \equiv N(0, \omega_{xx})$,
 $J_x^c(r) = \int_0^r \exp\{(r-s)c\} dB_x(s)$, $\frac{1}{\sqrt{T}} \sum_{j=1}^{\lfloor Ts \rfloor} u_{xj} \xrightarrow{\mathcal{L}} B_x(s)$,

$$\bar{J}_x^c(r) = J_x^c(r) - \int_0^1 J_x^c(r) dr, \text{ and } D_T = \begin{cases} \sqrt{T}, & \text{for } I0; \\ T^{(1+\alpha)/2}, & \text{for } MI; \\ T, & \text{for } NI \text{ or } I1; \\ T^\alpha R_T^T, & \text{for } ME. \end{cases}$$

Test Statistics

Theorem 4: self normalized statistics of $\hat{\beta}_{2\tau}$ by 2SQR(2)

Under $H_0 : \beta_{2\tau} = 0$,

$$\frac{\sqrt{T} \hat{f}_{v\tau}(0) \hat{\beta}_{2\tau} \sqrt{\sum_{t=1}^{T-1} x_t^2 - \frac{1}{T-1} \left(\sum_{t=1}^{T-1} x_t \right)^2}}{\sqrt{\tau(1-\tau)}} \xrightarrow{\mathcal{L}} N(0, 1),$$

where $\hat{f}_{v\tau}(0)$ is a consistent estimators of $f_{v\tau}(0)$.

Note: Theorem 2 and Theorem 4 can be extended to Wald statistics following χ^2 distribution with multiple predictors.

Test Statistics

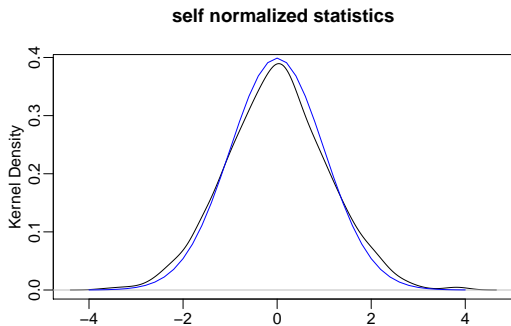



Figure 2: Kernel density of the self normalized statistics and **standard normal pdf**, with $R_T = 0.95$, $\tau = 0.1$, $\text{cov}(u_{xt}, u_{0t}) = 0.3$, normal errors.

 BQPR_statdensity

Power Performance

One NI predictor:

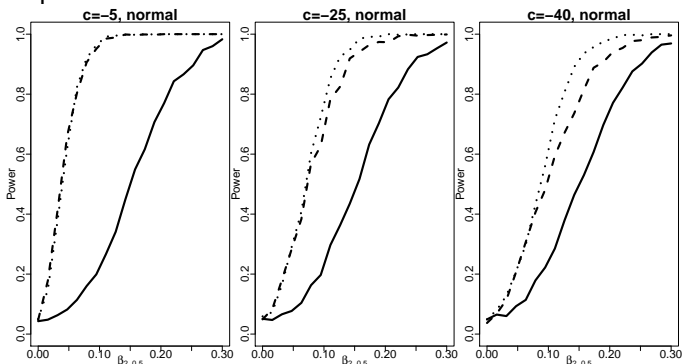


Figure 3: Power performance with one predictor, $T = 250$, $\tau = 0.5$, $\beta_1 = 0.3$, normal errors. Solid - 2SQR(1), dotted - 2SQR(2), dashed - IVX-QR.

 BQPR_power

Power Performance

One NI predictor:

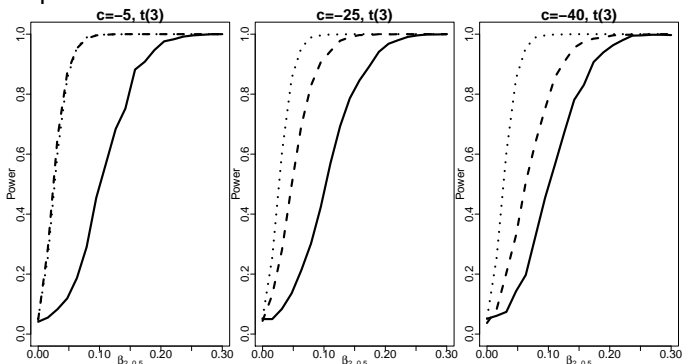


Figure 4: Power performance with one predictor, $T = 250$, $\tau = 0.5$, $\beta_1 = 0.3$, $t(3)$ errors. Solid - 2SQR(1), dotted - 2SQR(2), dashed - IVX-QR.

 BQPR_power

Size Performance

With one predictor in more cases:

τ	0.05	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	0.95
2SQR(1)											
IO: $c = -0.1, r = 0$	6.8	4.7	4.9	4.0	4.2	3.3	3.3	5.2	4.5	6.4	6.6
MI: $c = -0.1, r = 0.9$	6.2	5.0	3.9	4.3	3.6	5.4	4.6	3.8	5.0	5.3	7.8
NI: $c = -0.6, r = 1$	7.1	5.5	4.6	4.9	4.2	4.2	4.2	4.1	5.1	4.5	6.8
I1: $c = 0, r = 1$	6.9	6.3	4.8	5.0	4.6	3.9	3.9	4.3	4.9	5.2	6.4
ME: $c = 0.3, r = 0.6$	5.6	5.3	5.7	6.0	4.5	3.5	3.5	3.4	5.5	6.1	5.1
2SQR(2)											
IO: $c = -0.1, r = 0$	8.1	7.5	5.2	4.5	4.8	7.0	4.7	6.2	7.1	7.4	7.7
MI: $c = -0.1, r = 0.9$	8.4	6.3	7.3	7.1	4.9	5.8	4.9	6.0	6.2	7.2	10.5
NI: $c = -0.6, r = 1$	8.5	8.7	6.1	5.0	5.4	4.4	3.8	4.6	5.3	6.5	7.7
I1: $c = 0, r = 1$	7.5	6.5	5.5	6.0	3.9	6.0	4.8	5.1	5.2	7.9	8.0
ME: $c = 0.3, r = 0.6$	8.0	7.0	6.8	6.5	5.8	6.0	5.1	4.3	7.2	6.6	8.1
IVX-QR											
IO: $c = -0.1, r = 0$	7.8	7.0	3.6	4.6	4.3	4.2	3.6	4.9	4.6	5.4	7.5
MI: $c = -0.1, r = 0.9$	16.9	12.8	9.5	8.2	9.3	8.1	7.4	8.0	10.6	13.0	15.8
NI: $c = -0.6, r = 1,$	15.1	11.5	8.4	7.0	8.9	7.8	7.7	9.4	8.1	12.4	13.7
I1: $c = 0, r = 1,$	16.1	13.3	10.2	9.8	9.1	5.4	7.4	10.0	9.8	14.5	15.0
ME: $c = 0.3, r = 0.6$	18.4	15.3	12.9	12.6	12.7	13.0	13.6	12.7	13.8	15.1	17.8

Table 1: Size performance (%) with normal errors and $\beta_1 = 0.3$.

 BQPR_size



Size Performance

With one predictor in more cases:

τ	0.05	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	0.95
2SQR(1)											
IO: $c = -0.1, r = 0$	8.4	5.0	5.6	3.6	3.6	4.6	3.6	6.3	6.1	6.8	7.8
MI: $c = -0.1, r = 0.9$	8.0	7.3	5.6	4.6	5.3	2.7	4.2	5.2	6.2	7.5	7.0
NI: $c = -0.6, r = 1$	7.0	7.0	4.8	5.5	4.3	3.8	4.2	4.8	6.2	6.7	7.7
I1: $c = 0, r = 1$	8.4	5.0	5.6	3.6	3.6	4.6	3.6	6.3	6.1	6.8	7.8
ME: $c = 0.3, r = 0.6$	7.8	8.0	6.2	3.9	4.4	3.7	4.2	5.6	6.6	8.1	7.9
2SQR(2)											
IO: $c = -0.1, r = 0$	10.8	6.7	6.9	7.2	6.1	5.8	6.4	6.1	6.8	9.6	10.3
MI: $c = -0.1, r = 0.9$	10.3	7.9	6.6	5.8	4.8	4.8	3.6	4.5	5.8	8.6	10.7
NI: $c = -0.6, r = 1$	9.4	7.4	6.4	6.4	5.3	5.9	3.9	4.1	5.6	8.1	9.8
I1: $c = 0, r = 1$	11.6	9.3	6.3	5.4	5.8	5.8	5.0	6.1	8.2	8.0	11.9
ME: $c = 0.3, r = 0.6$	9.5	9.8	6.6	6.3	5.2	5.1	6.1	5.9	7.3	8.1	12.8
IVX-QR											
IO: $c = -0.1, r = 0$	9.9	8.4	6.1	4.6	5.3	3.0	3.8	4.0	5.9	7.2	9.2
MI: $c = -0.1, r = 0.9$	16.8	14.3	12.6	10.2	10.9	7.7	9.9	11.6	11.0	14.2	16.3
NI: $c = -0.6, r = 1$	15.4	13.4	11.5	9.6	7.7	8.1	9.7	10.1	10.2	13.4	14.8
I1: $c = 0, r = 1$	17.9	15.3	12.1	10.1	8.6	9.7	10.6	11.4	12.3	15.4	16.6
ME: $c = 0.3, r = 0.6$	21.8	17.0	15.1	12.3	14.8	10.1	8.9	13.7	17.1	19.1	19.6

Table 2: Size performance (%) with $t(3)$ errors and $\beta_1 = 0.3$.

 BQPR_size




Predictability of US Stock Returns

- Same dataset as Lee (2016): monthly data from 1927:01 to 2005:12
- Eight persistent predictors: d/p , e/p , b/m , $ntis$, d/e , tbl , dfy , tms
- Predict S&P 500 index including dividends minus one month Treasury bill rate




Predictability of US Stock Returns by 2SQR(1)

τ	0.05	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	0.95
d/p	0.72	0.96	0.34	0.99	0.60	0.21	0.06	0.00	0.00	0.00	0.00
e/p	0.39	0.01	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
b/m	0.34	0.92	0.95	0.99	0.99	0.89	0.61	0.06	0.09	0.01	0.20
ntis	0.62	0.27	0.71	0.45	0.35	0.95	0.95	0.79	0.19	0.23	0.16
d/e	0.90	0.80	0.73	0.60	0.92	0.98	0.93	0.89	0.85	0.59	0.99
tbl	0.55	0.58	0.21	0.24	0.29	0.69	0.57	0.17	0.09	0.11	0.08
dfy	0.15	0.11	0.10	0.04	0.60	0.78	0.00	0.04	0.00	0.02	0.00
tms	0.99	0.85	0.10	0.17	0.12	0.24	0.65	0.68	0.46	0.15	0.37

Table 3: p -values in hypothesis tests on univariate quantile predictor. Values are in bold if they are less or equal to 0.05.  BQPR_predictability

Predictability of US Stock Returns by 2SQR(2)


τ	0.05	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	0.95
d/p	0.01	0.00	0.00	0.00	0.00	0.10	0.36	0.85	0.75	0.14	0.00
e/p	0.00	0.01	0.23	0.13	0.00	0.00	0.00	0.00	0.01	0.63	0.21
b/m	0.41	0.40	0.29	0.94	0.22	0.00	0.00	0.00	0.00	0.03	0.59
ntis	0.04	0.02	0.01	0.07	0.02	0.01	0.13	0.17	0.30	0.73	0.93
d/e	0.00	0.00	0.00	0.03	0.50	0.72	0.95	0.09	0.02	0.00	0.00
tbl	0.41	0.76	0.44	0.15	0.01	0.03	0.01	0.00	0.00	0.01	0.01
dfy	0.00	0.00	0.00	0.00	0.02	0.46	0.01	0.00	0.00	0.00	0.00
tms	0.97	0.51	0.10	0.38	0.17	0.36	0.47	0.83	0.26	0.00	0.00

Table 4: p -values in hypothesis tests on univariate quantile predictor.  BQPR_predictability

Predictability of US Stock Returns by 2SQR(2)

τ	0.05	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	0.95
e/p, dfy	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.03	0.04	0.00
ntis, dfy	0.00	0.00	0.00	0.00	0.00	0.06	0.23	0.00	0.00	0.00	0.00
d/e, dfy	0.00	0.00	0.00	0.00	0.08	0.76	0.04	0.01	0.00	0.00	0.00


Table 5: p -values in hypothesis tests on bivariate quantile predictors.

 BQPR_predictability

VaR Backtesting

	Violation frequency
d/p	0.0787
e/p	0.0694
ntis	0.0586
d/e	0.0540
dfy	0.0602
e/p, dfy	0.0571
ntis, dfy	0.0694
d/e, dfy	0.0648

Table 6: Violation frequency in backtesting on the estimated $\text{VaR}_{0.05}$ (one step ahead) by 2SQR(2), in rolling window samples with window width = 300.

 BQPR_backtesting

VaR Backtesting

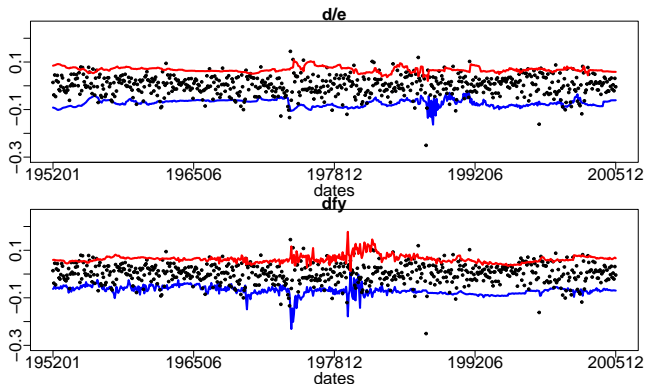



Figure 5: One step ahead estimated $VaR_{0.05}$, $VaR_{0.95}$ by 2SQR(2) with univariate predictor, and real observations.  BQPR_backtesting

VaR Backtesting

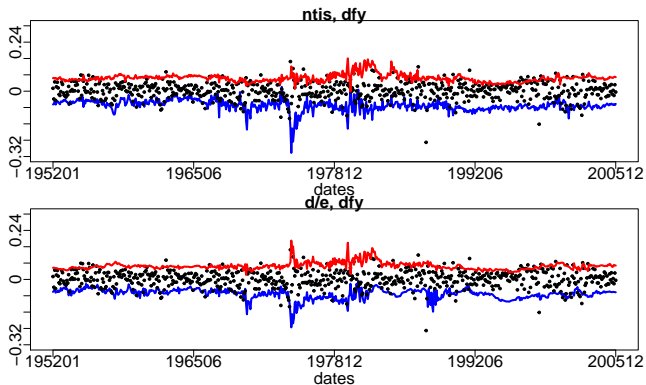



Figure 6: One step ahead estimated $VaR_{0.05}$, $VaR_{0.95}$ by 2SQR(2) with bivariate predictors, and real observations.

 BQPR_backtesting

Robust Inference for Quantile Predictive Regressions with Persistent Predictors

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Persistence Categories

- (I0) Stationary: $\alpha = 0, |1 + c| < 1$
- (MI) Mildly integrated: $0 < \alpha < 1, c < 0$
- (NI) Nearly integrated: $\alpha = 1, c < 0$
- (I1) Unit root: $c = 0$
- (ME) Mildly explosive: $0 < \alpha < 1, c > 0$

▶ Return

