

TEDAS - Tail Event Driven ASset allocation

Wolfgang Karl Härdle

David Lee Kuo Chuen

Sergey Nasekin

Alla Petukhina

Ladislaus von Bortkiewicz Chair of Statistics

C.A.S.E. – Center for Applied Statistics
and Economics

Humboldt–Universität zu Berlin

<http://lvb.wiwi.hu-berlin.de>

<http://www.case.hu-berlin.de>



S&P 500 Stocks

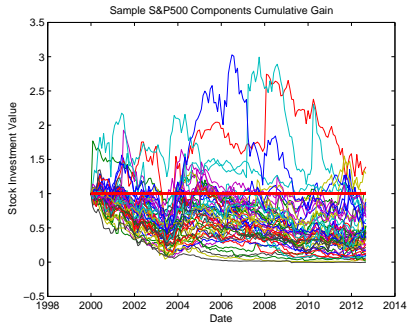


Figure 1: 50 random S&P 500 Sample Components' Cumulative Return: **94%** of stocks lost the value of the initial investment (**thick red line**)



Core & Satellites

Hedge funds and S&P500

- diversification - reduction of the portfolio risk
- construction - a more diverse universe of assets
- allocation - a higher risk-adjusted return.



Hedge Funds

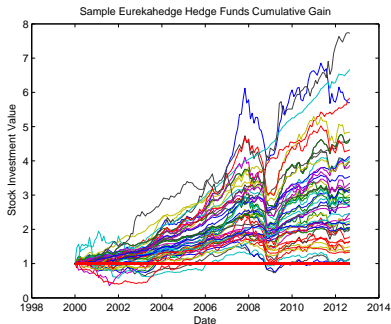


Figure 2: 50 EurekaHedge Hedge Funds Indices' Cumulative Return: **0%** of funds lost the value of the initial investment (**thick red line**)



Diversification

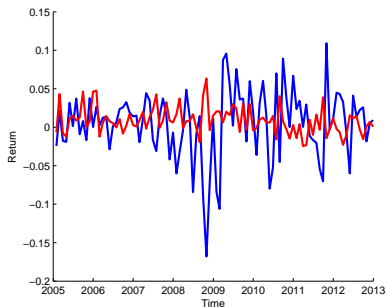


Figure 3: S&P 500 and Eureka Hedge North America Macro Hedge Fund Index monthly returns in 20050131-20121231



Traditional Assets/Hedge Funds

Hedge Funds	US	UK	SW	GER	JAP
Conv. arb.	0.10	0.08	0.07	0.10	-0.02
Dedic. sh. bias	-0.77	-0.53	-0.33	-0.46	-0.48
Fix. inc. arb.	0.10	0.13	0.01	0.08	-0.10
Glob. macro	0.30	0.19	0.10	0.27	-0.11
Man. fut.	-0.10	0.02	-0.09	-0.03	0.03

Table 1: Correlation statistics for traditional asset class and hedge funds' indices; based on monthly data Jan. 1994 - Aug. 2001; table from Lhabitant (2002, p.164)

▶ TLND Hedge Funds' Strategies

▶ More



Tail Risk

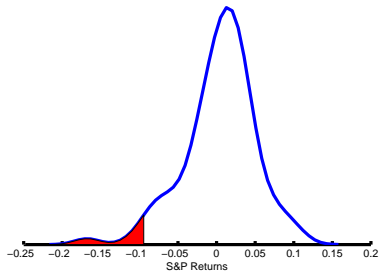


Figure 4: Estimated density of S&P 500 returns



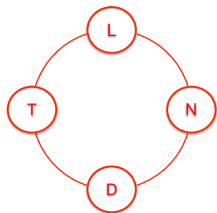
The TLND challenge

- Tail dependence
- Large universe: $p > n$
- Non normality
- Dynamics



TEDAS Objectives

- Hedge tail events
 - ▶ Quantile regression
 - ▶ Variable selection in high dimensions
- Improve Asset Allocation
 - ▶ Higher-order moments' optimization
 - ▶ Modelling of moments' dynamics



Outline

1. Motivation ✓
2. TEDAS framework
3. Empirical Application
4. Conclusions
5. Technical Details

Tail Events

- $Y \in \mathbb{R}^n$ core log-returns; $X \in \mathbb{R}^{n \times p}$ satellites' log-returns, $p > n$

-

$$q_\tau(x) \stackrel{\text{def}}{=} F_{Y|X}^{-1}(\tau) = x^\top \beta(\tau),$$

$$\beta(\tau) = \arg \min_{\beta \in \mathbb{R}^p} E_{Y|X} \rho_\tau\{Y - X^\top \beta\},$$

$$\rho_\tau(u) = u\{\tau - \mathbf{I}(u < 0)\}$$

- L_1 penalty $\lambda_n \|\hat{\omega}^\top \beta\|_1$ to nullify "excessive" coefficients; λ_n and $\hat{\omega}$ controlling penalization; constraining $\beta \leq 0$ yields ALQR [▶ Details](#)

$$\hat{\beta}_{\tau, \lambda_n}^{\text{adapt}} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_\tau(Y_i - X_i^\top \beta) + \lambda_n \|\hat{\omega}^\top \beta\|_1 \quad (1)$$



TEDAS Step 1

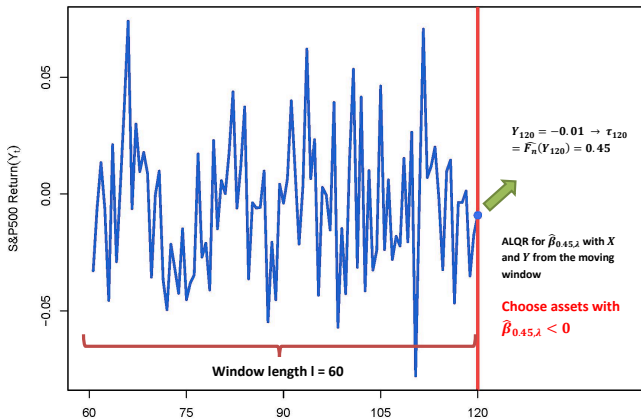
Initial wealth $W_0 = \$1$, $t = l, \dots, n$; $l = 60$ length of the moving window

□ Portfolio constituents' selection

1. determine core asset return Y_t , set $\tau_t = \widehat{F}_n(Y_t)$ ▶ Notation
2. ALQR for $\widehat{\beta}_{\tau_t, \lambda_n}$ using the observations $X \in \mathbb{R}^{t-l+1, \dots, t \times p}$, $Y \in \mathbb{R}^{t-l+1, \dots, t}$
3. if $Y_t < 0$, choose X_j , $j = 1, \dots, k < p$ with $\widehat{\beta}_{\tau_t, \lambda_n} < 0$; if $Y_t > 0$, choose X_j , $j = 1, \dots, k < p$ with $\widehat{\beta}_{\tau_t, \lambda_n} > 0$



TEDAS Step 1



TEDAS Step 2

□ Portfolio selection

1. apply TEDAS Gestalt to X_j , obtain $\hat{w}_t \in \mathbb{R}^k$
2. determine the realized portfolio wealth for $t + 1$,
$$\hat{X}_{t+1} \stackrel{\text{def}}{=} (X_{t+1,1}, \dots, X_{t+1,k})^\top: W_{t+1} = W_t(1 + \hat{w}_t^\top \hat{X}_{t+1})$$



TEDAS Example

1. Suppose $t = 86$ (Feb. 2007), $W_{86} = \$1.125$, accumulated wealth $W_{86} = \$1.429$, $Y_{86} = -1.85\% < 0$
2. $\hat{F}_n(Y_{86}) = 0.25$, so estimate $\hat{\beta}_{0.25} < 0$
3. ALQR on $X \in \mathbb{R}^{60 \times 163}$, $Y \in \mathbb{R}^{60}$ yields $\hat{\beta}_{0.25} = (-0.77, -1.12, -0.41)^\top$, *Latin American Arbitrage, North America Macro, Emerging Markets CTA/Managed Futures*
4. TEDAS CF-CVaR optimization [Details](#) yields $\hat{w}_{86} = (0.22, 0.16, 0.62)^\top$; $\hat{X}_{87} = (0.38\%, 0.45\%, 0.76\%)^\top$, $W_{87} = W_{86}(1 + \hat{w}_{86}^\top \hat{X}_{87}) = \1.438 (return of 0.62%) while $Y_{87} = -1.53\%$



Optimization strategies

TEDAS strategy	Model	Weights' optimization
TEDAS 1	NA	Equal weights
TEDAS 2	Time-Varying Details	Cornish-Fisher-CVaR optimization Details
TEDAS 3	conditional distribution	Expected utility optimization Details



Hedge funds

- Monthly data
 - ▶ Core asset (Y): S&P 500, Nikkei225, DAX 30, FTSE 100
 - ▶ Satellite assets (X): 164 Eureka hedge hedge funds indices
- Span: 20000131-20131031 (166 months)
- Source: Bloomberg



Benchmark Strategies

1. **RR**: dynamic risk-return optimization [▶ Details](#)
2. **PESS**: tail risk optimization [▶ Details](#)
3. S&P500 Buy-and-Hold



TEDAS with $Y = \text{S\&P 500}$

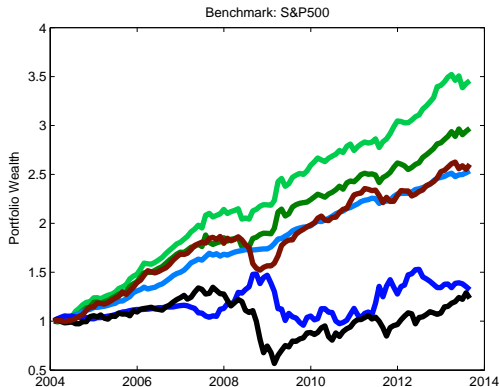



Figure 5: Cumulative portfolio wealth comparison: TEDAS 1, TEDAS 3, TEDAS 2, RR, PESS, S&P 500 buy & hold; X = hedge funds' indices' returns matrix  TEDASstrategies2



TEDAS with $Y = \text{Nikkei 225}$

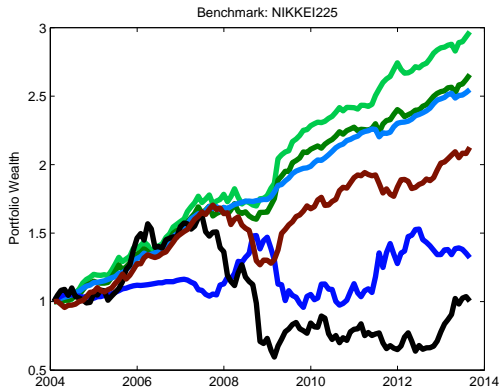



Figure 6: Cumulative portfolio wealth comparison: TEDAS 1, TEDAS 3, TEDAS 2, RR, PESS, Nikkei 225 buy & hold; X = hedge funds' indices' returns matrix  TEDASstrategies2



TEDAS with $Y = \text{FTSE 100}$

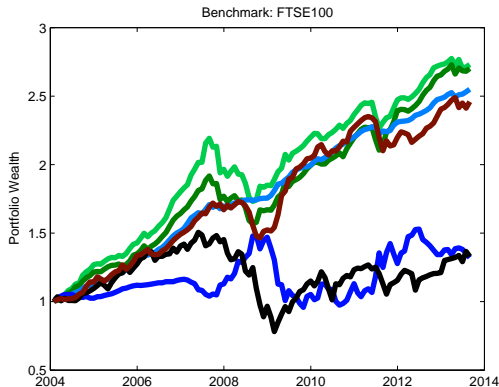



Figure 7: Cumulative portfolio wealth comparison: TEDAS 1, TEDAS 3, TEDAS 2, RR, PESS, FTSE100 buy & hold; X = hedge funds' indices' returns matrix  TEDASstrategies2



TEDAS with $Y = \text{DAX30}$

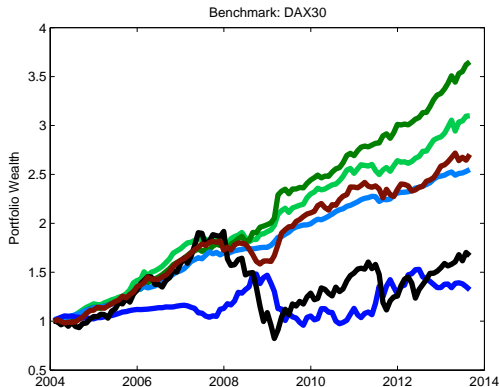



Figure 8: Cumulative portfolio wealth comparison: TEDAS 1, TEDAS 3, TEDAS 2, RR, PESS, DAX30 buy & hold; X = hedge funds' indices' returns matrix  TEDASstrategies2



Histograms of \hat{q}

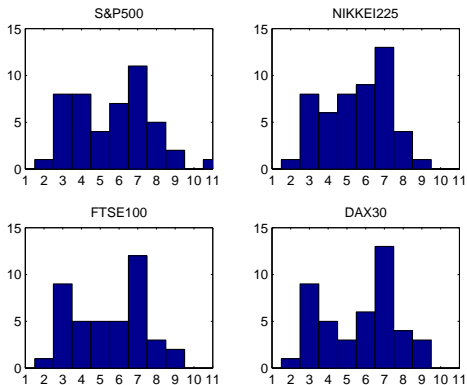


Figure 9: Frequency of the number of selected variables for 4 different Y



Selected Hedge Funds: S&P 500

Table 2: The selected hedge funds for S&P 500 benchmark

Top 5 influential hedge funds	Frequency
Latin American Onshore Fixed Income Hedge Fund Index	19
Emerging Markets Dual Approach Absolute Return Fund Index	18
Large North American Hedge Fund Index	15
Europe Macro Hedge Fund Index	15
Asia Macro Hedge Fund Index	14



Selected Hedge Funds: Nikkei 225

Table 3: The selected hedge funds for Nikkei 225 benchmark

Top 5 influential hedge funds	Frequency
Small Latin American Hedge Fund Index	22
Taiwan Hedge Fund Index	20
North America Top-Down Absolute Return Fund Index	16
Large North American Hedge Fund Index	15
Latin American Onshore Fixed Income Hedge Fund Index	15



Selected Hedge Funds: FTSE 100

Table 4: The selected hedge funds for FTSE 100 benchmark

Top 5 influential hedge funds	Frequency
Small Latin American Hedge Fund Index	15
Europe Macro Hedge Fund Index	15
Latin American Onshore Fixed Income Hedge Fund Index	14
Asia Pacific Top-Down Absolute Return Fund Index	14
Latin American Fixed Income Hedge Fund Index	13



Selected Hedge Funds: DAX 30

Table 5: The selected hedge funds for DAX 30 benchmark

Top 5 influential hedge funds	Frequency
Emerging Markets Dual Approach Absolute Return Fund Index	21
North America Macro Hedge Fund Index	19
Taiwan Hedge Fund Index	16
Asia CTA Hedge Fund Index	14
Europe Macro Hedge Fund Index	14



Dynamic Moment Parameters

Table 6: Parameter estimates in (14) and (15) at $t = 160$

	s_1	s_2	s_3	s_4	s_5	s_6
ω	0.12***	0.33***	0.40***	0.61***	0.60***	0.47***
α_1	0.72***	0.29***	0.47***	0.30***	0.33***	0.45***
β_1	0.13***	0.35***	0.12***	0.05**	0.07***	0.06
a_0	0.05	-0.02	-0.27	0.08***	0.11	-0.08
a_1^-	-0.33*	0.08	-0.55	-0.13***	-0.13	-0.69**
a_1^+	-0.43	0.22***	1.17	-0.46***	-0.33	-0.47***
a_2	1.45***	2.16***	0.96	2.14***	1.77***	1.10**
b_0	-5.03***	-1.96*	-5.53	-3.65***	-4.16**	-1.89***
b_1^-	0.70**	0.21	0.88	1.56	-1.53	0.51***
b_1^+	0.79***	-1.40***	0.87	0.89	-1.87	-0.79***
b_2	-11.67	-0.69	0.12	-0.64***	0.38	1.11***
$Factor_{LL}$	-193.87	-201.40	-215.14	-216.53	-218.57	-225.10
$Model_{LL}$	2394					

parameter estimates under the *NIG* distribution, for the log-returns of EurekaHedge 6 hedge funds (160 data monthly returns, 01.01.2000 - 30.04.2013)

the conditional variance of factors follows a GARCH(1,1) model

the conditional dynamics of skew and shape parameters is bounded by a logistic transformation

the *, ** and *** denote significance at the 10%, 5% and 1% levels, respectively



Evolution of ξ_t and ν_t

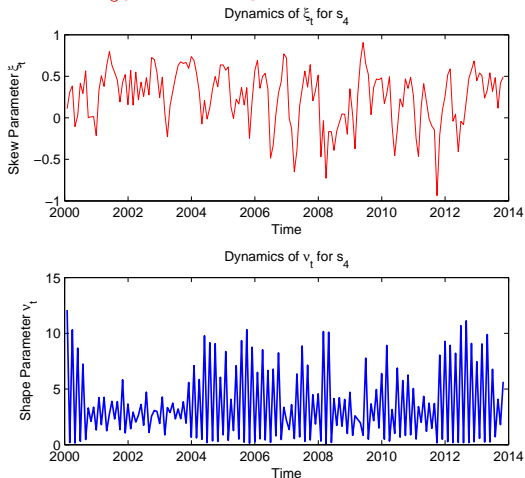


Figure 10: Evolution of the skew and shape parameters for s_4 in Table 6



Conditional Skewness and Kurtosis

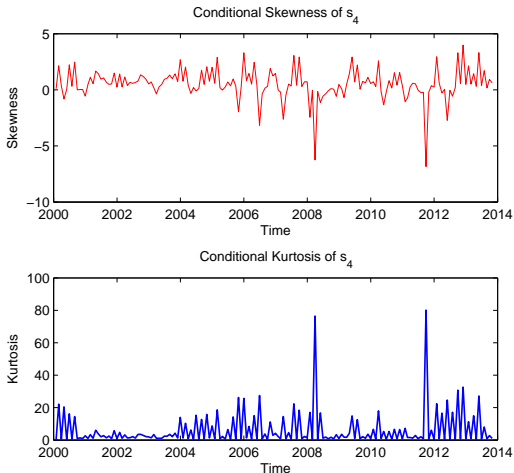


Figure 11: Evolution of the conditional skewness and kurtosis for s_4



Choice of τ

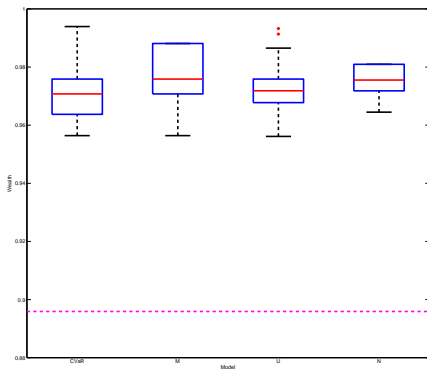


Figure 12: Next-day portfolio value for τ -values within 0.1 from τ_t under different models; DAX30 portfolio value (magenta); negative tail event, 5 assets



Choice of τ

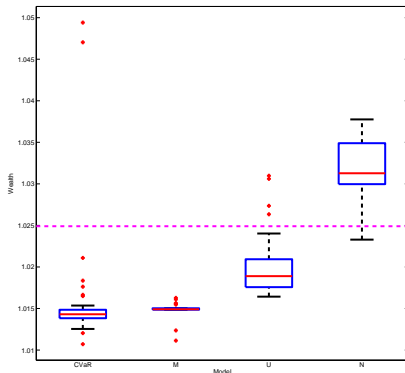


Figure 13: Next-day portfolio value for τ -values within 0.1 from τ_t under different models; DAX30 portfolio value (magenta); positive tail event, 11 assets



Choice of τ

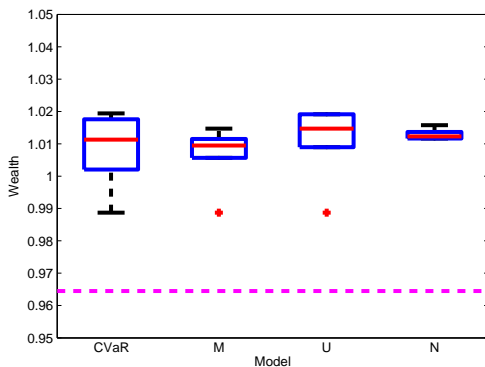


Figure 14: Next-day portfolio value for τ -values within 0.1 from τ_t under different models; FTSE100 portfolio value (magenta); negative tail event, 5 assets



Choice of τ

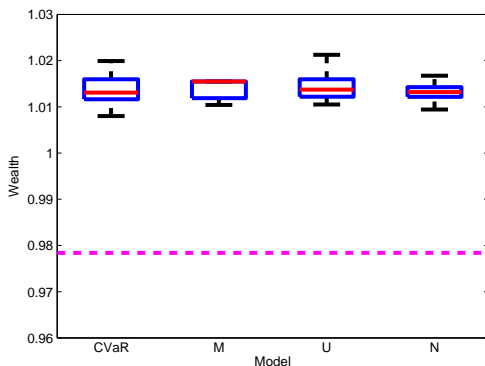


Figure 15: Next-day portfolio value for τ -values within 0.1 from τ_t under different models; FTSE100 portfolio value (magenta); negative tail event, 11 assets



Choice of τ

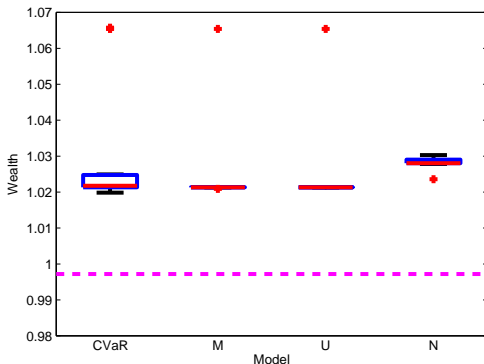


Figure 16: Next-day portfolio value for τ -values within 0.1 from τ_t under different models; S&P500 portfolio value (magenta); negative tail event, 3 assets



Choice of τ

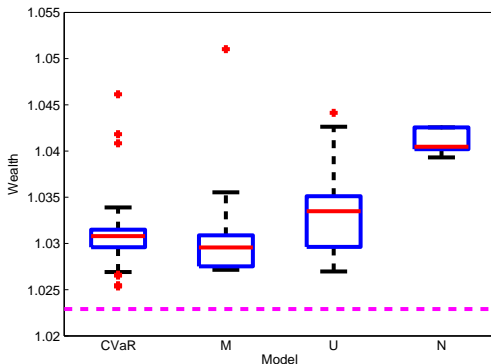


Figure 17: Next-day portfolio value for τ -values within 0.1 from τ_t under different models; S&P500 portfolio value (magenta); positive tail event, 9 assets



Conclusions: TLND challenge

- Lasso quantile regression captures **T**ail events
- Lasso asset selection resolves **L**arge dimensionality problem
- Higher-moment optimization for **N**on-normality
- **D**ynamic portfolio optimization through conditional distribution modelling
- TEDAS: out-of-sample performance superior to benchmark strategies



TEDAS - Tail Event Driven Asset Allocation

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Sergey Nasekin

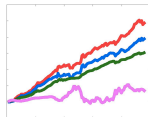
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Notation

□ $\hat{q}_\tau \stackrel{\text{def}}{=} \hat{F}_n^{-1}(\tau)$, with

$$\hat{F}_n(Y_t) \stackrel{\text{def}}{=} \int_{-\infty}^{Y_t} \hat{f}_n(u) du = \frac{1}{n} \sum_{i=1}^n H\left(\frac{Y_t - Y_i}{h}\right), \quad (2)$$

where $\hat{f}_n(Y_t) \stackrel{\text{def}}{=} (1/nh) \sum_{i=1}^n K\{(Y_t - Y_i)/h\}$,

$H(x) = \int_{-\infty}^x K(u) du$, $K(\cdot) = \varphi(\cdot)$;

Silverman (1986) rule-of-thumb:

$$h = 1.06sn^{-1/5}, \quad s \text{ sample standard deviation of } Y$$

□ $\hat{\beta}_{\tau, \lambda_n}$ are the estimated non-zero ALQR coefficients

▶ Back to "TEDAS Step 1"



Lasso Shrinkage

Linear model: $Y = X\beta + \varepsilon$; $Y \in \mathbb{R}^n$, $X \in \mathbb{R}^{n \times p}$, $\beta \in \mathbb{R}^p$, $\{\varepsilon_i\}_{i=1}^n$ i.i.d., independent of $\{X_i; i = 1, \dots, n\}$

The optimization problem for the lasso estimator:

$$\begin{aligned} \hat{\beta}^{\text{lasso}} &= \arg \min_{\beta \in \mathbb{R}^p} f(\beta) \\ &\text{subject to } g(\beta) \geq 0 \end{aligned} \tag{3}$$

where

$$\begin{aligned} f(\beta) &= \frac{1}{2} (y - X\beta)^\top (y - X\beta) \\ g(\beta) &= t - \|\beta\|_1 \end{aligned}$$

where t is the size constraint on $\|\beta\|_1$ [▶ Back to "Tail Events"](#)



Lasso Duality

If (1) is convex programming problem, then the Lagrangian is

$$L(\beta, \lambda) = f(\beta) - \lambda g(\beta).$$

and the primal-dual relationship is

$$\underbrace{\text{minimize}_{\beta} \sup_{\lambda \geq 0} L(\beta, \lambda)}_{\text{primal}} \geq \underbrace{\text{maximize}_{\lambda \geq 0} \inf_{\beta} L(\beta, \lambda)}_{\text{dual}}$$

Then the dual function $L^*(\lambda) = \inf_{\beta} L(\beta, \lambda)$ is

$$L^*(\lambda) = \frac{1}{2} y^T y - \frac{1}{2} \hat{\beta}^T X^T X \hat{\beta} - t \frac{(y - X \hat{\beta})^T X \hat{\beta}}{\|\hat{\beta}\|_1}$$

with $(y - X \hat{\beta})^T X \hat{\beta} / \|\hat{\beta}\|_1 = \lambda$ [▶ Back to "Tail Events"](#)



Paths of Lasso Coefficients

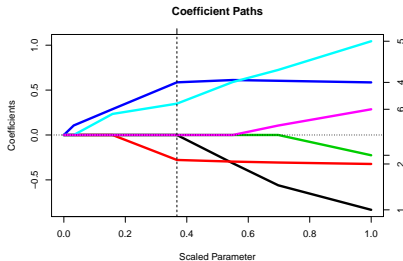


Figure 18: Lasso shrinkage of coefficients in the hedge funds dataset example (6 covariates were chosen for illustration); each curve represents a coefficient as a function of the scaled parameter $\hat{s} = t/\|\beta\|_1$; the dashed line represents the model selected by the BIC information criterion ($\hat{s} = 3.7$)

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Example of Lasso Geometry

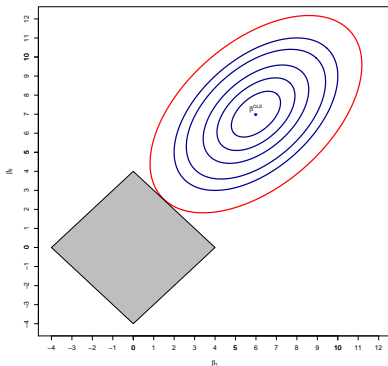



Figure 19: Contour plot of the residual sum of squares objective function centered at the OLS estimate $\hat{\beta}^{ols} = (6, 7)$ and the constraint region $\sum |\beta_j| \leq t$  MVAlassocontour



Quantile Regression

The loss $\rho_\tau(u) = u\{\tau - \mathbf{I}(u < 0)\}$ gives the (conditional) quantiles $F_{y|x}^{-1}(\tau) \stackrel{\text{def}}{=} q_\tau(x)$.

Minimize

$$\hat{\beta}_\tau = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_\tau(Y_i - X_i^\top \beta).$$

Re-write:

$$\underset{(\xi, \zeta) \in \mathbb{R}_+^{2n}}{\text{minimize}} \quad \left\{ \tau \mathbf{1}_n^\top \xi + (1 - \tau) \mathbf{1}_n^\top \zeta \mid X\beta + \xi - \zeta = Y \right\}$$

with ξ, ζ are vectors of "slack" variables [▶ Back to "Tail Events"](#)



Non-Positive (NP) Lasso-Penalized QR

The **lasso-penalized** QR problem with an additional non-positivity constraint takes the following form:

$$\begin{aligned} & \underset{(\xi, \zeta, \eta, \tilde{\beta}) \in \mathbb{R}_+^{2n+p} \times \mathbb{R}^p}{\text{minimize}} && \tau \mathbf{1}_n^\top \xi + (1 - \tau) \mathbf{1}_n^\top \zeta + \lambda \mathbf{1}_n^\top \eta \\ & \text{subject to} && \xi - \zeta = Y + X\tilde{\beta}, \\ & && \xi \geq \mathbf{0}, \\ & && \zeta \geq \mathbf{0}, \\ & && \eta \geq \tilde{\beta}, \\ & && \eta \geq -\tilde{\beta}, \\ & && \tilde{\beta} \geq \mathbf{0}, \quad \tilde{\beta} \stackrel{\text{def}}{=} -\beta \end{aligned} \tag{4}$$

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Solution

Transform into matrix (I_p is $p \times p$ identity matrix; $E_{p \times n} = \begin{pmatrix} I_n \\ 0 \end{pmatrix}$):

$$\begin{aligned} & \text{minimize} && c^\top x \\ & \text{subject to} && Ax = b, \quad Bx \leq 0 \end{aligned}$$

where $A = \begin{pmatrix} I_n & -I_n & 0 & X \end{pmatrix}$, $b = Y$, $x = \begin{pmatrix} \xi & \zeta & \eta & \beta \end{pmatrix}^\top$,

$$c = \begin{pmatrix} \tau 1_n \\ (1 - \tau) 1_n \\ \lambda 1_p \\ 0 1_p \end{pmatrix}, \quad B = \begin{pmatrix} -E_{p \times n} & 0 & 0 & 0 \\ 0 & -E_{p \times n} & 0 & 0 \\ 0 & 0 & -I_p & I_p \\ 0 & 0 & -I_p & -I_p \\ 0 & 0 & 0 & I_p \end{pmatrix}$$

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Solution - Continued

The previous problem may be reformulated into *standard form*

$$\begin{aligned} & \text{minimize} && c^\top x \\ & \text{subject to} && Cx = d, \\ & && x + s = u, \quad x \geq 0, s \geq 0 \end{aligned}$$

and the dual problem is:

$$\begin{aligned} & \text{maximize} && d^\top y - u^\top w \\ & \text{subject to} && C^\top y - w + z = c, \quad z \geq 0, w \geq 0 \end{aligned}$$

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Solution - Continued

The KKT conditions for this linear program are

$$F(x, y, z, s, w) = \begin{pmatrix} Cx - d \\ x + s - u \\ C^T y - w + z - c \\ x \circ z \\ s \circ w \end{pmatrix} = 0,$$

with $y \geq 0$, $z \geq 0$ dual slacks, $s \geq 0$ primal slacks, $w \geq 0$ dual variables.

This can be solved by a primal-dual path following algorithm based on the *Newton method*

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Adaptive Lasso Procedure

Lasso estimates $\hat{\beta}$ can be inconsistent (Zou, 2006) in some scenarios.

Lasso soft-threshold function gives biased results

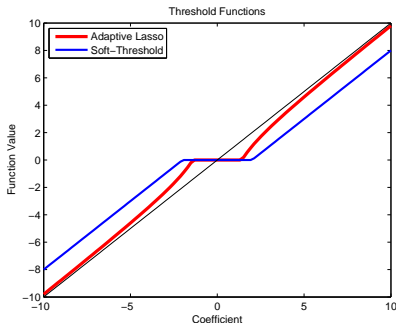


Figure 20: Threshold functions for simple and adaptive Lasso



Adaptive Lasso Procedure

The adaptive Lasso (Zou, 2006) yields a sparser solution and is less biased.

L_1 - penalty replaced by a re-weighted version; $\hat{\omega}_j = 1/|\hat{\beta}_j^{\text{init}}|^\gamma$,
 $\gamma = 1$, $\hat{\beta}^{\text{init}}$ is from (3)

The adaptive lasso estimates are given by:

$$\hat{\beta}_\lambda^{\text{adapt}} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (Y_i - X_i^\top \beta)^2 + \lambda \|\hat{\omega}^\top \beta\|_1$$

(Bühlmann, van de Geer, 2011): $\hat{\beta}_j^{\text{init}} = 0$, then $\hat{\beta}_j^{\text{adapt}} = 0$

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Simple and Adaptive Lasso Penalized QR

Simple lasso-penalized QR optimization problem is:

$$\hat{\beta}_{\tau,\lambda} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_{\tau}(Y_i - X_i^{\top} \beta) + \lambda \|\beta\|_1 \quad (5)$$

Adaptive lasso-penalized QR model uses the re-weighted penalty:

$$\hat{\beta}_{\tau,\lambda}^{\text{adapt}} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_{\tau}(Y_i - X_i^{\top} \beta) + \lambda \|\hat{\omega}^{\top} \beta\|_1 \quad (6)$$

Adaptive lasso-penalized QR procedure can ensure oracle properties for the estimator [▶ Details](#)

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Algorithm for Adaptive Lasso Penalized QR

The optimization for the adaptive lasso quantile regression can be re-formulated as a lasso problem:

- the covariates are rescaled: $\tilde{X} = (X_1 \circ \hat{\beta}_1^{\text{init}}, \dots, X_p \circ \hat{\beta}_p^{\text{init}})$;
- the lasso problem (5) is solved:

$$\hat{\beta}_{\tau, \lambda} = \arg \min_{\beta \in \mathbb{R}^p} \sum_{i=1}^n \rho_{\tau}(Y_i - \tilde{X}_i^{\top} \beta) + \lambda \|\beta\|_1$$

- the coefficients are re-weighted as $\hat{\beta}_{\tau, \lambda}^{\text{adapt}} = \hat{\beta}_{\tau, \lambda} \circ \hat{\beta}^{\text{init}}$

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Oracle properties

In the linear model, let $Y = X\beta + \varepsilon = X^1\beta^1 + X^2\beta^2 + \varepsilon$, where $X = (X^1, X^2)$, $X^1 \in \mathbb{R}^{n \times q}$, $X^2 \in \mathbb{R}^{n \times (p-q)}$; β_q^1 are true nonzero coefficients, $\beta_{p-q}^2 = 0$ are noise coefficients; $q = \|\beta\|_0$.

Also assume that $\lambda q / \sqrt{n} \rightarrow 0$ and $\lambda / \{\sqrt{q} \log(n \vee p)\} \rightarrow \infty$ and certain regularity conditions are satisfied [▶ Details](#)



Oracle properties

Then the adaptive L_1 QR estimator has the "oracle" properties (Zheng et al., 2013):

1. Variable selection consistency:

$$P(\beta^2 = 0) \geq 1 - 6 \exp \left\{ -\frac{\log(n \vee p)}{4} \right\}.$$

2. Estimation consistency: $\|\beta - \hat{\beta}\| = \mathcal{O}_p(\sqrt{q/n})$

3. Asymptotic normality: $u_q^2 \stackrel{\text{def}}{=} \alpha^T \Sigma_{11} \alpha, \forall \alpha \in \mathbb{R}^q, \|\alpha\| < \infty,$

$$n^{1/2} u_q^{-1} \alpha^T (\beta^1 - \hat{\beta}^1) \xrightarrow{\mathcal{L}} N \left\{ 0, \frac{(1 - \tau)\tau}{f^2(\gamma^*)} \right\}$$

where γ^* is the τ th quantile and f is the pdf of ε

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Selected Hedge Funds' Strategies

1. *Convertible arbitrage* hedge funds focus on the mispricing of convertible bonds. A typical position involves a long position in the convertible bond and a short position in the underlying asset.
2. *Fixed income arbitrage* hedge funds tend to profit from price anomalies between related securities and/or bet on the evolution of interest rates spreads. Typical trading strategies are butterfly-like structures, cash/futures basis trading strategies or relative swap spread trades.
3. *Event-driven* hedge funds focus on price movements generated by an anticipated corporate event, such as a merger, an acquisition, a bankruptcy, etc. [▶ Back](#)



Selected Hedge Funds' Strategies

4. *Long/short equity* hedge funds represent the original hedge fund model. They invest in equities both on the long and the short sides, and generally have a small net long exposure. They are genuinely opportunistic strategies and could be classified as "double alpha, low beta" funds.
5. *Market neutral* hedge funds seek to neutralize certain market risks by taking offsetting long and short positions in instruments with actual or theoretical relationships. Most of them are in fact long/short equity hedge funds.
6. *Dedicated short bias* hedge funds are essentially long/short equity hedge funds, that maintain a consistent net short exposure, therefore attempting to capture profits when the market declines. [▶ Back](#)



Selected Hedge Funds' Strategies

7. *Emerging market* hedge funds invest in equities and fixed-income securities of emerging markets around the world.
8. *Global macro* hedge funds take very large directional bets on overall market directions that reflect their forecasts of major economic trends and/or events.
9. *Managed futures* hedge funds implement discretionary or systematic trading in listed financial, commodity and currency futures around the world. The managers of these funds are known as commodity trading advisors (CTAs).
10. *Multi-strategy* hedge funds regroup managers acting in several of the above-mentioned strategies. [▶ Return](#)



Traditional Assets/Hedge Fund Indices

Table 7: Correlation statistics for MSCI and hedge funds' indices returns

Hedge Fund Indices	MSCI Indices								
	WRD	EUR	US	UK	FR	SW	GER	JAP	PAC
Asia CTA	-0.01	0.02	-0.02	-0.06	0.01	-0.09	0.04	-0.03	0.02
Asia Distressed Debt	0.30	0.30	0.24	0.31	0.31	0.26	0.27	0.26	0.34
Asia Macro	-0.01	-0.01	-0.04	0.01	-0.02	0.07	-0.03	0.06	0.06
Global CTA FoF	0.02	0.08	-0.08	0.09	0.10	0.09	0.07	0.06	0.10
Global Event Driven FoF	0.65	0.59	0.58	0.66	0.59	0.50	0.57	0.47	0.67
Global Macro FoF	0.19	0.22	0.07	0.24	0.22	0.18	0.20	0.23	0.31
CTA/Managed Futures	-0.04	0.02	-0.13	0.03	0.03	0.07	-0.01	0.04	0.05
Event Driven	0.82	0.75	0.75	0.78	0.75	0.64	0.75	0.62	0.83
Fixed Income	0.70	0.65	0.63	0.70	0.65	0.56	0.62	0.51	0.78
Long Short Equities	0.82	0.78	0.74	0.76	0.77	0.64	0.77	0.64	0.82
Asia inc Japan Distr. Debt	0.30	0.30	0.24	0.31	0.31	0.26	0.27	0.26	0.34
Asia inc Japan Macro	0.34	0.33	0.31	0.27	0.33	0.24	0.35	0.31	0.40

Calculations based on monthly data Jan. 2000 - Jul. 2012

WRD - World, EUR - Eurozone, FR - France, SW - Switzerland, PAC - Pacific ex. Japan

FoF means "fund of funds"

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Risk-Return Asset Allocation

Log returns $X_t \in \mathbb{R}^p$:

$$\begin{aligned} \min_{w_t \in \mathbb{R}^p} \quad & \sigma_{P,t}^2(w_t) \stackrel{\text{def}}{=} w_t^\top \Sigma_t w_t \\ \text{s.t.} \quad & \mu_{P,t}(w_t) = r_T, \\ & w_t^\top \mathbf{1}_p = 1, \\ & w_{i,t} \geq 0 \end{aligned} \tag{7}$$

where r_T "target" return, $\Sigma_t \stackrel{\text{def}}{=} E_{t-1}\{(X_t - \mu)(X_t - \mu)^\top\}$, Σ_t is modeled with a GARCH model

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Tail Risk Asset Allocation

Portfolio returns $X \in \mathbb{R}^{n \times p}$, Bassett et al. (2004)

$$\begin{aligned} \min_{(\beta, \alpha)^T \in \mathbb{R}^p} \quad & \sum_{t=1}^n \rho_{\tau} \left\{ X_{t1} - \sum_{j=2}^p (X_{t1} - X_{tj}) \beta_j - \alpha \right\} \\ \text{s.t.} \quad & w^{\top} \hat{\mu} = r_{\mathcal{T}}, \\ & w^{\top} \mathbf{1}_p = 1, \end{aligned} \quad (8)$$

where $r_{\mathcal{T}}$ is the "target" return for the portfolio and $w = w(\beta) = (1 - \sum_{j=2}^p \beta_j, \beta^{\top})^{\top}$, $\tau \in (0, 1)$, $\hat{\mu} \stackrel{\text{def}}{=} \bar{X}$ sample returns' mean

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Multi-Moment Utility Optimization

The (dynamic) investment decision: $U(\cdot)$ utility function; $X_t \in \mathbb{R}^p$ log-returns, w_t weights, $\mu_{P,t}(w_t) \stackrel{\text{def}}{=} w_t^\top \mu$, $\mu \stackrel{\text{def}}{=} E_{t-1}(X_t)$, r_T "target" return:

$$\max_{w_t \in \mathbb{R}^p} E_{t-1} \{U(W_t)\}, \quad \text{s.t. } \mu_{P,t}(w_t) = r_T, \quad w^\top \mathbf{1}_p = 1, \quad w_{i,t} \geq 0, \quad (9)$$

$$\begin{aligned} E_{t-1} \{U(W_t)\} &\approx U\{\bar{W}_t\} + \frac{1}{2} U^{(2)}\{\bar{W}_t\} \sigma_{W_t}^2 + \\ &+ \frac{1}{3!} U^{(3)}\{\bar{W}_t\} S_{W_t} + \frac{1}{4!} U^{(4)}\{\bar{W}_t\} K_{W_t}, \end{aligned}$$

where $W_t \stackrel{\text{def}}{=} 1 + w_t^\top X_t$ is the end-of-period t wealth, $\bar{W}_t \stackrel{\text{def}}{=} E_{t-1}(W_t)$, $\sigma_{W_t}^2 \stackrel{\text{def}}{=} E_{t-1} \{(W_t - \bar{W}_t)^2\}$, $S_{W_t} \stackrel{\text{def}}{=} E_{t-1} \{(W_t - \bar{W}_t)^3\}$, $K_{W_t} \stackrel{\text{def}}{=} E_{t-1} \{(W_t - \bar{W}_t)^4\}$; $U^{(n)}(\cdot)$ is the n th derivative of $U(\cdot)$

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Utility Function Example

- CARA utility:

$$U(W) = -\exp(-\eta W),$$

where η coefficient of risk aversion

- then:

$$\begin{aligned} E_{t-1} \{U(W_t)\} &= E_{t-1} \{-\exp(-\eta W_t)\} \\ &\approx -\exp(-\eta \bar{W}_t) \left(1 + \frac{\eta^2}{2} \sigma_{W_t}^2 - \frac{\eta^3}{3!} S_{W_t} + \frac{\eta^4}{4!} K_{W_t} \right) \end{aligned}$$

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Portfolio Moments

The portfolio moments:

$$\sigma_{W_t}^2 = w_t^\top M_t^2 w_t$$

$$S_{W_t} = w_t^\top M_t^3 (w_t \otimes w_t)$$

$$K_{W_t} = w_t^\top M_t^4 (w_t \otimes w_t \otimes w_t),$$

where \otimes Kronecker product,

$$M_t^2 \stackrel{\text{def}}{=} E_{t-1}(r_t - \mu)^2 \quad (10)$$

$$M_t^3 \stackrel{\text{def}}{=} E_{t-1}\{(r_t - \mu)(r_t - \mu)^\top \otimes (r_t - \mu)^\top\} \quad (11)$$

$$M_t^4 \stackrel{\text{def}}{=} E_{t-1}\{(r_t - \mu)(r_t - \mu)^\top \otimes (r_t - \mu)^\top \otimes (r_t - \mu)^\top\}, \quad (12)$$

A dynamic distribution model is used to obtain M_t^2 , M_t^3 , M_t^4 in (10), (11), (12) [▶ Return to "TEDAS Gestalten"](#)



Dynamic Distribution Model

- joint normality is questionable
- possible persistence in the dynamics of moments
- reaction of distribution parameters to past shocks
- computational feasibility

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Descriptive Statistics

Table 8: Monthly returns of 3 Eureka hedge funds' indices

	Japan Multi-Strategy		North America Fixed Income		Europe Arbitrage	
Univariate statistics						
Normality tests						
JB	533.775	(0.000)	294.089	(0.000)	610.407	(0.000)
KS	0.503	(0.000)	0.473	(0.000)	0.485	(0.000)
Omnibus	82.773	(0.000)	43.761	(0.000)	171.079	(0.000)
Dynamic conditional moments' tests						
ARCH	11.227	(0.000)	34.966	(0.000)	26.592	(0.000)
Bera-Lee	48.469	(0.000)	36.475	(0.000)	40.783	(0.000)
Bera-Zuo	203.723	(0.000)	20.149	(0.166)	421.847	(0.000)
Multivariate statistics						
Test						
Omnibus	326.226	(0.000)				
Mardia	301.199	(0.000)				
Henze-Zirkler	9.862	(0.000)				

Standard errors and p -values are given in parentheses.

ARCH, Bera-Lee and Bera-Zuo stand for the test statistics of the ARCH test by Engle (1982) and information matrix tests for testing variation in second, third and fourth conditional moments

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Generalized Hyperbolic (GH) Distribution

A vector X has a multivariate GH distribution if

$$X = \mu + W\delta + \sqrt{W}AZ, \quad (13)$$

where

- (i) $Z \sim N(0, I_k)$
- (ii) $A \in \mathbb{R}^{d \times k}$
- (iii) $\mu, \delta \in \mathbb{R}^d$
- (iv) $W \geq 0$, scalar-valued random variable, independent of Z ,
 $W \sim GIG(\lambda, \alpha, \beta)$; GIG is the generalized inverse Gaussian distribution

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Multivariate Affine GH Distribution

- margins of the (MGH) distribution not mutually independent for some choice of $\Sigma = AA^T$
- MAGH distribution, Schmidt et al. (2006), models margins and dependency independently

$Y \sim \text{MAGH}(\lambda, \alpha, \beta, \mu, \Sigma)$ if

- $X = (X_1, \dots, X_d)^T$, $X_i \sim \text{GH}(0, 1, \alpha_i, \beta_i)$, $i = 1, \dots, d$
- $Y = AX + \mu$, $AA^T = \Sigma$ positive definite

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Normal Inverse Gaussian (NIG) Distribution

- obtained from the GH distribution with $\lambda = -0.5$
- "semi-heavy tails" property: fits financial data well

The density is written as:

$$f_{NIG}(x) = \frac{\alpha\delta}{\pi} \exp\left\{\delta\sqrt{\alpha^2 - \beta^2} + \beta(x - \mu)\right\} \frac{K_1\left\{\alpha\sqrt{\delta^2 + (x - \mu)^2}\right\}}{\sqrt{\delta^2 + (x - \mu)^2}},$$

where $0 \leq |\beta| \leq \alpha$, $\delta > 0$, K_1 is the modified Bessel function of the third kind and order 1

Location-Scale Property: let $\bar{\alpha} \stackrel{\text{def}}{=} \delta\alpha$ and $\bar{\beta} \stackrel{\text{def}}{=} \delta\beta$, then
 $X \sim NIG(\bar{\alpha}, \bar{\beta}, \mu, \delta) \Leftrightarrow (X - \mu)/\delta \sim NIG(\bar{\alpha}, \bar{\beta}, 0, 1)$

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Choice of the Matrix A

- assume $X = As$, X random signal generated by another random vector $s = (s_1, \dots, s_d)$, s_i statistically independent, $i = 1, \dots, d$ and a mixing matrix A , both unknown
- the *independent component analysis* (ICA) technique separates source signals s from a set of mixed signals X without or with very little aid of information about f or the mixing process A
- ICA estimates A and s by maximizing the nongaussianity of linear combinations of X

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The Model for Portfolio Returns

- assume $\varepsilon_t = A s_t$, $E \varepsilon_t = 0$, $E \varepsilon_t \varepsilon_t^\top = I_d$, $E s_t = 0$, $E s_t s_t^\top = I_d$
- define $E(s_t | \mathcal{F}_t) = 0$, $D_t \stackrel{\text{def}}{=} E(s_t s_t^\top | \mathcal{F}_t) \stackrel{\text{def}}{=} \text{diag}(d_{1t}, \dots, d_{dt})$
- let $z_{it} \sim NIG(\bar{\alpha}_{it}, \bar{\beta}_{it}, 0, 1)$, then $s_{it} \sim NIG(\bar{\alpha}_{it}/\sqrt{d_{it}}, \bar{\beta}_{it}/\sqrt{d_{it}}, 0, \sqrt{d_{it}})$
- *MANIG*: multivariate affine normal inverse Gaussian distribution
- model for portfolio returns $r_t = m_t + \varepsilon_t$,
 $r_t | \mathcal{F}_t \sim \text{MANIG}(m_t, \Sigma_t, \omega_t)$, where $\omega_t = (\omega_{1t}, \dots, \omega_{dt})^\top$ and $\omega_{it} = (\alpha_{it}, \beta_{it})^\top$, $i = 1, \dots, d$, $\Sigma_t = M_t^2 = A D_t A^\top$, d_{it} can be modeled as GARCH-type processes

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Moment Dynamics

- reparametrize the model to have asymmetry and shape parameters $\xi_{it} = \beta_{it}/\alpha_{it}$, $\nu_{it} = \sqrt{\alpha_{it}^2 - \beta_{it}^2}$
- introduce asymmetric GARCH-like dynamics:

$$\nu_{i,t} = a_{i,0} + a_{i,1}^- |s_{i,t-1}| N_{i,t-1} + a_{i,1}^+ |s_{i,t-1}| P_{i,t-1} + a_{i,2} \nu_{i,t-1} \quad (14)$$

$$\xi_{i,t} = b_{i,0} + b_{i,1}^- s_{i,t-1} N_{i,t-1} + b_{i,1}^+ s_{i,t-1} P_{i,t-1} + b_{i,2} \xi_{i,t-1}, \quad (15)$$

where $N_{i,t} = \mathbf{I}(z_{i,t} \leq 0)$, $P_{i,t} = 1 - N_{i,t}$

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Portfolio Moments

$$M_t^3 = AM_{S_t}^3(A \otimes A)^\top, \quad M_t^4 = AM_{S_t}^4(A \otimes A \otimes A)^\top,$$

where

$$M_{S_t}^3 = E_{t-1}(s_{i,t}s_{j,t}s_{k,t}) = \sum_{r=1}^p d_{ir,t}d_{jr,t}d_{kr,t}sk_{rt}^s$$

$$\begin{aligned} M_{S_t}^4 &= E_{t-1}(s_{i,t}s_{j,t}s_{k,t}s_{l,t}) \\ &= \sum_{r=1}^p d_{ir,t}d_{jr,t}d_{kr,t}d_{lr,t}kurt_{rt}^s + \sum_{r=1}^p \sum_{s \neq r} \psi_{rs,t}, \end{aligned}$$

$$\psi_{rs,t} = d_{ir,t}d_{jr,t}d_{ks,t}d_{ls,t} + d_{ir,t}d_{js,t}d_{kr,t}d_{ls,t} + d_{is,t}d_{jr,t}d_{kr,t}d_{ls,t},$$

$$D_t^{1/2} = (d_{ij,t})_{i,j=1,\dots,p}, \quad sk_{it}^s, \quad kurt_{it}^s \text{ are obtained with } \alpha_{it}, \beta_{it}$$

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Conditional VaR (CVaR) Optimization

Given $\alpha > 0.5$ confidence level,

$$\min_{w_t \in \mathbb{R}^p} \text{CVaR}_\alpha(w_t), \quad \text{s.t. } \mu_{P,t}(w_t) = r_T, w_t^\top \mathbf{1}_p = 1, w_{i,t} \geq 0, \quad (16)$$

$$\text{CVaR}_\alpha(w_t) = -\frac{1}{1-\alpha} q_{\alpha^*}^*(w_t) \sigma_{P,t}(w_t), \quad \text{▶ Proof} \quad (17)$$

where (via Cornish-Fisher (CF) expansion):

$$q_{\alpha^*}^*(w_t) = \left\{ 1 + \frac{S_{P,t}(w_t)}{6} z_{\alpha^*} + \frac{K_{P,t}(w_t)}{24} (z_{\alpha^*}^2 - 1) - \frac{S_{P,t}^2(w_t)}{36} (2z_{\alpha^*}^2 - 1) \right\} \varphi(z_{\alpha^*}), \quad (18)$$

where $\alpha^* \stackrel{\text{def}}{=} 1 - \alpha$ [▶ Return to "TEDAS Gestalten"](#) [▶ Back to "TEDAS Example"](#)



The Orthogonal GARCH Model

- $X_t \in \mathbb{R}^{n \times p}$, $\Gamma_t = B_t \in \mathbb{R}^{p \times p}$ matrix of standardized eigenvectors of $n^{-1}X_t^\top X_t$ ordered according to decreasing magnitude of eigenvalues
- $F_t = P_t \stackrel{\text{def}}{=} X_t \Gamma_t$ PCs of X_t
- factors f , introduce noise u_i , i.e.
 $y_j = b_{j1}f_1 + b_{j2}f_2 + \dots + b_{jk}f_k + u_j$ or $Y_t = F_t B_t^\top + U_t$
- then $\Sigma_t = \text{Var}(X_t) = \text{Var}(F_t B_t^\top) + \text{Var}(U_t) = B_t \Delta_t B_t^\top + \Omega_t$,
 $\Delta_t = \text{Var}(F_t)$ diagonal matrix of PC variances at t

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Cornish-Fisher VaR Optimization

Log returns $X_t \in \mathbb{R}^p$:

$$\begin{aligned} & \underset{w \in \mathbb{R}^d}{\text{minimize}} && W_t \{-q_\alpha(w_t) \cdot \sigma_p(w_t)\} \\ & \text{subject to} && w_t^\top \mu = \mu_p, \quad w_t^\top \mathbf{1} = 1, \quad w_{t,i} \geq 0 \end{aligned}$$

here $W_t \stackrel{\text{def}}{=} W_0 \cdot \prod_{j=1}^{t-1} w_{t-j}^\top (1 + X_{t-j})$, \tilde{w} , W_0 initial wealth,

$$\sigma_p^2(w) \stackrel{\text{def}}{=} w_t^\top \Sigma_t w_t,$$

$$q_\alpha(w) \stackrel{\text{def}}{=} z_\alpha + (z_\alpha^2 - 1) \frac{S_p(w)}{6} + (z_\alpha^3 - 3z_\alpha) \frac{K_p(w)}{24} - (2z_\alpha^3 - 5z_\alpha) \frac{S_p(w)^2}{36},$$

here $S_p(w)$ skewness, $K_p(w)$ kurtosis, z_α is $N(0, 1)$ α -quantile. If $S_p(w)$, $K_p(w)$ zero, then obtain Markowitz allocation

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Portfolio Skewness and Kurtosis

Skewness S_P and excess kurtosis K_P are given by moment expressions

$$S_P(w) = \frac{1}{\sigma_P^3(w)} (m_3 - 3m_2m_1 + 2m_1^3)$$

$$K_P(w) = \frac{1}{\sigma_P^4(w)} (m_4 - 4m_3m_1 + 6m_2m_1^2 + 3m_1^4) - 3$$

where portfolio non-central moments also depend on w :

$$m_1 = \mu_P(w) \stackrel{\text{def}}{=} w^\top \mu$$

$$m_2 = \sigma_P^2 + m_1^2$$

$$m_3 = \sum_{i=1}^d \sum_{j=1}^d \sum_{k=1}^d w_i w_j w_k S_{ijk}$$



Portfolio Skewness and Kurtosis - Continued

$$m_4 = \sum_{i=1}^d \sum_{j=1}^d \sum_{k=1}^d \sum_{l=1}^d w_i w_j w_k w_l K_{ijkl},$$

where $\sigma_p^2(w) = w^\top \Sigma w$ and $S_{ijk} = E(r_i \times r_j \times r_k)$, $K_{ijkl} = E(r_i \times r_j \times r_k \times r_l)$ can be computed via sample averages from returns data.

S_{ijk} , K_{ijkl} determine the d -dimensional portfolio co-skewness and co-kurtosis tensors

$$S \stackrel{\text{def}}{=} \{S_{ijk}\}_{i,j,k=1,\dots,d} \in \mathbb{R}^{d \times d \times d}$$
$$K \stackrel{\text{def}}{=} \{K_{ijkl}\}_{i,j,k,l=1,\dots,d} \in \mathbb{R}^{d \times d \times d \times d}.$$

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Regularity Conditions for Adaptive Lasso QR

A1 Sampling and smoothness: $\forall x$ in the support of X_i , $\forall y \in \mathbb{R}$,
 $f_{Y_i|X_i}(y|x)$, $f \in \mathcal{C}^k(\mathbb{R})$, $|f_{Y_i|X_i}(y|x)| < \bar{f}$, $|f'_{Y_i|X_i}(y|x)| < \bar{f}'$; $\exists \underline{f}$,
 such that $f_{Y_i|X_i}(x^\top \beta_\tau | x) > \underline{f} > 0$

A2 Restricted identifiability and nonlinearity: let $\delta \in \mathbb{R}^p$,
 $T \subset \{0, 1, \dots, p\}$, δ_T such that $\delta_{Tj} = \delta_j$ if $j \in T$, $\delta_{Tj} = 0$ if
 $j \notin T$; $T = \{0, 1, \dots, s\}$, $\bar{T}(\delta, m) \subset \{0, 1, \dots, p\} \setminus T$, then
 $\exists m \geq 0$, $c \geq 0$ such that

$$\inf_{\delta \in A, \delta \neq 0} \frac{\delta^\top \mathbb{E}(X_i X_i^\top) \delta}{\|\delta_{T \cup \bar{T}(\delta, m)}\|^2} > 0, \quad \frac{3\underline{f}^{3/2}}{8\bar{f}'} \inf_{\delta \in A, \delta \neq 0} \frac{\mathbb{E}[|X_i^\top \delta|^2]^{3/2}}{\mathbb{E}[|X_i^\top \delta|^3]} > 0,$$

where $A \stackrel{\text{def}}{=} \{\delta \in \mathbb{R}^p : \|\delta_{T^c}\|_1 \leq c \|\delta_T\|_1, \|\delta_{T^c}\|_0 \leq n\}$

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Regularity Conditions - Continued

A3 Growth rate of covariates:

$$\frac{q^3 \{\log(n \vee p)\}^{2+\eta}}{n} \rightarrow 0, \eta > 0$$

A4 Moments of covariates: Cramér condition

$$E[|x_{ij}|^k] \leq 0.5 C_m M^{k-2} k!$$

for some constants $C_m, M, \forall k \geq 2, j = 1, \dots, p$

A5 Well-separated regression coefficients: $\exists b_0 > 0$, such that
 $\forall j \leq q, |\hat{\beta}_j| > b_0$



Proof of the CF-CVaR Expansion 1

Define the Cornish-Fisher expansion:

$$q_{1-\alpha} \stackrel{\text{def}}{=} z_{1-\alpha} + (z_{1-\alpha}^2 - 1)s + (z_{1-\alpha}^3 - 3z_{1-\alpha})k - (2z_{1-\alpha}^3 - 5z_{1-\alpha})s^2,$$

where $s \stackrel{\text{def}}{=} S/6$, $k \stackrel{\text{def}}{=} K/24$, S and K are skewness and excess kurtosis, respectively; $z_{1-\alpha} \stackrel{\text{def}}{=} \Phi^{-1}(1 - \alpha)$.

Re-write:

$$q_{1-\alpha} = a_0 + a_1 z_{1-\alpha} + a_2 z_{1-\alpha}^2 + a_3 z_{1-\alpha}^3, \quad (19)$$

where $a_0 = -s$, $a_1 = 1 - 3k + 5s^2$, $a_2 = s$, $a_3 = k - 2s^2$



Proof of the CF-CVaR Expansion 2

Define the *conditional Value-at-Risk* (CVaR) or *expected shortfall* (ES):

$$\text{CVaR}_\alpha \stackrel{\text{def}}{=} \frac{1}{1-\alpha} \int_\alpha^1 \text{VaR}_q \, d q,$$

where $\text{VaR}_q \stackrel{\text{def}}{=} -\Phi^{-1}(\alpha)\sigma\sqrt{T}$

Observe:

$$\text{CVaR}_\alpha = -\frac{1}{1-\alpha} \int_\alpha^1 \Phi^{-1}(q)\sigma\sqrt{T} \, d q \quad (20)$$

$$= -\frac{\sigma\sqrt{T}}{1-\alpha} \int_\alpha^1 \Phi^{-1}(q) \, d q \quad (21)$$

$$= -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} u\varphi(u) \, d u, \quad (22)$$

where (22) follows from the change of variable: $u = z_q = \Phi^{-1}(q)$



Proof of the CF-CVaR Expansion 3

Substitute (19) into (22):

$$\begin{aligned}\text{CVaR}_\alpha &= -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} (a_0 + a_1z + a_2z^2 + a_3z^3) \varphi(z) dz \\ &= a_0A_0 + a_1A_1 + a_2A_2 + a_3A_3,\end{aligned}$$

$$A_0 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} \varphi(z) dz = -\sigma\sqrt{T},$$

$$A_1 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} z\varphi(z) dz = \frac{\sigma\sqrt{T}}{1-\alpha} \varphi(z_\alpha),$$

$$A_2 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} z^2\varphi(z) dz = -\sigma\sqrt{T} \left(\frac{\varphi(z_\alpha)z_\alpha}{1-\alpha} + 1 \right),$$

$$A_3 = -\frac{\sigma\sqrt{T}}{1-\alpha} \int_{-\infty}^{-z_\alpha} z^3\varphi(z) dz = \frac{\sigma\sqrt{T}}{1-\alpha} (z_\alpha^2 + 2)\varphi(z_\alpha).$$

Collecting terms and simplifying gives the desired result. □

[▶ Return to "Conditional VaR Optimization"](#)



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