

# Uniform Confidence for Pricing Kernels

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## Motivation

- Arbitrage free market; riskless bond with rate  $r$
- Underlying price process  $\{S_t\}$

Price  $z_t$  at time  $t$  from a payoff  $\psi(S_T)$

$$\begin{aligned} z_t &= \int_0^\infty \exp(-r\tau) \psi(x) dQ_{S_T}(x) \\ &= \int_0^\infty \exp(-r\tau) \psi(x) \frac{q_t(x)}{p_t(x)} dP_{S_T}(x) \end{aligned}$$



Figure 1: **conditional measure** at time of maturity  $T$  built upon a path of the stochastic process for **underlying asset** with information up to time  $t$ .



## Empirical Pricing Kernel (EPK)

Pricing Kernel (PK) a stochastic discount factor, i.e.

$$\mathcal{K}_{t,\tau}(x) = \exp(-r\tau) \frac{q_t(x)}{p_t(x)}$$

EPK is therefore an estimate of PK:

$$\widehat{\mathcal{K}}_{t,\tau}(x) = \exp(-r\tau) \frac{\widehat{q_t(x)}}{\widehat{p_t(x)}}$$



## The EPK Paradox

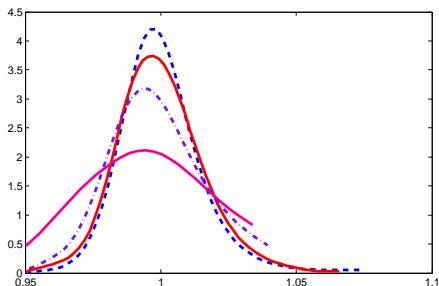


Figure 2: Examples of inter-temporal pricing kernels with maturity 0.00833(3D) respectively on 17-Jan-2006 (blue), 18-Apr-2006 (red), 16-May-2006 (magenta), 13-June-2006 (black).



## The EPK Paradox

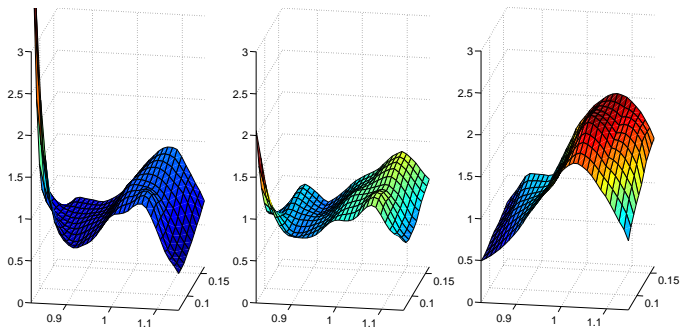


Figure 3: Estimated PK across moneyiness and maturity



## The EPK Paradox

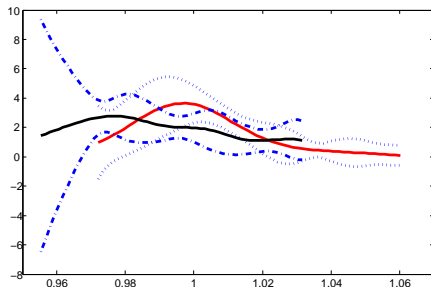


Figure 4: Examples of inter-temporal pricing kernels with various maturities in years: 0.02222 (8D) (red) 0.1(36D) (green) on 12-Jan-2006 and their confidence bands



## Aims

- Nonparametric confidence band to test alternatives
- Check the statistical significance of the EPK puzzle
- Investigate shapes of EPKs: investor preference
- Understand the dynamics of risk patterns



## Outline

1. Motivation ✓
2. Uniform Confidence Band
3. Monte-Carlo Study
4. Empirical Data Analysis





## Risk Neutral Density (RND) Estimation

The RND may be estimated from option prices, Breeden and Litzenberger (1978):

$$q_t(S_T) = \exp(r\tau) \frac{\partial^2 H_t(k, \tau)}{\partial k^2} \Big|_{k=S_T}$$

with call price function  $H_t(k, \tau)$ .

Aït-Sahalia and Lo (1998) estimate  $H_t(k, \tau)$  nonparametrically and differentiate it twice w.r.t.  $k$ .



Call prices  $(X_i, Y_i)$ , with fixed  $\tau$ , we have:

$$Y_i = H(x_i) + \varepsilon_i, i = 1, \dots, n_q$$

$(X_i, Y_i)$ s are i.i.d.

Define  $L\{y; H(u)\}$  as the conditional density of  $Y$  given  $K = u$

Local polynomial estimate, ( $x \approx u$ ):

$$H(u) \approx H(x, u) \stackrel{\text{def}}{=} \sum_{j=0}^3 H_j(x)(u - x)^j$$

Local likelihood

$$L_{n_q}\{H(x)\} \stackrel{\text{def}}{=} \frac{1}{n_q} \sum_{i=1}^{n_q} K_{h_{n_q}}(x - x_i) \log L\{Y_i; H(x_i, x)\},$$



## Data

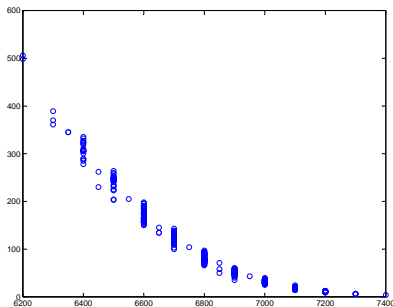


Figure 5: Plot of call prices against strikes  $k$ ,  $n_q = 1000$ ,  $n_p = 500$ .

□ **Source:** Research Data Center (RDC)

<http://sfb649.wiwi.hu-berlin.de>



Solutions:

$$\widehat{\mathbf{H}}(\mathbf{x}) \stackrel{\text{def}}{=} \operatorname{argmax}_H L_{n_q}\{H(x)\},$$

where

$$\widehat{\mathbf{H}}(\mathbf{x}) = \{\widehat{H_0(x)}, \widehat{H_1(x)}, \widehat{H_2(x)}, \widehat{H_3(x)}\}^\top$$

Estimate for  $q_t(x)$ :

$$\widehat{q_t(x)} \propto 2! \widehat{H_2(x)}$$

The kernel density estimate for  $p_t(x)$  is based on historical  $\{S_t\}$ :

$$\widehat{p_t(x)} = n_p^{-1} \sum_{j=1}^{n_p} K_{h_{n_p}}(x - S_j)$$



## Uniform Convergence

### Theorem

*Under regularity conditions, for all  $x$  in an interval  $J$ , we have a.s.,*

$$\sup_{x \in J} |\widehat{\mathcal{K}_{t,\tau}(x)} - \mathcal{K}_{t,\tau}(x)| = \mathcal{O}[\max\{(n_p h_{n_p} / \log n_p)^{-0.5}, h_{n_p}^2, h_{n_q}^3, h_{n_q}^{-2} \{n_q h_{n_q} / \log n_q\}^{-0.5}\}]$$



## Uniform Confidence Band

### Theorem

*Under regularity conditions,*

$$\mathcal{K}_{nt,\tau}(x) \stackrel{\text{def}}{=} n_q^{1/2} h_{n_q}^{5/2} \{ \widehat{\mathcal{K}_{t,\tau}(x)} - \mathcal{K}_{t,\tau}(x) \} \text{Var}\{ \widehat{\mathcal{K}_{t,\tau}(x)} \}^{-1/2}.$$

*We have:*

$$\begin{aligned} \mathbb{P} \left\{ (-2 \log h_{n_q})^{1/2} \left\{ \sup_{x \in J} |\mathcal{K}_{nt,\tau}(x)| - c_{nt} \right\} < z \right\} \\ \longrightarrow \exp\{-2 \exp(-z)\}, \end{aligned}$$

where  $c_{nt} = (-2 \log h_{n_q})^{1/2} + (-2 \log h_{n_q})^{-1/2} \{x_\alpha + \log(C/2\pi)\}$



## Uniform Confidence Band

Thus, a  $(1 - \alpha)100\%$  confidence band for pricing kernel  $\mathcal{K}_{t,\tau}$  is:

$$[f(x) : \sup_x \{|\widehat{\mathcal{K}_{t,\tau}}(x) - f(x)| \widehat{\text{Var}}(\widehat{\mathcal{K}_{t,\tau}}(x))^{-1/2}\} \leq L_\alpha]$$

where

$$L_\alpha = 2!(n_q h_{n_q}^5)^{-1/2} c_{nt}$$

and

$$x_\alpha = -\log\{-1/2 \log(1 - \alpha)\}$$



## Extension on $\tau$

Let  $\mathfrak{x}$  be the possible set of maturities, the extension of our results over  $\tau$  :

$$[f_{t,\tau}(x) : \sup_{x \in E, \tau \in \mathfrak{x}} \{|\widehat{\mathcal{K}}_{t,\tau}(x) - f_{t,\tau}(x)| \widehat{\text{Var}}(\widehat{\mathcal{K}}_{t,\tau}(x))^{-1/2}\} \leq L_\alpha].$$

In the BS setup, the evolution of bands over time, for fixed  $\tau_1$   
 $(g(\tau_1 - \tau_2) = \mathcal{K}_{t,\tau_1}(x)/\mathcal{K}_{t,\tau_2}(x))$

$$[f_{t,\tau_2} : \widehat{g}(\tau_1 - \tau_2) \{-L_\alpha \widehat{\text{Var}}(\widehat{\mathcal{K}}_{t,\tau_1}(x))^{1/2} + \widehat{\mathcal{K}}_{t,\tau_1}(x)\} \leq f_{t,\tau_2}(x) \leq \widehat{g}(\tau_1 - \tau_2) \{L_\alpha \widehat{\text{Var}}(\widehat{\mathcal{K}}_{t,\tau_1}(x))^{1/2} + \widehat{\mathcal{K}}_{t,\tau_1}(x)\}],$$





## Bootstrap

### Theorem

*Under regularity conditions*

$$[f_{t,\tau} : \sup_{x \in E} \{|\widehat{\mathcal{K}}_{t,\tau}(x) - f_{t,\tau}(x)| \text{Var}(\widehat{\mathcal{K}}_{t,\tau})^{-1/2}\} \leq L_{\alpha}^*]$$

*where the bound  $L_{\alpha}^*$  satisfies*

$$\begin{aligned} & \mathbb{P}^*(-\{U_{n_q}(x)^{-1}H_{n_q}^{-1}A_{n_q}^*(x)/B_{t,\tau}(x)^{-1}N(x)^{-1}M^*(x)N(x)^{-1}\}_{3,3} \\ & \leq L_{\alpha}^*) = 1 - \alpha \end{aligned}$$



## A Monte-Carlo Study

For  $q_t(x)$ , generate data from BS model, interest rate  $r = 0.04$ ,  $S_t = 6500$ ,  $k \in [6200, 7400]$ ,  $\tau = 1M$ ,  $\varepsilon_i \in U[0, 6]$ ,  $\sigma = 0.1878$ .

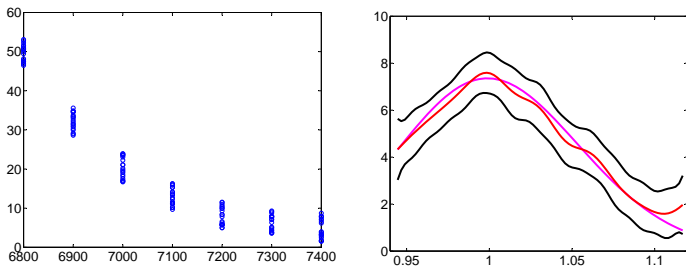


Figure 6: (Left)  $H$  against  $k$  (Right) Plot of confidence bands (black), **estimated value**, the Black Scholes SPD (magenta) of the EPK,  $h_{n_q} = 0.085$ ,  $\alpha = 0.05$ ,  $n_q = 300$ .



## A Monte-Carlo Study

For historical density, simulate data from Geometric Brownian Motion, with  $\mu = 0.23$ ,  $\sigma = 0.1878$ .

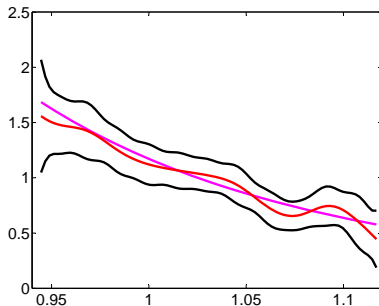


Figure 7: Plot of confidence bands, **estimated value**, the Black Scholes EPK (magenta),  $h_{n_q} = 0.060$ ,  $\alpha = 0.05$ ,  $n_q = 500$ ,  $n_p = 600$ .

Uniform Confidence for PKs



## Coverage Probability

Case	( $n =$ )300	450	600
( $\tau =$ )3	0.9063(2.402)	0.9144(2.204)	0.9233(1.998)
6	0.8964(2.438)	0.9056(2.134)	0.9203(2.069)

Table 1: Cov. prob. (area) of the uniform confidence band for  $q_t(x)$  at  $\alpha = 5\%$  with  $\sigma = 0.1878$ ,  $\text{sim} = 500$

Case	( $n =$ )300	450	600
3	0.7820(2.5434)	0.7980(2.4978)	0.8020(2.4131)
6	0.8602(2.4987)	0.8749(2.4307)	0.8900(2.4131)

Table 2: Same for EPK at  $\alpha = 10\%$



## Empirical Data Analysis

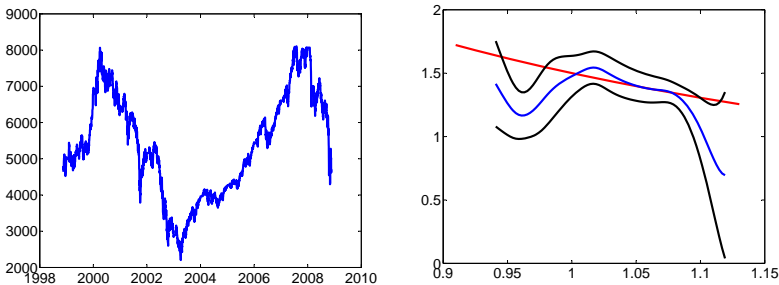


Figure 8: (Left) Plot of DAX index (Right) Plot of confidence bands (black), EPK by **Black Scholes fitting**, **nonparametric EPK**,  $h_{n_q} = 0.075$ ,  $\alpha = 0.05$ ,  $n_p = 506$ ,  $n_q = 715$ .



## Empirical Data Analysis

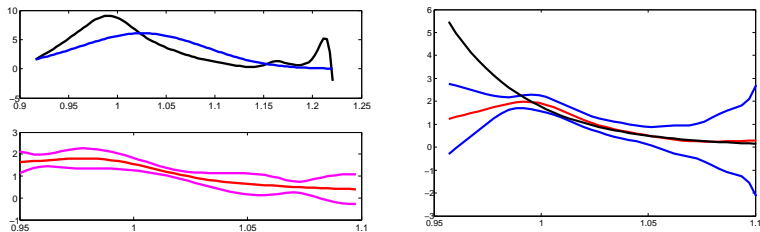


Figure 9: (Left) Plot of  $q$  and  $p$  (upper panel) Plot EPK and its bands, 2006 Feb 28th (lower panel) (Right) Plot of confidence bands (blue), EPK by Black Scholes fitting (black), **EPK**, 2006, April, 24th.



## Empirical Data Analysis

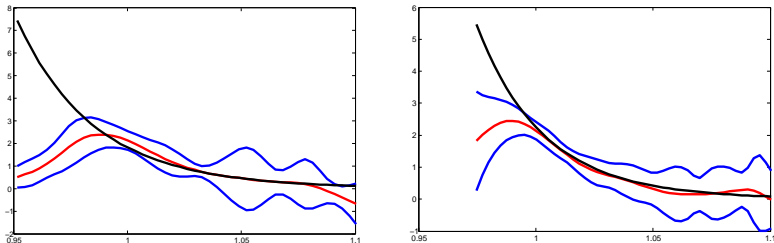


Figure 10: (Left) Plot of confidence bands (blue), EPK by Black Scholes fitting (black), **EPK**, 2006, July, 24th. (Right) Same for 2006, Aug, 18th.



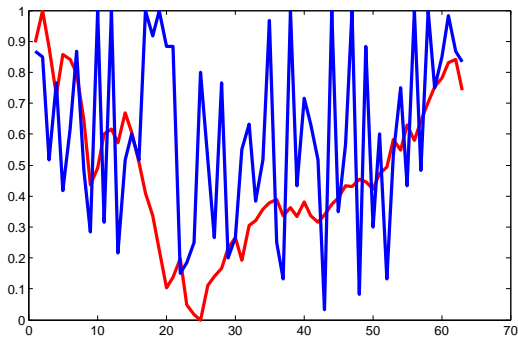


Figure 11: Plot of estimation of the BS EPK covered in band, DAX price (red)  $\tau = 2M.(200001-200006)$





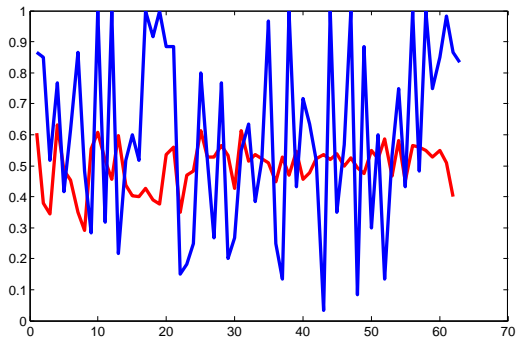


Figure 12: Plot of estimation of the BS EPK covered in band (blue), DAX price difference (red)  $\tau = 2M.$ (200001-200006)



## Conclusions

- Uniform confidence bands tell us about risk patterns
- Smoothing of EPK is best done via IVS
- Bootstrap does not improve coverage probability significantly
- BS for  $\tau = 0.5M$  is mostly rejected
- Bootstrap improvement possible for robust smoothers



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## Bibliography



Aït-Sahalia, Y., and A. W. Lo.

*Nonparametric Estimation of State-Price Densities Implicit in Financial Asset Prices*

Journal of Finance, 1998, 53, 499-547



Aït-Sahalia, Y., and A. W. Lo.

*Nonparametric Risk Management and Implied Risk Aversion*

Journal of Econometrics, 2000, 94, 9-51




Breeden, D. T. and Litzenberger, R. H.

*Prices of state-contingent claims implicit in option prices*

Journal of Business, 1978, 51, 621-651






## Bibliography

-  Detlefsen, K., Härdle, W., and Moro, R.  
*Empirical Pricing Kernels and Investor Preference*  
Mathematical Methods in Economics and Finance, forthcoming
-  Engle, R. F. and Rosenberg, J. V.  
*Empirical pricing kernels*  
Journal of Financial Economics, 2002, 64, 341-372
-  Claeskens, G. and Van Keilegom, I.  
*Bootstrap Confidence Bands for Regression Curves and Their Derivatives*  
The Annals of Statistics, 2003, 31, 1852-1884



## Bibliography

-  Giacomini, E., Handel, M. and Härdle, W.  
*Time Dependent Relative Risk Aversion*  
Risk Assessment: Decisions in Banking and Finance, Bol, G.,  
Rachev, S., Würth, R., Springer, 2008, 15-46
-  Härdle, W., Krätschmer, V. and Moro, R.  
*A Microeconomic Explanation of the EPK Paradox*  
Journal of Financial Econometrics, submitted on 18.09.09
-  Härdle, W. and Hlavka, Z.  
*Dynamics of state price densities*  
Journal of Econometrics, 2009, 150, 1-15



## Bibliography



Jackwerth, J.C.

*Risk Aversion from Option Prices and Realized Returns*  
The Review of Financial Studies, 2000, 13, 433-451



Leland, H. E.

*Who should buy portfolio insurance?*  
Journal of Finance, 1980, 32, 581-594



Rookley, C.

*Fully exploiting the information content of intra day option quotes: Applications in option pricing and risk management*  
University of Arizona working paper, 1997



## Rookley (1997)

Let  $C_{it}$  be the price of the  $i^{th}$  option at time  $t$  and  $K_{it}$  its strike price, and define the rescaled call option  $c = C/S_t$  in terms of moneyness  $M = S_t/K$  s.t.

$$\begin{aligned}c_{it} &= c\{M_{it}; \sigma(M_{it})\} = \Phi(d_1) - \frac{e^{-r\tau}\Phi(d_2)}{M_{it}} \\d_1 &= \frac{\log(M_{it}) + \left\{r_t + \frac{1}{2}\sigma(M_{it})^2\right\}\tau}{\sigma(M_{it})\sqrt{\tau}} \\d_2 &= d_1 - \sigma(M_{it})\sqrt{\tau}\end{aligned}$$





The RND is then

$$q(\cdot) = e^{r\tau} \frac{\partial^2 C}{\partial K^2} = e^{r\tau} S \frac{\partial^2 c}{\partial K^2}$$

with

$$\frac{\partial^2 c}{\partial K^2} = \frac{d^2 c}{dM^2} \left( \frac{M}{K} \right)^2 + 2 \frac{dc}{dM} \frac{M}{K^2}$$

and

$$\begin{aligned} \frac{d^2 c}{dM^2} = & \Phi'(d_1) \left\{ \frac{d^2 d_1}{dM^2} - d_1 \left( \frac{dd_1}{dM} \right)^2 \right\} \\ & - \frac{e^{-r\tau} \Phi'(d_2)}{M} \left\{ \frac{d^2 d_2}{dM^2} - \frac{2}{M} \frac{dd_2}{dM} - d_2 \left( \frac{dd_2}{dM} \right)^2 \right\} \\ & - \frac{2e^{-r\tau} \Phi(d_2)}{M^3} \end{aligned}$$



$$\text{With } V = \sigma(M), \quad V' = \frac{\partial \sigma(M)}{\partial M}, \quad V'' = \frac{\partial^2 \sigma(M)}{\partial M^2}$$

$$\begin{aligned} \frac{d^2 d_1}{dM^2} = & -\frac{1}{M * V(M) \sqrt{\tau}} \left\{ \frac{1}{M} + \frac{V'(M)}{V(M)} \right\} \\ & + V''(M) \left\{ \frac{\sqrt{\tau}}{2} - \frac{\log(M) + r\tau}{V(M)^2 \sqrt{\tau}} \right\} \\ & + V'(M) \left\{ 2V'(M) \frac{\log(M) + r\tau}{V(M)^3 \sqrt{\tau}} \right. \\ & \left. - \frac{1}{M * V(M)^2 \sqrt{\tau}} \right\} \end{aligned}$$



$$\begin{aligned}
 \frac{d^2 d_2}{dM^2} = & -\frac{1}{M * V(M)\sqrt{\tau}} \left\{ \frac{1}{M} + \frac{V'(M)}{V(M)} \right\} \\
 & - V''(M) \left\{ \frac{\sqrt{\tau}}{2} + \frac{\log(M) + r\tau}{V(M)^2 \sqrt{\tau}} \right\} \\
 & + V'(M) \left\{ 2V'(M) \frac{\log(M) + r\tau}{V(M)^3 \sqrt{\tau}} \right. \\
 & \left. - \frac{1}{M * V(M)^2 \sqrt{\tau}} \right\}
 \end{aligned}$$

