

TENET: Tail-Event-driven NETWORK Risk

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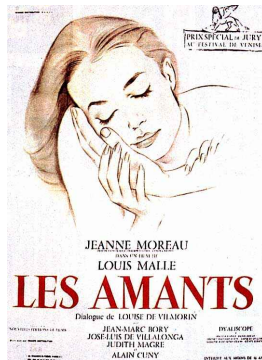
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What is Systemic Risk?

"I know it when I see it".

Justice Potter Stewart, 1964.



What is Systemic Risk?

Systemic risk is a "risk of financial instability so widespread that it impairs the functioning of a financial system to the point where economic growth and welfare suffer materially".

ECB, Financial Network and Financial Stability, 2010.

"Financial institutions are **systemically important** if the failure of the firm to meet its obligations to creditors and customers would have significant adverse consequences for the financial system and the broader economy".

Daniel Tarullo, Regulatory Restructuring, 2009.



What is Systemic Risk?



Figure 1: Systemic Risk?

CoVaR as a Systemic Risk Measure

Step 1. Estimate linear quantile regressions

$$X_{i,t} = \alpha_i + \gamma_i M_{t-1} + \varepsilon_{i,t},$$

$$X_{j,t} = \alpha_{j|i} + \gamma_{j|i} M_{t-1} + \beta_{j|i} X_{i,t} + \varepsilon_{j|i,t},$$

where

- $X_{i,t}$ is the log return of a financial institution i ,
- M_{t-1} are lagged macro state variables.

Adrian and Brunnermeier (2011)

▶ Macro state variables



CoVaR as a Systemic Risk Measure

Step 2. Generate predicted values under assumption

$$F_{\varepsilon_{i,t}}^{-1}(\tau|M_{t-1}) = 0 \text{ and } F_{\varepsilon_{j|i,t}}^{-1}(\tau|M_{t-1}, X_{i,t}) = 0, \tau = (0, 1),$$

$$\widehat{\text{VaR}}_{i,t}^{\tau} = \hat{\alpha}_i + \hat{\gamma}_i M_{t-1},$$

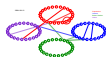
$$\widehat{\text{CoVaR}}_{j|i,t}^{\tau} = \hat{\alpha}_{j|i} + \hat{\gamma}_{j|i} M_{t-1} + \hat{\beta}_{j|i} \widehat{\text{VaR}}_{i,t}^{\tau}.$$

Adrian and Brunnermeier (2011)



Elements of Systemic Risk

- Network Effects
- Single Institution's Contribution to Systemic Risk
- Single Institution's Exposure to Systemic Risk



Challenges

- Linear tail behavior
 - ▶ Adrian and Brunnermeier (2011)
 - ▶ Acharya et al. (2012)
 - ▶ Brownlees and Engle (2012)

- Linear tail behavior in **high dimensions**
 - ▶ Hautsch, Schaumburg, and Schienle (2014)

- **Non-linear** tail behavior in **ultra-high dimensions**
 - ▶ Method by Fan, Härdle, Wang, and Zhu (2014)



Non-Linearity

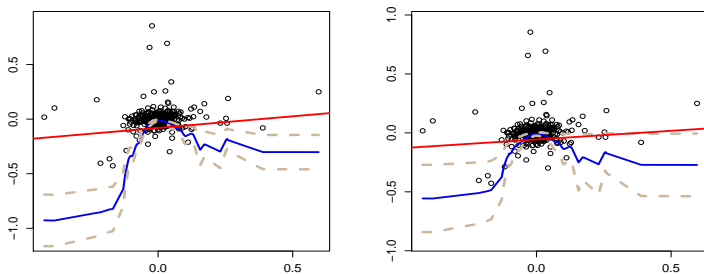
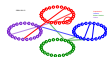


Figure 2: Bank of America (BOA) and Citi (C) weekly returns 0.05 (left) and 0.1 (right) quantile functions, y-axis = BOA returns, x-axis = C returns. **Local linear quantile regression** and **Linear quantile regression**. 95% confidence band, $T = 546$, weekly returns, 2005.01.31-2010.01.31. Chao, Härdle and Wang (2014).



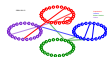
Outline

1. Motivation ✓
2. Statistical Methodology
3. Systemic Risk Modelling
4. Empirical Analysis
5. Conclusion
6. References



Model Components

- ▣ **Tail Behavior:** Generalized Quantile Regression
- ▣ **Non-Linearity:** Single-Index Model
- ▣ **Ultra-High Dimensions:** Variable Selection



Generalized Quantile Regression

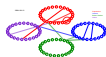
Let $\{X_i, Y_i\}_{i=1}^n$ be independent r. v., $X \in \mathbb{R}^p$, $\tau \in (0, 1)$.

$$Y_i = X_i^\top \theta + \varepsilon_i,$$
$$\hat{\theta} = \arg \min_{\theta \in \mathbb{R}^p} \sum_{i=1}^n \rho_\tau(Y_i - X_i^\top \theta),$$

where $\rho_\tau(\cdot)$ is an asymmetric loss function

$$\rho_\tau(u) = |u|^\alpha |\mathbf{1}(u \leq 0) - \tau|,$$

with $\alpha = 1$ corresponding to a quantile and $\alpha = 2$ corresponding to an expectile regression.



Asymmetric Loss Functions

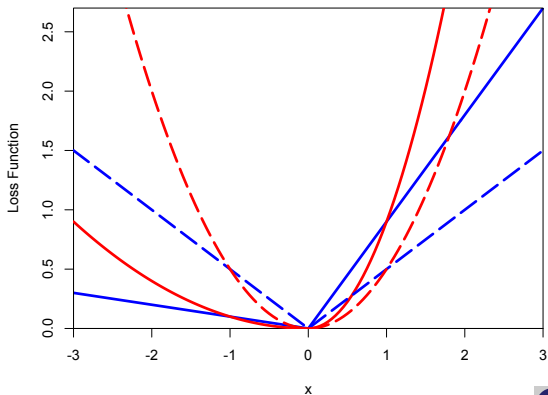
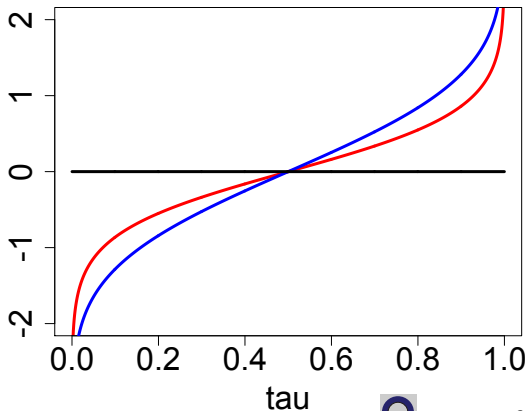


Figure 3: Asymmetric Loss Functions for **Quantile** and **Expectile**, $\tau = 0.9$: a solid line, $\tau = 0.5$: a dashed line.



Linear Quantile and Expectile



 SFScnfexpectile0.95

Figure 4: **Quantile** and **Expectile** for $N(0, 1)$.



Single-Index Model

Let $\{X_i, Y_i\}_{i=1}^n$ be independent r. v., $X \in \mathbb{R}^p$.

$$Y_i = g(\beta^\top X_i) + \varepsilon_i,$$

where

- $g(\cdot)$ is the link function,
- $\beta \in \mathbb{R}^p$ is the vector of index parameters,
- $p = \mathcal{O}\{\exp(n^\alpha)\}$ for some $\alpha \in (0, 1)$.



Estimation

Recall (1):

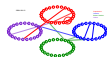
$$Y_i = g(\beta^\top X_i) + \varepsilon_i$$

A quasi-likelihood approach under assumption $F_{\varepsilon_i}^{-1}(\tau|X) = 0$

$$\min_{\beta \in \mathbb{R}^p} E \rho\{Y - g(\beta^\top X)\} \quad (1)$$

Further assumptions:

$\|\beta\|_2 = 1$ and first component of β is positive.



Estimation

Taylor approximation:

$$g(\beta^\top X_t) \approx g(\beta^\top x) + g'(\beta^\top x)\beta^\top(X_t - x) \quad (2)$$

Theoretically:

$$L_x(\beta) \stackrel{\text{def}}{=} \frac{E \rho\{Y - g(\beta^\top x) - g'(\beta^\top x)\beta^\top(X - x)\}}{K_h\{\beta^\top(X - x)\}} \quad (3)$$

Empirically:

$$L_{n,x}(\beta) \stackrel{\text{def}}{=} \frac{n^{-1} \sum_{t=1}^n \rho\{Y_t - g(\beta^\top x) - g'(\beta^\top x)\beta^\top(X_t - x)\}}{K_h\{\beta^\top(X_t - x)\}} \quad (4)$$

where $K_h(\cdot) = K(\cdot/h)/h$ with $K(\cdot)$ a kernel and h a bandwidth.



Minimum Average Contrast Estimation

$$\begin{aligned} L_n(\beta) &\stackrel{\text{def}}{=} n^{-1} \sum_{j=1}^n L_{n, X_j}(\beta) \\ &= n^{-2} \sum_{j=1}^n \sum_{t=1}^n \rho \left\{ Y_t - g(\beta^\top X_j) - g'(\beta^\top X_j) \beta^\top (X_t - X_j) \right\} \\ &\quad K_h \{ \beta^\top (X_t - X_j) \} \end{aligned} \quad (5)$$

$$\hat{\beta} \approx \arg \min_{\beta} L_n(\beta) \quad (6)$$

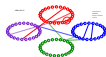


Variable Selection

$$\hat{\beta} = \arg \min_{g, g', \beta} n^{-1} \sum_{j=1}^n \sum_{t=1}^n \rho \left\{ Y_t - g(\beta^\top X_j) - g'(\beta^\top X_j) X_{tj}^\top \beta \right\} \omega_{tj}(\beta) \\ + \sum_{l=1}^p \gamma_\lambda(|\beta_l|^\theta),$$

where

- $X_{tj} = X_t - X_j,$
- $\omega_{tj}(\beta) \stackrel{\text{def}}{=} \frac{K_h(X_{tj}^\top \beta)}{\sum_{t=1}^n K_h(X_{tj}^\top \beta)},$
- $\theta \geq 0,$
- $\gamma_\lambda(t)$ is some nondecreasing function concave for $t \in [0, +\infty)$ with a continuous derivative on $(0, +\infty).$



Theory

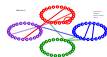
Denote $\hat{\beta}$ as the final estimate of β^* .

Theorem

*Under A 1-5, the estimators $\hat{\beta}^0$ and $\hat{\beta}$ exist and $P(\hat{\beta}^0 = \hat{\beta}) \rightarrow 1$.
Moreover,*

$$P(\hat{\beta}^0 = \hat{\beta}) \geq 1 - (p - q) \exp(-C' n^\alpha). \quad (7)$$

► Assumptions



Theory

Theorem

Under A 1-5, $\widehat{\beta}_{(1)} \stackrel{\text{def}}{=} (\widehat{\beta}_l)_{l \in \mathcal{M}_*}$, $b \in \mathbb{R}^q$, $\|b\| = 1$:

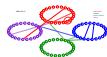
$$\|\widehat{\beta}_{(1)} - \beta_{(1)}^*\| = \mathcal{O}_p\{(\lambda D_n + n^{-1/2})\sqrt{q}\} \quad (8)$$

$$b^\top C_{0(1)}^{-1} \sqrt{n}(\widehat{\beta}_{(1)} - \beta_{(1)}^*) \xrightarrow{\mathcal{L}} N(0, \sigma^2) \quad (9)$$

where $\sigma^2 = E[\psi(\varepsilon_i)]^2 / [\partial^2 E \rho(\varepsilon_i)]^2$

$$\partial^2 E \rho(\cdot) = \left. \frac{\partial^2 E \rho(\varepsilon_i - v)^2}{\partial v^2} \right|_{v=0} \quad (10)$$

[Go to details](#)



Theory

Theorem

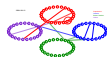
Under A 1-5, $\mathcal{B}_n \stackrel{\text{def}}{=} \{\hat{\beta} = \beta^*\} : P(\mathcal{B}_n) \rightarrow 1$. Let $\mu_j \stackrel{\text{def}}{=} \int u^j K(u) du$, $\nu_j \stackrel{\text{def}}{=} \int u^j K^2(u) du$, $j = 0, 1, \dots$. If $nh^3 \rightarrow \infty$ and $h \rightarrow 0$, then

$$\sqrt{nh} \sqrt{f_{Z(1)}(z) / (\nu_0 \sigma^2)} \left\{ \hat{g}(x^\top \hat{\beta}) - g(x^\top \beta^*) - \frac{1}{2} h^2 g''(x^\top \beta^*) \mu_2 \partial \mathbf{E} \psi(\varepsilon) \right\} \\ \xrightarrow{\mathcal{L}} \mathbf{N}(0, 1),$$

and

$$\sqrt{nh^3} \sqrt{\{f_{Z(1)}(z) \mu_2^2\} / (\nu_2 \sigma^2)} \left\{ \hat{g}'(x^\top \hat{\beta}) - g'(x^\top \beta^*) \right\} \xrightarrow{\mathcal{L}} \mathbf{N}(0, 1).$$

[Go to details](#)



Adaptive LASSO

$$\dots \sum_{l=1}^p \gamma_{\lambda}(|\beta_l|^{\theta}) = \lambda \sum_{l=1}^p w_l |\beta_l|,$$

where

- λ is a penalty term,
- $\theta = 1$,
- $w_l = 1/|\widehat{\beta}_l^0|^{\delta}$ are weights, $l = 1, \dots, p$, $\delta > 0$,
- $\widehat{\beta}^0$ is an initial estimator of β .

Zou (2006), Wu and Liu (2009)



Lambda

- Empirical choice of λ : $\lambda_n = 0.25\sqrt{\|\beta_0\|} \log n \vee p(\log n)^{0.5}$
- λ for ultra-high dimensions (Wang and Leng (2007))
- Schwarz Information Criterion (SIC)
(Schwarz (1978), Koenker, Ng, and Portnoy (1994))

$$\text{SIC}(\lambda) = \log[n^{-1} \sum_{i=1}^n \rho_{\tau}\{Y_i - f(X_i)\}] + \frac{\log n}{2n} \text{df}$$

where df is a measure of the effective dimensionality of the fitted model.

► Effective dimension



Bandwidth

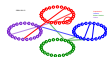
Symmetrized nearest neighbor estimation implies

$$\hat{m}_h(X_0) = (nh)^{-1} \sum_{i=1}^n Y_i K_h\{F_n(X_i) - F_n(x_0)\}$$

where

- $\hat{m}(x)$ denotes an estimator of the regression function,
- h is some bandwidth tending to zero.

Härdle and Carroll (1989)



Methodology of AB

- VaR: $\widehat{\text{VaR}}_{i,t,\tau} = \hat{\alpha}_i + \hat{\gamma}_i M_{t-1}$,
- $\widehat{\text{CoVaR}}^{\text{AB}}$: $\widehat{\text{CoVaR}}_{s|i,t,\tau} = \hat{\alpha}_{s|i} + \hat{\gamma}_{s|i} M_{t-1} + \hat{\beta}_{s|i} \widehat{\text{VaR}}_{i,t,\tau}$,
 - ▶ AB's information set: firm i 's VaR and macro state variables.
 - ▶ Systemic risk contribution: $\hat{\beta}_{s|i}$
- Limitations:
 - ▶ Linear assumption between a single firm and system.
 - ▶ Mechanical correlation between a single firm and the value-weighted system.



Methodology of TENET

$$\square \text{ VaR: } \widehat{\text{VaR}}_{i,t,\tau} = \hat{\alpha}_i + \hat{\gamma}_i M_{t-1},$$

$$\square \widehat{\text{CoVaR}}^{\text{TENET}} : \widehat{\text{CoVaR}}_{j|\tilde{R}_j,t,\tau} = \hat{g}(\hat{\beta}_{j|\tilde{R}_j}^{\text{T}} \tilde{R}_{j,t}),$$

- ▶ TENET's information set: internal factors, many other firms' VaRs and macro state variables.

- ▶ Spillover effects: $\hat{g}'(\hat{\beta}_{j|\tilde{R}_j}^{\text{T}} \tilde{R}_{j,t}) \hat{\beta}_{j|\tilde{R}_j}$.

$$\square \widehat{\text{CoVaR}}^{\text{SYSTEM}} : \widehat{\text{CoVaR}}_{s|\tilde{F}_j,t,\tau} = \hat{g}(\hat{\beta}_{s|\tilde{F}_j}^{\text{T}} \tilde{F}_{j,t}),$$

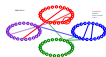
- ▶ SYSTEM's information set: firm j 's VaR, selected internal factors, selected other firms' VaRs and selected macro state variables.

- ▶ Systemic risk contribution: $\hat{g}'(\hat{\beta}_{s|\tilde{F}_j}^{\text{T}} \tilde{F}_{j,t}) \hat{\beta}_{s|\tilde{F}_j}$



Advantages of TENET

- Nonlinear structure.
- High dimensional setting with variable selection.
- Network dynamics.



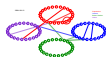
Step 1: VaR

Estimate linear QR

$$X_{i,t} = \alpha_i + \gamma_i M_{t-1} + \varepsilon_{i,t}, \quad (11)$$

$$\widehat{\text{VaR}}_{i,t,\tau} = \hat{\alpha}_i + \hat{\gamma}_i M_{t-1}, \quad (12)$$

- $X_{i,t}$ is the log-return of company i ,
- M_{t-1} are macro state variables as in Adrian and Brunnermeier (2011).



Step 2: Spillover Effects based Network

Estimate SIM-based QRs with variable selection

$$X_{j,t} = g(\beta_{j|R_j}^\top R_{j,t}) + \varepsilon_{j,t}, \quad (13)$$

$$\widehat{\text{CoVaR}}_{j|\tilde{R}_j}^{\text{TENET}} \stackrel{\text{def}}{=} \widehat{\text{CoVaR}}_{j|\tilde{R}_j,t,\tau}^{\text{SIM}} = \hat{g}(\hat{\beta}_{j|\tilde{R}_j}^\top \tilde{R}_{j,t}), \quad (14)$$

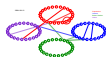
$$\hat{D}_{j|\tilde{R}_j} \stackrel{\text{def}}{=} \frac{\partial \hat{g}(\hat{\beta}_{j|\tilde{R}_j}^\top R_{j,t})}{\partial R_{j,t}} \Big|_{R_{j,t}=\tilde{R}_{j,t}} = \hat{g}'(\hat{\beta}_{j|\tilde{R}_j}^\top \tilde{R}_{j,t}) \hat{\beta}_{j|\tilde{R}_j} \quad (15)$$

- $R_{j,t} = \{X_{-j,t}, M_{t-1}, B_{j,t-1}\}$ the p dimensional information set.
- $X_{-j,t} = \{X_{1,t}, X_{2,t}, \dots, X_{k,t}\}$ log returns of all financial institutions except for a firm j , k : the number of financial institutions.
- $B_{j,t-1}$: the firm specific characteristics.



Step 2: Spillover Effects based Network

- $\beta_{j|R_j} \stackrel{\text{def}}{=} \{\beta_{j|-j}, \beta_{j|M}, \beta_{j|B_j}\}^\top$.
- $\tilde{R}_{j,t} \stackrel{\text{def}}{=} \{\widehat{\text{VaR}}_{-j,t,\tau}, M_{t-1}, B_{j,t-1}\}$.
- $\widehat{\text{VaR}}_{-j,t,\tau}$ are the estimated VaRs from (12) for financial institutions except for j in step 1.
- $\hat{\beta}_{j|\tilde{R}_j} \stackrel{\text{def}}{=} \{\hat{\beta}_{j|-j}, \hat{\beta}_{j|M}, \hat{\beta}_{j|B_j}\}^\top$.
- $\hat{D}_{j|\tilde{R}_j}$ is the gradient measuring the marginal effect of covariates evaluated at $R_{j,t} = \tilde{R}_{j,t}$, and the componentwise expression is $\hat{D}_{j|\tilde{R}_j} = \{\hat{D}_{j|-j}, \hat{D}_{j|M}, \hat{D}_{j|B_j}\}^\top$.
- $\hat{D}_{j|-j}$ allows to measure spillover effects across the financial institutions and to characterize their evolution as a system represented by a network.



Step 2: Total Connectedness Matrix

$$A_t = \begin{matrix} & l_{1,t} & l_{2,t} & l_{3,t} & \cdots & l_{k,t} \\ \begin{matrix} l_{1,t} \\ l_{2,t} \\ l_{3,t} \\ \vdots \\ l_{k,t} \end{matrix} & \left(\begin{array}{ccccc} 0 & |\widehat{D}_{1|2}| & |\widehat{D}_{1|3}| & \cdots & |\widehat{D}_{1|k}| \\ |\widehat{D}_{2|1}| & 0 & |\widehat{D}_{2|3}| & \cdots & |\widehat{D}_{2|k}| \\ |\widehat{D}_{3|1}| & |\widehat{D}_{3|2}| & 0 & \cdots & |\widehat{D}_{3|k}| \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ |\widehat{D}_{k|1}| & |\widehat{D}_{k|2}| & |\widehat{D}_{k|3}| & \cdots & 0 \end{array} \right) \end{matrix}$$

Table 1: A $k \times k$ adjacency matrix for financial institutions at time t .



Step 2: Network Measures

□ The firm level:

$$\blacktriangleright DC_{j|i,t} \stackrel{\text{def}}{=} |\hat{D}_{j|i}|$$

$$\blacktriangleright FC_{j,t}^{IN} \stackrel{\text{def}}{=} \sum_{i=1}^k |\hat{D}_{j|i}|$$

$$\blacktriangleright FC_{j,t}^{OUT} \stackrel{\text{def}}{=} \sum_{j=1}^k |\hat{D}_{j|i}|$$

□ The group level:

$$GC_{g,t}^{IN} \stackrel{\text{def}}{=} \sum_{i=1}^k \sum_{j \in g} |\hat{D}_{j|i}|, \quad GC_{g,t}^{OUT} \stackrel{\text{def}}{=} \sum_{i \in g} \sum_{j=1}^k |\hat{D}_{j|i}|$$

□ The overall level:

$$TC_t = TC_t^{IN} = TC_t^{OUT} \stackrel{\text{def}}{=} \sum_{i=1}^k \sum_{j=1}^k |\hat{D}_{j|i}|$$



Step 3: Systemic risk Contribution

Estimate SIM-based QRs without variable selection:

$$X_{s,t} = g(\beta_{s|F_j}^\top F_{j,t}) + \varepsilon_{s,t}, \quad (16)$$

$$\widehat{\text{CoVaR}}^{\text{SYSTEM}} \stackrel{\text{def}}{=} \widehat{\text{CoVaR}}_{s|\tilde{F}_j,t,\tau}^{\text{SIM}} = \widehat{g}(\widehat{\beta}_{s|\tilde{F}_j}^\top \tilde{F}_{j,t}), \quad (17)$$

$$\widehat{D}_{s|\tilde{F}_j} \stackrel{\text{def}}{=} \frac{\partial \widehat{g}(\widehat{\beta}_{s|\tilde{F}_j}^\top F_{j,t})}{\partial F_{j,t}} \Big|_{F_{j,t}=\tilde{F}_{j,t}} = \widehat{g}'(\widehat{\beta}_{s|\tilde{F}_j}^\top \tilde{F}_{j,t}) \widehat{\beta}_{s|\tilde{F}_j} \quad (18)$$

- $X_{s,t}$ refer to log returns of this financial system.

$$X_{s,t} = \frac{\sum_{i=1}^k X_{i,t} \cdot \text{Asset}_{i,t-1}}{\sum_{i=1}^k \text{Asset}_{i,t-1}}$$



Step 3: Systemic risk Contribution

- $F_{j,t} = \{X_{j,t}, C_{j,t}\}$, $C_{j,t} = \{X_{-j,t}^*, M_{t-1}^*, B_{j,t-1}^*\}$, $C_{j,t}$ includes control variables selected from step 2,
- $\beta_{s|F_j} = \{\beta_{s|j}, \beta_{s|C_j}\}^\top$.
- $\tilde{F}_{j,t} = \{\widehat{VaR}_{j,t,\tau}, \tilde{C}_{j,t}\}$, $\tilde{C}_{j,t} = \{\widehat{VaR}_{-j,t,\tau}^*, M_{t-1}^*, B_{j,t-1}^*\}$,
- $\hat{\beta}_{s|\tilde{F}_j} = \{\hat{\beta}_{s|j}, \hat{\beta}_{s|\tilde{C}_j}\}^\top$.
- $\hat{D}_{s|\tilde{F}_j} = \{\hat{D}_{s|j}, \hat{D}_{s|\tilde{C}_j}\}^\top$ is the partial derivative of system CoVaR with respect to the variables in $F_{j,t}$ evaluated at level $F_{j,t} = \tilde{F}_{j,t}$,
- $\hat{D}_{s|j}$ is the the partial derivative of system CoVaR with respect to institution j . In terms of identification of the system risk contributions we focus here on $\hat{D}_{s|j}$.



Dataset

- ▣ Asset log returns of 100 U.S. publicly traded financial firms.
- ▣ Firms classified by SIC codes: **Depositories (25)**, **Insurance (25)**, **Broker-Dealers (25)** and **Others (25)**.
- ▣ 4 firm specific characteristics: LEV, MM, MTB, SIZE.
- ▣ 7 macro state variables: VIX, 3MTB, LIQUIDITY, YIELD, CREDIT, D_J, S&P.
- ▣ Time period: January 5, 2007 - January 4, 2013, $T = 266$, $n = 48$.
- ▣ Frequency: weekly.

▸ Firms

▸ Macro state variables



Network Dynamics

Figure 5: Financial risk network dynamics Depositories, Insurance, Broker-Dealers, Others ; $T = 266$, $\tau = 0.05$, $n = 48$.

TENET



Network Analysis—Overall Level

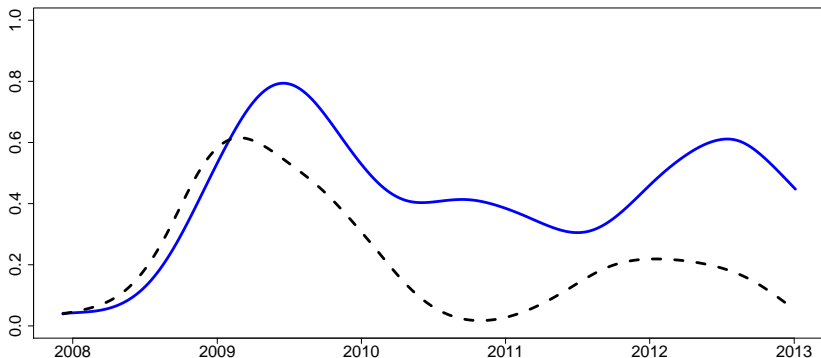
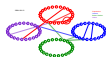


Figure 6: Total connectedness (solid line) and averaged λ of 100 financial institutions (dashes line): 20071207–20130105, both are standardized on [0, 1] scale.

Financial Risk Meter: <http://sfb649.wiwi.hu-berlin.de/frm/index.html>

TENET



Network Analysis–Group Level

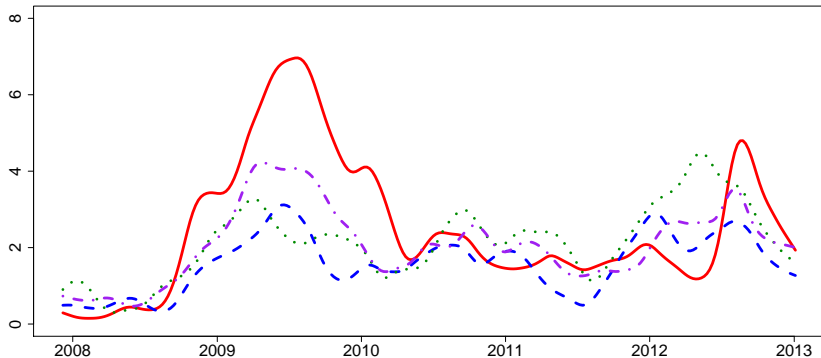


Figure 7: Incoming links for four industry groups. Depositories, Insurance, Broker-Dealers, Others ; $\tau = 0.05$, window size $n = 48$, $T = 266$.



Network Analysis—Group Level

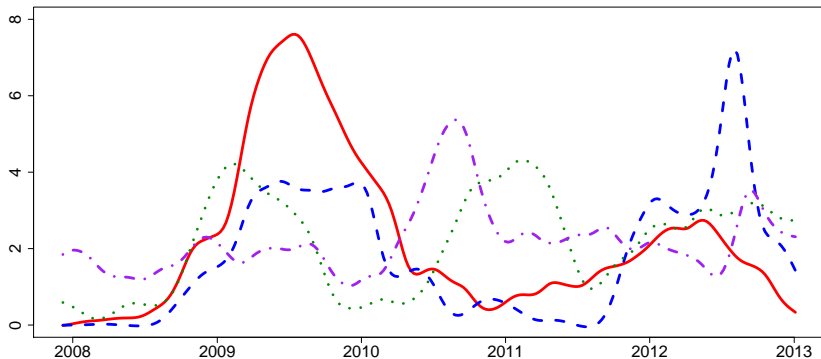
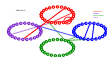


Figure 8: Outgoing links for four industry groups. Depositories, Insurance, Broker-Dealers, Others ; $\tau = 0.05$, window size $n = 48$, $T = 266$.



Network analysis—Firm Level

- Most connected institution wrt Incoming links:
Oppenheimer Holding, Inc. (OPY). ▶ IN-link
- Most connected institution wrt Outgoing links:
Lincoln National Corporation (LNC). ▶ OUT-link
- Directional most connected institutions:
from NewStar Financial, Inc. (NEWS) to Oppenheimer
Holdings, Inc.(OPY). ▶ DIRECT-link



Summary of Network analysis

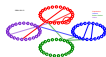
- The connections between institutions tend to increase before the financial crisis.
- The connections between institutions get weaker as the financial system stabilized.
- Whereas banks dominate both incoming and outgoing links, the insurers are less affected by the financial crisis and exhibit less contribution in terms of risk transmission.
- Several institutions with moderate or small sizes and also some non bank institutions received or transmitted more risk, as there are "too connected" firms.



Systemic Risk Contribution

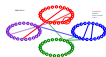
Ranking of SRC	Ticker	Averaged Sum	Ranking of MC
1	JPM	0.27	2
2	BAC	0.26	3
3	WFC	0.23	1
4	C	0.19	4
5	PRU	0.17	13
6	L	0.16	28
7	GS	0.13	7
8	MET	0.12	9
9	MTB	0.11	33
10	AXP	0.10	5

Table 2: Top 10 financial institutions ranked according to the systemic risk contribution (SRC) calculated by the averaged sum of partial derivatives, and the Ranking of market capitalization (MC) in this 100 financial institutions' list is also shown in this table.



Link Function Dynamics

Figure 9: Link function dynamics for JPM, 5th January 2007 - 30th December 2011, $\tau = 0.05$, window size $n = 48$.



Conclusion

- Network can identify the interconnectedness among financial institutions.
- Partial Derivative of Systemic CoVaR can find the systemic risk contributors.
- Both major systemic risk contributors and major interconnected companies are systemically important.



TENET: Tail Event driven NETWORK risk

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<http://www.case.hu-berlin.de>



Expectile-Quantile Correspondence

Let $v(x)$ represents expectile regression, $I(x)$ represents quantile regression.

Fixed x , define $w(\tau)$ such that $v_{w(\tau)}(x) = I(x)$ then $w(\tau)$ is related to $I(x)$ via

$$w(\tau) = \frac{\tau I(x) - \int_{-\infty}^{I(x)} y dF(y|x)}{2 E(Y|x) - 2 \int_{-\infty}^{I(x)} y dF(y|x) - (1 - 2\tau)I(x)}$$

For example, $Y \sim U(-1, 1)$, then $w(\tau) = \tau^2 / (2\tau^2 - 2\tau + 1)$

Expectile corresponds to quantile with transformation w .

[Return](#)



Numerical procedure

- Given $\widehat{\beta}^{(t)}$, standardize $\widehat{\beta}^{(t)}$ so that $\|\widehat{\beta}^{(t)}\| = 1$, $\widehat{\beta}_1^{(t)} > 0$.
Then compute

$$(\widehat{a}_j^{(t)}, \widehat{b}_j^{(t)}) \stackrel{\text{def}}{=} \arg \min_{(a_j, b_j)'s} \sum_{i=1}^n \rho(Y_i - a_j - b_j X_{ij}^\top \widehat{\beta}^{(t)}) \omega_{ij}(\widehat{\beta}^{(t)}),$$

where

- $\widehat{\beta}_0$ initial estimator of β^* ,
- $X_{ij} = X_i - X_j$,
- $a_j = g(\beta^\top X_j)$,
- $b_j = g'(\beta^\top X_j)$,
- $\omega_{ij}(\widehat{\beta}_0^{(t)}) \stackrel{\text{def}}{=} \frac{K_h(X_{ij}^\top \beta_0^{(t)})}{\sum_{i=1}^n K_h(X_{ij}^\top \beta_0^{(t)})}$,
- $t = 1, 2, \dots$ are iterations.



Numerical procedure

2. Given $(\hat{a}_j^{(t)}, \hat{b}_j^{(t)})$, solve

$$\hat{\beta}^{(t+1)} = \arg \min_{\beta} n^{-1} \sum_{j=1}^n \sum_{i=1}^n \rho(Y_i - \hat{a}_j^{(t)} - \hat{b}_j^{(t)} X_{ij}^T \beta) \omega_{ij}(\hat{\beta}^{(t)}),$$
$$+ \sum_{l=1}^p \hat{d}_l^{(t)} |\beta_l|.$$

where

- $\hat{d}_l^{(t)} = \gamma_\lambda(|\hat{\beta}_l^{(t)}|)$,
- $\omega_{ij}(\cdot)$ are from the step before.

▶ Return



Effective dimension

Let $\{X_i, Y_i\}_{i=1}^n$ be independent r. v.

Given X , let $Y_i \sim (\mu(X), \sigma^2)$, where $\mu(X)$ is the true mean and σ^2 is the common variance.

$$df(\hat{f}) = \sum_{i=1}^n \frac{\text{Cov}\{\hat{f}(X_i), Y_i\}}{\sigma^2}.$$

Under certain mild conditions an unbiased estimator of df is

$$df(\hat{f}) = \sum_{i=1}^n \frac{\partial \hat{f}(X_i)}{\partial Y_i}$$

Stein (1981)



Assumptions

A1 K a cts symmetric pdf, $g(\cdot) \in C^2$.

A2 $\rho(x)$ convex. Suppose $\psi(x)$, subgradient of $\rho(x)$:

i) Lipschitz continuous; ii) $E \psi(\varepsilon_i) = 0$ and
 $\inf_{|v| \leq c} \partial E \psi(\varepsilon_i - v) = C_1$.

A3 ε_i is independent of X_i . Let $Z_i = X_i^\top \beta^*$ and $Z_{ij} = Z_i - Z_j$.

$C_{0(1)} \stackrel{\text{def}}{=} E\{g'(Z_i)^2 (E(X_{i(1)}|Z_i) - X_{i(1)})(E(X_{i(1)}|Z_i - X_{i(1)}))\}^\top$,
 and the matrix $C_{0(1)}$ satisfies

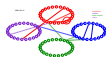
$0 < L_1 \leq \lambda_{\min}(C_{0(1)}) \leq \lambda_{\max}(C_{0(1)}) \leq L_2$ for positive
 constants L_1 and L_2 . There exists a constant $c_0 > 0$ such that
 $\sum_{i=1}^n \{\|X_{i(1)}\|/\sqrt{n}\}^{2+c_0} \xrightarrow{P} 0$, with $0 < c_0 < 1$. Also

$\|\sum_i \sum_j X_{(0)ij} \omega_{ij} X_{(1)ij}^\top \partial E \psi(v_{ij})\|_{2,\infty} = \mathcal{O}_p(n^{1-\alpha_1})$.



Assumptions

- A4** The penalty parameter λ is chosen such that $\lambda D_n = \mathcal{O}\{n^{-1/2}\}$, with $D_n \stackrel{\text{def}}{=} \max\{d_l : l \in \mathcal{M}_*\} = \mathcal{O}(n^{\alpha_1 - \alpha_2/2})$, $d_l \stackrel{\text{def}}{=} \gamma_\lambda(|\beta_l^*|)$, $\mathcal{M}_* = \{l : \beta_l^* \neq 0\}$ be the true model. Furthermore assume $qh \rightarrow 0$ as $n \rightarrow \infty$, $q = \mathcal{O}(n^{\alpha_2})$, $p = \mathcal{O}(\exp\{n^\delta\})$, $nh^3 \rightarrow \infty$ and $h \rightarrow 0$. Also, $0 < \delta < \alpha < \alpha_2/2 < 1/2$, $\alpha_2/2 < \alpha_1 < 1$. For example, $\delta = 1/5$, $\alpha = 1/4$, $\alpha_2 = 3/5$, $\alpha_1 = 3/5$.
- A5** The error term ε_j satisfies $E\varepsilon_j = 0$ and $\text{Var}(\varepsilon_j) < \infty$. Assume that $E|\psi^m(\varepsilon_j)/m!| \leq s_0 c^m$ where s_0 and c are constants.

[Return](#)

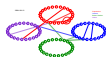
Subgradient

If $f : U \rightarrow \mathbb{R}$ is a real-valued convex function defined on a convex open set in the Euclidean space \mathbb{R}^n , a vector v in that space is called a subgradient at a point x_0 in U if for any x in U one has

$$f(x) - f(x_0) \geq v \cdot (x - x_0)$$

where the dot denotes the dot product.

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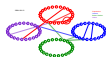


Matrix norm

Assume A is a $m \times n$ matrix

$$\|A\|_{\alpha,\beta} = \max_{x \neq 0} \frac{\|Ax\|_{\beta}}{\|x\|_{\alpha}}$$

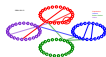
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Sparsistency

The result of (7) is stronger than the oracle property defined in Fan and Li (2001) once the properties of $\hat{\beta}^0$ are established. It was formulated by Kim et al. (2008) for the SCAD estimator with polynomial dimensionality p . It implies not only the model selection consistency and but also sign consistency (Zhao and Yu, 2006; Bickel et al., 2008, 2009):

$$P\{\text{sgn}(\hat{\beta}) = \text{sgn}(\beta^*)\} = P\{\text{sgn}(\hat{\beta}^0) = \text{sgn}(\beta^*)\} \rightarrow 1$$

[▶ Return](#)

The confidence interval

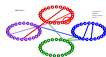
The $100(1 - \alpha)\%$ confidence interval:

$$\left[\hat{g}(z) - \frac{1}{\sqrt{nh}} \cdot \frac{\sigma\sqrt{\nu_0}}{\sqrt{\hat{f}_{Z(1)}(z)}} \cdot \mathfrak{z}_\alpha + \frac{1}{2}h^2\hat{g}''(z)\mu_2\partial\hat{E}\psi(\varepsilon); \right. \\ \left. \hat{g}(z) + \frac{1}{\sqrt{nh}} \cdot \frac{\sigma\sqrt{\nu_0}}{\sqrt{\hat{f}_{Z(1)}(z)}} \cdot \mathfrak{z}_\alpha + \frac{1}{2}h^2\hat{g}''(z)\mu_2\partial\hat{E}\psi(\varepsilon) \right]$$

where \mathfrak{z}_α is the α -Quantile of the standard normal distribution, and

$$\hat{f}_{Z(1)}(z) = n^{-1} \sum_{i=1}^n K_h(z - Z_{i(1)}), \text{ where } Z_{i(1)} = X_{i(1)}^\top \hat{\beta}_{(1)}.$$

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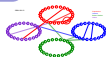


Network Analysis: IN-link

Ranking of IN-link	Ticker	Total IN Sum	Ranking of MC
1	OPY	68.63	98
2	IVZ	67.54	35
3	SFE	65.38	93
4	FITB	64.64	30
5	KEY	64.01	40
6	JPM	54.81	2
7	WFC	50.31	1
8	ZION	48.95	63
9	COF	48.36	10
10	STI	47.41	29

Table 3: Top 10 financial institutions ranked according to Incoming links calculated by the sum of absolute value of the partial derivatives, and the Ranking of market capitalization (MC) in this 100 financial institutions' list is also shown in this table.

[Return](#)

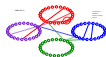


Network Analysis: OUT-link

Ranking of OUT-link	Ticker	Total Out Sum	Ranking of MC
1	LNC	260.72	43
2	C	174.46	4
3	LTS	164.48	97
4	MS	163.91	12
5	CBG	121.48	32
6	AGM	114.38	89
7	FITB	97.21	30
8	RF	84.65	36
9	ZION	84.52	63
10	NNI	80.87	77

Table 4: Top 10 financial institutions ranked according to Outgoing links calculated by the sum of absolute value of the partial derivatives, and the Ranking of market capitalization (MC) in this 100 financial institutions' list is also shown in this table.

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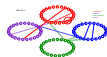


Network Analysis: DIRECT-link

Ranking of Sum	From Ticker	To Ticker	Total Sum
1	NEWS	OPY	33.06
2	LNC	CBG	32.74
3	C	MS	28.26
4	RF	STI	23.72
5	C	BAC	22.99
6	LNC	SFE	17.61
7	MS	LM	16.82
8	C	OPY	16.36
9	CBG	JLL	15.54
10	LNC	CLMS	15.34

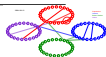
Table 5: Top 10 directional connectedness from one financial institution to another. The ranking is calculated by the sum of absolute value of the partial derivatives.

▶ Return



Financial firms

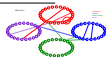
Depositories (25)		Insurances (25)	
WFC	Wells Fargo & Company	AIG	American International Group, Inc.
JPM	J P Morgan Chase & Co	MET	MetLife, Inc.
BAC	Bank of America Corporation	TRV	The Travelers Companies, Inc.
C	Citigroup Inc.	AFL	Aflac Incorporated
USB	U.S. Bancorp	PRU	Prudential Financial, Inc.
COF	Capital One Financial Corporation	CB	Chubb Corporation (The)
PNC	PNC Financial Services Group, Inc. (The)	MMC	Marsh & McLennan Companies, Inc.
BK	Bank Of New York Mellon Corporation (The)	ALL	Allstate Corporation (The)
STT	State Street Corporation	AON	Aon plc
BBT	BB&T Corporation	L	Loews Corporation
STI	SunTrust Banks, Inc.	PGR	Progressive Corporation (The)
FITB	Fifth Third Bancorp	HIG	Hartford Financial Services Group, Inc. (The)
MTB	M&T Bank Corporation	PFG	Principal Financial Group Inc
NTRS	Northern Trust Corporation	CNA	CNA Financial Corporation
RF	Regions Financial Corporation	LNC	Lincoln National Corporation
KEY	KeyCorp	CINF	Cincinnati Financial Corporation
CMA	Comerica Incorporated	Y	Alleghany Corporation
HBAN	Huntington Bancshares Incorporated	UNM	Unum Group
HCBC	Hudson City Bancorp, Inc.	WRB	W.R. Berkley Corporation
PBCT	People's United Financial, Inc.	FNF	Fidelity National Financial, Inc.
BOKF	BOK Financial Corporation	TMK	Torchmark Corporation
ZION	Zions Bancorporation	MKL	Markel Corporation
CFR	Cullen/Frost Bankers, Inc.	AJG	Arthur J. Gallagher & Co.
CBSH	Commerce Bancshares, Inc.	BRO	Brown & Brown, Inc.
SBNY	Signature Bank	HCC	HCC Insurance Holdings, Inc.



Financial firms

[▶ Return](#)

Broker-Dealers (25)		others (25)	
GS	Goldman Sachs Group, Inc. (The)	AXP	American Express Company
BLK	BlackRock, Inc.	BEN	Franklin Resources, Inc.
MS	Morgan Stanley	CBG	CBRE Group, Inc.
CME	CME Group Inc.	IVZ	Invesco Plc
SCHW	The Charles Schwab Corporation	JLL	Jones Lang LaSalle Incorporated
TROW	T. Rowe Price Group, Inc.	AMG	Affiliated Managers Group, Inc.
AMTD	TD Ameritrade Holding Corporation	OCN	Ocwen Financial Corporation
RJF	Raymond James Financial, Inc.	EV	Eaton Vance Corporation
SEIC	SEI Investments Company	LM	Legg Mason, Inc.
NDAQ	The NASDAQ OMX Group, Inc.	CACC	Credit Acceptance Corporation
WDR	Waddell & Reed Financial, Inc.	FII	Federated Investors, Inc.
SF	Stifel Financial Corporation	AB	Alliance Capital Management Holding L.P.
GBL	Gamco Investors, Inc.	PRAA	Portfolio Recovery Associates, Inc.
MKTX	MarketAxess Holdings, Inc.	JNS	Janus Capital Group, Inc
EEFT	Euronet Worldwide, Inc.	NNI	Nelnet, Inc.
WETF	WisdomTree Investments, Inc.	WRLD	World Acceptance Corporation
DLLR	DFC Global Corp	ECPG	Encore Capital Group Inc
BGCP	BGC Partners, Inc.	NEWS	NewStar Financial, Inc.
PJC	Piper Jaffray Companies	AGM	Federal Agricultural Mortgage Corporation
ITG	Investment Technology Group, Inc.	WHG	Westwood Holdings Group Inc
INTL	INTL FCStone Inc.	AVHI	AV Homes, Inc.
GFIG	GFI Group Inc.	SFE	Safeguard Scientifics, Inc.
LTS	Ladenburg Thalmann Financial Services Inc	ATAX	America First Tax Exempt Investors, L.P.
OPY	Oppenheimer Holdings, Inc.	TAXI	Medallion Financial Corp.
CLMS	Calamos Asset Management, Inc.	NICK	Nicholas Financial, Inc.



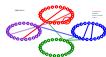
Macro state variables

-
-
1. VIX
 2. Short term liquidity spread (liquidity)
 3. Daily change in the 3-month Treasury maturities (3MT)
 4. Change in the slope of the yield curve (yield)
 5. Change in the credit spread (credit)
 6. Daily Dow Jones U.S. Real Estate index returns (D_J)
 7. S&P500 returns (S&P)
-
-




Source: Adrian and Brunnermeier (2011), Datastream.

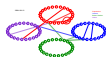
▶ [Return to Introduction](#)

▶ [Return to Empirical Analysis](#)






References

-  Acharya, V. and Engle, R. and Richardson, M.
Capital shortfall: A new approach to ranking and regulating systemic risks.
The American Economic Review,102(3): 59-64. 2012.
-  Adrian, T. and Brunnermeier, M. K.
CoVaR.
Staff Reports 348, Federal Reserve Bank of New York, 2011.
-  Beale, N., Rand, D. G., Battey, H., Croxson, K., May, R. M., and Nowak, M. A.
Individual versus systemic risk and the regulator's dilemma.
Proceedings of the National Academy of Sciences,
108(31):12647-12652, 2011.



References

-  Belloni, A. and Chernozhukov, V.
L1-penalized quantile regression in highdimensional sparse models.
The Annals of Statistics, 39(1):82-130, 2011.
-  Berkowitz, J., Christoffersen, P. and Pelletier, D.
Evaluating value-at-risk models with desk-level data.
Management Science, 57(12):2213-2227, 2011.
-  Billio, M., Getmansky, M., Lo, A. W., and Pelizzon, L.
Econometric measures of connectedness and systemic risk in the finance and insurance sectors.
Journal of Financial Economics, 104(3):535-559, 2012.



References



Bisias, D., Flood, M., Lo, A. W., and Valavanis
A Survey of Systemic Risk Analytics.
Annual Review of Financial Economics, Vol. 4: 255-296, 2012.



Borisov, I. and Volodko, N.

Exponential inequalities for the distributions of canonical u-and
v-statistics of dependent observations,
Siberian Advances in Mathematics,19(1):1-12, 2009.






Boss, M., Krenn, G., Pühr, C., and Summer, M.

Systemic risk monitor: A model for systemic risk analysis and
stress testing of banking systems.
Financial Stability Report, 11:83-95, 2006.



References

-  Brownlees, C. T. and Engle, R. F.
SRISK: A Conditional Capital Shortfall Index for Systemic Risk Measurement.
[Available at SSRN 1611229, 2015.](#)
-  Carroll, R. J., and Härdle, W. K.
Symmetrized nearest neighbor regression estimates.
[Statistics and Probability Letters, 7\(4\): 315-318, 1989.](#)
-  Chan-Lau, J., Espinosa, M., Giesecke, K., and Solé, J.
Assessing the systemic implications of financial linkages.
[IMF Global Financial Stability Report, 2. 2009.](#)



References



Chao, S. K., Härdle, W. K. and Wang, W.

Quantile regression in Risk Calibration.

Handbook of Financial Econometric and Statistics, pages 1467-1489, 2015.



Diebold, F. X. and Yilmaz, K.

On the network topology of variance decompositions:

Measuring the connectedness of financial firms.

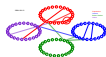
Journal of Econometrics, 182:119-134, 2014.



Fan, Y., Härdle, W. K., Wang, W., and Zhu, L.

Composite quantile regression for the single-index model.

Revised and resubmitted to *Journal of Business and Economic Statistics*, 2013.



References



Fan, J. and Li, R.

Variable selection via nonconcave penalized likelihood and its oracle properties.

J. Amer. Statist. Assoc. 96: 1348-1360, 2001.



Franke, J., Mwita, P., and Wang, W.

Nonparametric estimates for conditional quantiles of time series.

AStA Advances in Statistical Analysis, pages 1-24, 2014.






Gertler, M. and Kiyotaki, N.

Financial intermediation and credit policy in business cycle analysis.

Handbook of monetary economics, 3(11): 547-599, 2010.






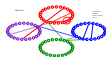
References

-  Giglio, S., Kelly, B., Pruitt, S., and Qiao, X.
Systemic risk and the macroeconomy: An empirical evaluation.
Fama-Miller Working Paper, 2012.
-  Härdle, W. K., Müller, M., Sperlich, S., and Werwatz, A.
Nonparametric and semiparametric models.
Springer, 2004.
-  Hautsch, N., Schaumburg, J. and Schienle, M.
Financial network systemic risk contributions.
Review of Finance, 19(2): 685-738, 2015.



References

-  Huang, X., Zhou, H., and Zhu, H.
A framework for assessing the systemic risk of major financial institutions.
Journal of Banking and Finance, 33(11): 2036-2049, 2009.
-  Koenker, R., Ng, P., and Portnoy, S.
Quantile smoothing splines.
Biometrika, 81(4): 673-680, 1994.
-  Kong, E., Linton, O., and Xia, Y.
Uniform bahadur representation for local polynomial estimates of m-regression and its application to the additive model.
Econometric Theory, 26(05):1529-1564, 2010.



References



Lehar, A.

Measuring systemic risk: A risk management approach.
Journal of Banking and Finance, 29(10): 2577-2603,
2005.



Li, Y. and Zhu, J.

L1- norm quantile regression.
Journal of Computational and Graphical Statistics, 17:
163-185, 2008.






Minsky, H. P.

A theory of systemic fragility.
*Financial crises: Institutions and markets in a fragile
environment*, pages 138-152, 1977.






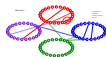
References

-  Rodriguez-Moreno, M. and Pe na, J. I.
Systemic risk measures: The simpler the better?
Journal of Banking and Finance, 37(6):1817-1831, 2013.
-  Schwarz, G.
Estimating the dimension of a model.
The annals of statistics, 6(2):461-464, 1978.
-  Tibshirani, R.
Regression shrinkage and selection via the lasso.
Journal of the Royal Statistical Society, B 58(1): 267-288,
1996.



References

-  Wang, H., Li, R., and Tsai, C.-L.
Tuning parameter selectors for the smoothly clipped absolute deviation method.
Biometrika, 94(3):553-568, 2007.
-  Wu, T. Z., Yu, K. and Yu, Y.
Single-index quantile regression.
Journal of Multivariate Analysis, 101(7):1607-1621, 2010.
-  Yu, K., Lu, Z., and Stander, J.
Quantile regression: applications and current research areas.
Journal of the Royal Statistical Society: Series D (The Statistician), 52(3):331-350, 2003.



References



Zheng, Q., Gallagher, C., and Kulasekera, K.
Adaptive penalized quantile regression for high dimensional data.

Journal of Statistical Planning and Inference,
143(6):1029-1038, 2013.



Zou, H.
The adaptive Lasso and its oracle properties.

Journal of the American statistical association, 101(476):
1418-1429, 2006.

