## Credit Risk Calibration based on CDS Spreads

Shih-Kang Chao
Wolfgang Karl Härdle
Hien Pham-Thu

Ladislaus von Bortkiewicz Chair of Statistics

C.A.S.E. - Center for Applied Statistics and Economics
Humboldt-Universität zu Berlin
http://lvb.wiwi.hu-berlin.de


## The impact of the subprime crisis



Lehman collapse sends
Financial
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risis on Wall Street as Lehman Totters Financial Crisic derillis Sold, AGG Seeks to Raise Cash be eon markets breakdown 1.2

 SOUVERFIGN DERT CRISIS: Pus debt growing tears sees brink of a default turmoil in bond markets if FPticiasmets reopen monday Frustration has grown debt downgradess banks on the edge of insolvencs Stock market crash! tr. ol banking inaustr.

## The consequences out of the financial crisis

Innocent \& not involved?


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## The Concept of Central Counterparty (CCP)

Central Counterparty interposes itself between counterparties and becomes the buyer to every seller and the seller to every buyer.


## Risk Mangement of CCP

## Main focus: credit risk



## Credit Risk Calibration by CCP

Is CCP in the position to monitoring the spillover of credit risk by its members?


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## Credit Risk Calibration: How to measure credit risk spillover effects?

High upward and downward co-movements in CDS spreads during the period 2007-2009.


## Risk measures

$\square$ Value at Risk (VaR)

$$
\operatorname{VaR}_{t+d}^{\alpha}=\inf \left\{x \in \mathbb{R}: \mathrm{P}\left(X_{t+d} \leq x \mid \mathcal{F}_{t}\right) \geq \alpha\right\}
$$

where $X_{t}=-\log \left(\frac{S_{t}}{S_{t-1}}\right)$ denotes the CDS spread $\log$ returns.

## Objectives

$\square$ Marginal credit risk analysis tool based on CDS spreads
$\square$ Measure of interconnectedness: quantification of mutual effects of credit risk
$\checkmark$ Relationship between CDS spreads in tail events: linear or non-linear?
$\square$ Uncover the relationship between CDS spreads and CDS determinants

## Outline

1. Motivation $\checkmark$
2. Linear quantile regression
3. PLM Methodology
4. Empirical study
5. Conclusions

## Linear Quantile Regression

$$
\begin{aligned}
& X_{i, t}=\alpha_{i}+\gamma_{i}^{\top} M_{t-1}+\varepsilon_{i, t} \\
& X_{j, t}=\alpha_{j \mid i}+\beta_{j \mid i} X_{i, t}+\gamma_{j \mid i}^{\top} M_{t-1}+\varepsilon_{j, t}
\end{aligned}
$$

$M_{t}$ : state variables. $F_{\varepsilon_{i, t}}^{-1}\left(\tau \mid M_{t-1}\right)=0$ and $F_{\varepsilon_{j, t}}^{-1}\left(\tau \mid M_{t-1}, X_{i, t}\right)=0$.

$$
\begin{aligned}
\widehat{V a R}_{i, t} & =\hat{\alpha}_{i}+\hat{\gamma}_{i}^{\top} M_{t-1} \\
\widehat{\operatorname{CoVaR}}_{j \mid i, t} & =\hat{\alpha}_{j \mid i}+\hat{\beta}_{j \mid i} \widehat{V a R}_{i, t}+\hat{\gamma}_{j \mid i}^{\top} M_{t-1}
\end{aligned}
$$

Systemic contribution of $i$ on $j$ :

$$
\triangle \widehat{\operatorname{CoVaR}}_{j \mid i, t}=\widehat{\operatorname{CoVaR}}_{j \mid i, t}-\widehat{\operatorname{CoVaR}}_{j \mid X_{i}=\text { Median }, t}
$$

See Adrian \& Brunnermeier (2011): CoVaR (AB (2011))
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Figure 1: Quantile regression at 0.01 level on CDS spread return. Linear quantile regression line. Partial linear quantile regression estimation. The dashed lines express the asymptotic and bootstrap confidence bands at 95\% level.
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## Partial Linear Quantile Regression:

$$
\begin{aligned}
& X_{i, t}=\alpha_{i}+\gamma_{i}^{\top} M_{t-1}+\varepsilon_{i, t} \\
& X_{j, t}=\tilde{\alpha}_{j \mid i}+\tilde{\beta}_{j \mid i}^{\top} M_{t-1}+I_{j \mid i}\left(X_{i, t}\right)+\varepsilon_{j, t}
\end{aligned}
$$

$I$ : a general function. $M_{t}$ : state variables. $F_{\varepsilon i, t}^{-1}\left(\tau \mid M_{t-1}\right)=0$ and $F_{\varepsilon_{j, t}}^{-1}\left(\tau \mid M_{t-1}, X_{i, t}\right)=0$.

$$
\begin{aligned}
\widehat{\operatorname{VaR}}_{i, t} & =\hat{\alpha}_{i}+\hat{\gamma}_{i}^{\top} M_{t-1}, \\
\widehat{\operatorname{CoVaR}}_{j \mid i, t} & =\hat{\alpha}_{j \mid i}+\hat{\gamma}_{j \mid i}^{\top} M_{t-1}+\hat{l}_{j \mid i}\left(\widehat{\operatorname{VaR}}_{i, t}\right) .
\end{aligned}
$$

See Chao, Härdle \& Wang (2013): Quantile Regression in Risk Calibration
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## State variables

$M_{t}: 7$ state variables suggested by AB and further extension:

1. VIX
2. Short term liquidity spread
3. Change in the 3 M T-bill rate
4. Change in the slope of the yield curve
5. Change in the credit spread between 10 years BAA-rated bonds and the T-bond rate
6. S\&P500 returns
7. Dow Jones U.S. Real Estate index returns
8. Constituent's specific stock log returns (15x)
9. Constituent's specific stock volatility log returns (15x)

## Least Absolute Shrinkage and Selection Operator (LASSO)

$\square$ Selection of variables with significant effect on CDS spread returns
$\square$ The quantile regression under LASSO penalty

$$
L^{\text {LASSO } \left.^{( } \beta\right)=\sum_{i=1}^{n} \rho_{\tau}\left(y_{i}-\beta^{\top} x_{i}\right)+\lambda_{n} \sum_{j=1}^{p}\left|\beta_{j}\right||.|c| l}
$$

where $0 \leq \tau \leq 1$ and $\lambda_{n}$ denotes the penalty parameter.
$\square \lambda_{n}$ is chosen via generalized approximate cross-validation (GACV) suggested by Yuan (2006) and Li et al. (2007)

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## CDS spread returns

$\square$ Daily CDS spreads of 14 biggest derivative dealers and 1 monoline
$\square$ Overall data period: Sept 2002 - Dec $2011(N=2208)$
$\square$ Segregation into two sub-periods

- pre-shock: Sept 122002 - Sept 122008
- shock event: Lehman Brothers filed for Chapter 11 bankruptcy protection on Sept 152008
- post-shock: Sept 162008 - Dec 312011

Table 1: Descriptive statistics of CDS spread log returns

|  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Std. Dev | Skewness | Kurtosis | Min | Max | Autocorr. |
| CITI | 0.023 | 0.871 | 27.203 | -0.174 | 0.286 | 0.032 |
| BOA | 0.023 | 0.579 | 14.454 | -0.182 | 0.247 | 0.008 |
| BARC | 0.021 | 1.045 | 24.028 | -0.155 | 0.270 | 0.115 |
| BNP | 0.021 | 0.160 | 17.017 | -0.192 | 0.214 | 0.117 |
| CS | 0.019 | 0.172 | 17.983 | -0.168 | 0.182 | 0.065 |
| DB | 0.020 | 0.682 | 22.554 | -0.156 | 0.252 | 0.143 |
| GS | 0.020 | -0.040 | 28.865 | -0.248 | 0.219 | 0.222 |
| HSBC | 0.019 | -0.294 | 13.582 | -0.147 | 0.151 | 0.067 |
| JPM | 0.019 | 0.453 | 15.169 | -0.138 | 0.213 | 0.117 |
| MS | 0.023 | 4.678 | 118.434 | -0.255 | 0.475 | -0.006 |
| RBS | 0.024 | 1.884 | 87.755 | -0.368 | 0.376 | -0.072 |
| SG | 0.020 | -0.209 | 21.404 | -0.223 | 0.187 | 0.129 |
| UBS | 0.020 | 0.439 | 20.372 | -0.153 | 0.218 | 0.090 |
| LEH | 0.019 | -2.040 | 30.336 | -0.226 | 0.148 | 0.138 |
| AIG | 0.024 | 1.106 | 61.673 | -0.253 | 0.402 | 0.237 |

## Estimated Coefficient: $\widehat{\beta_{V I X}}$ - pre-shock

Figure 2: $\widehat{\beta}$ of variable VIX of all 15 FI : 1-Citi, 2-BoA, 3-GS, 4-JPM, 5-MS, 6 -LEH, 7 -AIG, 8 -SG, $9-B N P, 10-\mathrm{CS}, 11-\mathrm{DB}, 12-\mathrm{BARC}, 13-\mathrm{HSBC}, 14-\mathrm{RBS}$, 15-UBS
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## Estimated Coefficient: $\widehat{\beta_{V I X}}$ - post-shock

Figure 3: $\widehat{\beta}$ of variable VIX of all 15 FI : 1-Citi, 2-BoA, 3-GS, 4-JPM, 5-MS, 6 -AIG, $7-S G, 8-B N P, 9-C S, 10-D B, 11-B A R C, 12-H S B C, 13-R B S, 14-U B S$


Figure 4: Backtesting results: Bank of America VaR exceedance under LASSO quantile regression (left) and under AB model (right) in pre-shock period.


Figure 5: Backtesting results: Royal Bank of Scotland VaR exceedance under LASSO quantile regression (left) and under AB model (right) in pre-shock period.

## Backtesting of calculated VaR under AB (2011)

|  | Exceedance | LRpoF | LR $_{\text {uncond }}$ | LRcc | Test Outcome |
| :--- | :---: | ---: | :---: | ---: | :---: |
| CITI | 38 | 38.69 | 0 | 38.69 | Rejected |
| BOA | 39 | 41.17 | 0 | 41.17 | Rejected |
| BARC | 28 | 17.22 | 0 | 17.22 | Rejected |
| BNP | 33 | 27.17 | 0 | 27.17 | Rejected |
| CS | 46 | 59.90 | 0 | 59.90 | Rejected |
| DB | 47 | 62.76 | 0 | 62.76 | Rejected |
| GS | 45 | 57.08 | 0 | 57.08 | Rejected |
| HSBC | 41 | 46.27 | 0 | 46.27 | Rejected |
| JPM | 57 | 93.73 | 0 | 93.73 | Rejected |
| MS | 60 | 103.77 | 0 | 103.77 | Rejected |
| RBS | 40 | 43.70 | 0 | 43.70 | Rejected |
| SG | 31 | 22.99 | 0 | 22.99 | Rejected |
| UBS | 36 | 33.91 | 0 | 33.91 | Rejected |
| LEH | 43 | 51.58 | 0 | 51.58 | Rejected |
| AIG | 57 | 93.73 | 0 | 93.73 | Rejected |

Table 2: Backtesting for $\mathrm{N}=1145$ observations; Test statistic: $\mathrm{LR}_{\text {Pof }}$ for Kupiec test, $\mathrm{LR}_{\text {uncond }}$ for Christoffersen test, $\mathrm{LR}_{\mathrm{CC}}$ for conditional coverage. Credit Risk Calibration based on CDS Spreads

## Backtesting of calculated VaR under QLPLM

|  | Exceedance | LR $\mathbf{P o F}$ | LR $_{\text {uncond }}$ | LR $_{\mathbf{c c}}$ | Test Outcome |
| :--- | :---: | :---: | :---: | ---: | :---: |
| CITI | 18 | 3.22 | 0 | 3.22 | Not Rejected |
| BOA | 20 | 5.27 | 0 | 5.27 | Not Rejected |
| BARC | 15 | 1.01 | 0 | 1.01 | Not Rejected |
| BNP | 19 | 4.19 | 0 | 4.19 | Not Rejected |
| CS | 15 | 1.01 | 0 | 1.01 | Not Rejected |
| DB | 22 | 7.73 | 0 | 7.73 | Not Rejected |
| GS | 26 | 13.73 | 0 | 13.73 | Rejected |
| HSBC | 18 | 3.22 | 0 | 3.22 | Not Rejected |
| JPM | 19 | 4.19 | 0 | 4.19 | Not Rejected |
| MS | 20 | 5.27 | 0 | 5.27 | Not Rejected |
| RBS | 18 | 3.22 | 0 | 3.22 | Not Rejected |
| SG | 21 | 6.45 | 0 | 6.45 | Not Rejected |
| UBS | 16 | 1.62 | 0 | 1.62 | Not Rejected |
| LEH | 33 | 27.17 | 0 | 27.17 | Rejected |
| AIG | 25 | 12.11 | 0 | 12.11 | Rejected |

Table 3: Backtesting for $N=1145$ observations; Test statistic: LR ${ }_{\text {PoF }}$ for Kupiec test, $L R_{\text {uncond }}$ for Christoffersen test, $L R_{\text {CC }}$ for conditional coverage. Credit Risk Calibration based on CDS Spreads

## $\triangle \mathrm{CoVaR}$ in pre-shock period

|  | Citi | BoA | BAR | DB | GS | JPM | MS | RBS | LEH | AIG | sum |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Citi | - | -0.04 | -0.03 | -0.02 | -0.03 | -0.03 | -0.03 | -0.03 | -0.04 | -0.04 | -0.41 |
| BoA | -0.07 | - | -0.04 | -0.03 | -0.05 | -0.05 | -0.04 | -0.04 | -0.04 | -0.04 | -0.58 |
| BAR | -0.01 | -0.04 | - | -0.05 | -0.03 | -0.04 | -0.03 | -0.07 | -0.03 | -0.03 | -0.61 |
| DB | 0.00 | -0.01 | -0.05 | - | -0.03 | -0.03 | -0.03 | -0.04 | -0.01 | -0.02 | -0.37 |
| GS | -0.05 | -0.04 | -0.02 | -0.02 | - | -0.04 | -0.04 | -0.03 | -0.03 | -0.04 | -0.46 |
| JPM | -0.05 | -0.05 | -0.03 | -0.03 | -0.04 | - | -0.03 | -0.03 | -0.03 | -0.04 | -0.52 |
| MS | -0.04 | -0.03 | -0.03 | -0.03 | -0.05 | -0.03 | - | -0.03 | -0.03 | -0.05 | -0.43 |
| RBS | -0.03 | -0.02 | -0.12 | -0.07 | -0.02 | -0.04 | -0.02 | - | -0.03 | -0.02 | -0.78 |
| LEH | -0.04 | -0.04 | -0.03 | -0.03 | -0.04 | -0.04 | -0.03 | -0.03 | - | -0.04 | -0.46 |
| AIG | -0.02 | -0.02 | -0.01 | -0.02 | -0.03 | -0.03 | -0.02 | -0.02 | -0.02 | - | -0.28 |

Table 4: Average $\triangle$ CoVaR overview for pre-shock period.

## $\triangle \mathrm{CoVaR}$ in post-shock period

|  | Citi | BoA | BAR | DB | GS | JPM | MS | RBS | SG | AIG | sum |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Citi | - | -0.16 | -0.07 | -0.05 | -0.15 | -0.15 | -0.11 | -0.07 | -0.08 | -0.11 | -0.96 |
| BoA | -0.19 | - | -0.14 | -0.13 | -0.20 | -0.19 | -0.18 | -0.13 | -0.16 | -0.11 | -1.45 |
| BAR | -0.11 | -0.15 | - | -0.10 | -0.12 | -0.12 | -0.08 | -0.14 | -0.13 | -0.10 | -1.06 |
| DB | -0.15 | -0.16 | -0.13 | - | -0.19 | -0.18 | -0.17 | -0.20 | -0.20 | -0.16 | -1.54 |
| GS | -0.21 | -0.20 | -0.13 | -0.15 | - | -0.22 | -0.18 | -0.14 | -0.17 | -0.14 | -1.53 |
| JPM | -0.17 | -0.18 | -0.09 | -0.12 | -0.17 | - | -0.17 | -0.14 | -0.15 | -0.13 | -1.32 |
| MS | -0.11 | -0.13 | -0.07 | -0.08 | -0.17 | -0.14 | - | -0.10 | -0.11 | -0.13 | -1.03 |
| RBS | -0.10 | -0.17 | -0.12 | -0.16 | -0.17 | -0.12 | -0.12 | - | -0.14 | -0.16 | -1.25 |
| SG | -0.15 | -0.25 | -0.13 | -0.14 | -0.21 | -0.24 | -0.18 | -0.22 | - | -0.17 | -1.69 |
| AIG | -0.01 | -0.03 | -0.05 | -0.06 | -0.04 | -0.03 | -0.04 | -0.05 | -0.04 | - | -0.35 |

Table 5: Average $\triangle$ CoVaR overview for post-shock period

## Average $\triangle$ CoVaR in the pre-shock period



Figure 6: Network of spread spillover effects described by average $\triangle \mathrm{CoVaR}$


## Average $\triangle \mathrm{CoVaR}$ in the post-shock period



Figure 7: Network of spread spillover effects described by average $\triangle \mathrm{CoVaR}$

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## Change in $\triangle$ CoVaR during the pre-shock period

Figure 8: Network of spread spillover effects described by $\triangle \mathrm{CoVaR}$ Credit Risk Calibration based on CDS Spreads

## Study of CDS spreads determinants

$\square$ CDS spread returns mainly described by implied volatility index VIX and real estate sector returns
$\square$ Strong positive relationship between CDS spread returns and equity volatility index
$\square$ Heterogeneous impact in regions: high sensitivity of US FIs to VIX after shock, delayed in sensitivity for European FIs.
$\square$ Effects of firm specific volatility is not as strong as market volatility indicated by VIX index

## Study of $\triangle$ CoVaR

$\square$ Continental effects shown by $\triangle$ CoVaR: higher value observed between FIs from the same region
$\square \triangle$ CoVaR more suitable for computing stressed $\mathrm{VaR}(\mathrm{VaR}$ under data of financial crisis) rather than for CDS spread forecasting, especially in late post-shock period Next steps:
$\square \triangle \mathrm{CoVaR}$ as risk weighting basis for transactions cleared through CCP
$\square \triangle$ CoVaR of CDS index on corporate companies for estimation of portfolio potential future exposure (PFE)

## Credit Risk Calibration based on CDS Spreads <br> Shih-Kang Chao <br> Wolfgang Karl Härdle Hien Pham-Thu

Ladislaus von Bortkiewicz Chair of Statistics
 C.A.S.E. - Center for Applied Statistics and Economics Humboldt-Universität zu Berlin http://lvb.wiwi.hu-berlin.de

## Partial Linear Model (PLM)

$\square$ The partial linearity observation implies:

$$
\begin{align*}
& X_{i, t}=\alpha_{i}+\gamma_{i}^{\top} M_{t-1}+\varepsilon_{i, t} \\
& X_{j, t}=\tilde{\beta}_{j \mid i}^{\top} M_{t-1}+I_{j \mid i}\left(X_{i, t}\right)+\varepsilon_{j, t} . \tag{1}
\end{align*}
$$

$I$ : a general function. $M_{t}$ : state variables. $F_{\varepsilon_{i, t}}^{-1}\left(\tau \mid M_{t-1}\right)=0$ and $F_{\varepsilon_{j, t}}^{-1}\left(\tau \mid M_{t-1}, X_{i, t}\right)=0$.
$\square$ Advantages

- Capturing nonlinear asset dependence
- Avoid curse of dimensionality


## Estimation of Partial Linear Model

$\checkmark$ PLM model: Liang, Härdle and Carroll (1999) and Härdle, Ritov and Song (2012)

$$
Y_{t}=\beta^{\top} M_{t-1}+I\left(X_{t}\right)+\varepsilon_{t}
$$

$\square$ Consider $[0,1]$ (standard rank space). Dividing $[0,1]$ into $a_{n}$ equally divided subintervals $I_{n t}, a_{n} \uparrow \infty$. On each subinterval, $I(\cdot)$ is roughly constant.

## Estimation of PLM QR

1. Linear element $\beta$ :

$$
\begin{aligned}
& \hat{\beta}= \\
& \underset{\beta}{\operatorname{argmin}} \min _{I_{1}, \ldots, I_{a_{n}}} \sum_{t=1}^{n} \rho_{\tau}\left\{Y_{t}-\beta^{\top} M_{t-1}-\sum_{m=1}^{a_{n}} I_{m} \mathbf{1}\left(X_{t} \in I_{n t}\right)\right\}
\end{aligned}
$$

2. Nonlinear element $I(\cdot)$ : With data $\left\{\left(X_{t}, Y_{t}-\hat{\beta}^{\top} M_{t-1}\right)\right\}_{t=1}^{n}$, applying LLQR.

## $\triangle$ CoVaR in pre-shock period

|  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Citi | BoA | BARC | DB | GS | JPM | MS | RBS | LEH | AIG |
| Citi | - | -0.37 | -0.23 | -0.27 | -0.35 | -0.32 | -0.27 | -0.34 | -0.42 | -0.45 |
| BoA | -0.52 | - | -0.33 | -0.26 | -0.29 | -0.27 | -0.21 | -0.50 | -0.33 | -0.43 |
| BARC | -0.42 | -0.29 | - | -0.35 | -0.42 | -0.35 | -0.30 | -0.46 | -0.58 | -0.52 |
| DB | -0.23 | -0.22 | -0.52 | - | -0.16 | -0.21 | -0.24 | -0.52 | -0.29 | -0.50 |
| GS | -0.27 | -0.28 | -0.29 | -0.22 | - | -0.22 | -0.27 | -0.61 | -0.34 | -0.28 |
| JPM | -0.29 | -0.25 | -0.20 | -0.23 | -0.24 | - | -0.46 | -0.50 | -0.45 | -0.26 |
| MS | -0.27 | -0.25 | -0.50 | -0.36 | -0.37 | -0.23 | - | -0.56 | -0.27 | -0.47 |
| RBS | -0.32 | -0.35 | -1.67 | -0.80 | -0.16 | -0.55 | -0.22 | - | -0.46 | -0.46 |
| LEH | -0.35 | -0.29 | -0.26 | -0.32 | -0.30 | -0.25 | -0.29 | -0.27 | - | -0.32 |
| AIG | -0.34 | -0.32 | -0.36 | -0.21 | -0.28 | -0.21 | -0.27 | -0.52 | -0.36 | - |
|  |  |  |  |  |  |  |  |  |  |  |

Table 6: Minimum $\triangle$ CoVaR overview for pre-shock period which demonstrates the maximum negative effects on CDS spreads returns.

## $\triangle \mathrm{CoVaR}$ in post-shock period

|  | Citi | BoA | BARC | DB | GS | JPM | MS | RBS | SG | AIG |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Citi | - | -0.79 | -0.97 | -0.79 | -1.03 | -1.55 | -1.36 | -1.06 | -0.51 | -1.24 |
| BoA | -0.84 | - | -0.55 | -0.58 | -0.83 | -0.58 | -1.19 | -0.45 | -0.65 | -0.56 |
| BARC | -1.72 | -0.78 | - | -0.58 | -0.90 | -0.46 | -0.42 | -0.95 | -0.47 | -0.74 |
| DB | -1.41 | -0.82 | -0.97 | - | -1.60 | -1.52 | -1.32 | -0.74 | -2.19 | -1.35 |
| GS | -0.90 | -1.18 | -0.63 | -1.09 | - | -0.73 | -1.99 | -1.51 | -0.94 | -1.66 |
| JPM | -0.58 | -0.54 | -0.34 | -0.42 | -0.55 | - | -1.07 | -0.44 | -0.61 | -0.77 |
| MS | -1.26 | -0.94 | -0.83 | -1.05 | -0.95 | -0.89 | - | -1.40 | -1.14 | -2.31 |
| RBS | -0.69 | -0.67 | -0.39 | -0.52 | -0.81 | -0.55 | -0.47 | - | -0.61 | -0.64 |
| SG | -0.89 | -1.02 | -0.38 | -0.44 | -0.90 | -0.79 | -0.71 | -0.63 | - | -0.54 |
| AIG | -0.61 | -0.41 | -0.65 | -0.71 | -0.37 | -0.49 | -0.58 | -0.78 | -0.31 | - |
|  |  |  |  |  |  |  |  |  |  |  |

Table 7: Minimum $\triangle$ CoVaR overview for post-shock period which demonstrates the maximum negative effects on CDS spreads returns.

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